

A Omitted proofs

Proof of Proposition 1. We recall the I-MMSE relation (11):

$$\frac{d}{d\beta} \frac{1}{\lambda_N} D(Q_{\beta\lambda_N, N} \| Q_{0, N}) = \frac{1}{2} - \frac{1}{2} \text{MMSE}_N(\beta\lambda_N).$$

Let us first assume that the all-or-nothing phenomenon holds. Since $D(Q_{0, N} \| Q_{0, N}) = 0$, we can write

$$\begin{aligned} \lim_{N \rightarrow \infty} \frac{1}{\lambda_N} D(Q_{\beta\lambda_N, N} \| Q_{0, N}) &= \lim_{N \rightarrow \infty} \int_0^\beta \frac{d}{d\kappa} \frac{1}{\lambda_N} D(Q_{\kappa\lambda_N, N} \| Q_{0, N}) d\kappa \\ &= \lim_{N \rightarrow \infty} \int_0^\beta \frac{1}{2} - \frac{1}{2} \text{MMSE}_N(\kappa\lambda_N) d\kappa \\ &\stackrel{(*)}{=} \int_0^\beta \frac{1}{2} - \lim_{N \rightarrow \infty} \text{MMSE}_N(\kappa\lambda_N) d\kappa \\ &= \frac{1}{2}(\beta - 1)_+, \end{aligned}$$

where in $(*)$ we have used the dominated convergence theorem and the fact that $\text{MMSE}_N(\kappa\lambda_N) \in [0, 1]$ and where the last equality follows from the all-or-nothing phenomenon.

In the other direction, we use the fact that $\text{MMSE}_N(\beta\lambda_N)$ is a non-increasing function of β [see, e.g., Mio19, Proposition 1.3.1]. Combined with the I-MMSE relation, this immediately yields that $\frac{1}{\lambda_N} D(Q_{\beta\lambda_N, N} \| Q_{0, N})$ is convex. We therefore have by standard facts in convex analysis [HUL93, Proposition 4.3.4] that

$$\frac{1}{2} - \frac{1}{2} \lim_{N \rightarrow \infty} \text{MMSE}_N(\beta\lambda_N) = \lim_{N \rightarrow \infty} \frac{d}{d\beta} \frac{1}{\lambda_N} D(Q_{\beta\lambda_N, N} \| Q_{0, N}) = \frac{d}{d\beta} \left(\lim_{N \rightarrow \infty} \frac{1}{\lambda_N} D(Q_{\beta\lambda_N, N} \| Q_{0, N}) \right)$$

for all β for which the right side exists. Since we have assumed that

$$\lim_{N \rightarrow \infty} \frac{1}{\lambda_N} D(Q_{\beta\lambda_N, N} \| Q_{0, N}) = \frac{1}{2}(\beta - 1)_+,$$

the right side is 0 when $\beta < 1$ and $\frac{1}{2}$ when $\beta > 1$. The all-or-nothing property immediately follows. \square

Proof of Proposition 2. The first claim follows directly from Lemma 6. Indeed, for the sparse vector model, $\log M_p = (1 + o(1))k \log \frac{p}{k}$, and by Lemma 6,

$$\lim \frac{1}{k \log \frac{p}{k}} \log P_p^{\otimes 2}[\langle \mathbf{X}, \mathbf{X}' \rangle \geq t] = -t. \quad (7)$$

Since $t < \frac{2t}{1+t}$ for all $t \in (0, 1)$, the claim holds.

We now turn to the proof of the all-or-nothing phenomenon. By Theorem 1, it suffices to show

$$D(Q_{2 \log M_p, p} \| Q_{0, p}) = o(\log M_p).$$

We write

$$\begin{aligned} D(Q_{2 \log M_p, p} \| Q_{0, p}) &= \mathbb{E}_{\mathbf{Y} \sim Q_{2 \log M_p, p}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \exp \left(\sqrt{2 \log M_p} \langle \mathbf{Y}, \mathbf{X}' \rangle - \log M_p \right) \\ &= \mathbb{E}_{\mathbf{X}} \mathbb{E}_{\mathbf{Z}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \exp \left(\sqrt{2 \log M_p} \langle \mathbf{Z}, \mathbf{X}' \rangle + 2 \log M_p \langle \mathbf{X}, \mathbf{X}' \rangle - \log M_p \right) \end{aligned}$$

Now, given \mathbf{X} and any vector $v \in \mathbb{R}^p$, let us denote by $v|_{\mathbf{X}} \in \mathbb{R}^p$ the vector given by

$$(v|_{\mathbf{X}})_i := \begin{cases} v_i & \text{if } \mathbf{X}_i \neq 0, \\ 0 & \text{otherwise.} \end{cases}$$

Similarly, let $v|_{\mathbf{X}^c} := v - v|_{\mathbf{X}}$. Given \mathbf{X} , the vectors $\mathbf{Z}|_{\mathbf{X}}$ and $\mathbf{Z}|_{\mathbf{X}^c}$ are independent; thus we can apply Jensen's inequality to the expectation with respect to $\mathbf{Z}|_{\mathbf{X}^c}$ to obtain

$$\begin{aligned} D(Q_{2 \log M_p, p} \| Q_{0, p}) &\leq \mathbb{E}_{\mathbf{X}} \mathbb{E}_{\mathbf{Z}|_{\mathbf{X}}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \mathbb{E}_{\mathbf{Z}|_{\mathbf{X}^c}} \exp \left(\sqrt{2 \log M_p} \langle \mathbf{Z}, \mathbf{X}' \rangle + 2 \log M_p \langle \mathbf{X}, \mathbf{X}' \rangle - \log M_p \right) \\ &= \mathbb{E}_{\mathbf{X}} \mathbb{E}_{\mathbf{Z}|_{\mathbf{X}}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \exp \left(\sqrt{2 \log M_p} \langle \mathbf{Z}|_{\mathbf{X}}, \mathbf{X}'|_{\mathbf{X}} \rangle + \log M_p (\|\mathbf{X}'|_{\mathbf{X}^c}\|^2 + 2 \langle \mathbf{X}, \mathbf{X}' \rangle - 1) \right). \end{aligned}$$

Since the entries of \mathbf{X} and \mathbf{X}' are all either 0 or $1/\sqrt{k}$ and \mathbf{X} has unit norm, we have that $\langle \mathbf{X}, \mathbf{X}' \rangle = \|\mathbf{X}'|_{\mathbf{X}}\|^2$, and since $\mathbf{X}'|_{\mathbf{X}^c}$ and $\mathbf{X}'|_{\mathbf{X}}$ are orthogonal, we obtain

$$\|\mathbf{X}'|_{\mathbf{X}^c}\|^2 + 2 \langle \mathbf{X}, \mathbf{X}' \rangle - 1 = \langle \mathbf{X}, \mathbf{X}' \rangle.$$

Continuing from above and using that $\|\mathbf{X}'|_{\mathbf{X}}\|_{\infty} \leq 1/\sqrt{k}$, we have

$$\begin{aligned} D(Q_{2 \log M_p, p} \| Q_{0, p}) &\leq \mathbb{E}_{\mathbf{X}} \mathbb{E}_{\mathbf{Z}|_{\mathbf{X}}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \exp \left(\sqrt{2 \log M_p/k} \|\mathbf{Z}|_{\mathbf{X}}\|_1 + \log M_p \langle \mathbf{X}, \mathbf{X}' \rangle \right) \\ &= \mathbb{E}_{\mathbf{X}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \exp (\log M_p \langle \mathbf{X}, \mathbf{X}' \rangle) + \mathbb{E}_{\mathbf{X}} \mathbb{E}_{\mathbf{Z}|_{\mathbf{X}}} \sqrt{(2 \log M_p/k)} \|\mathbf{Z}|_{\mathbf{X}}\|_1 \\ &\leq \mathbb{E}_{\mathbf{X}} \log \mathbb{E}_{\mathbf{X}' \sim P_p} \exp (\log M_p \langle \mathbf{X}, \mathbf{X}' \rangle) + O \left(\sqrt{2k \log M_p} \right) \end{aligned}$$

Since $k = o(p)$, we also have that $k = o(k \log \frac{p}{k}) = o(\log M_p)$; therefore, the second term is $o(\log M_p)$. Hence it suffices to focus on the first term.

We proceed via a large deviations argument as in the proof of Theorem 4. Write $\rho = \langle \mathbf{X}, \mathbf{X}' \rangle$ for the overlap; note that the law of ρ is the same for all \mathbf{X} in the support of P_p , so it suffices to understand $\log \mathbb{E} \exp (\rho \log M_p)$. We have, for any fixed positive integer ℓ ,

$$\begin{aligned} \mathbb{E} \exp (\rho \log M_p) &\leq \sum_{m=0}^{\ell-1} P_N[\rho \geq m/\ell] \exp \left(\frac{m+1}{\ell} \log M_p \right) \\ &\leq \ell \cdot \max_{0 \leq m < \ell} \exp \left(\frac{m+1}{\ell} \log M_p + \log P_N[\rho \geq m/\ell] \right), \end{aligned}$$

which implies

$$\limsup_{p \rightarrow \infty} \frac{1}{\log M_p} \log \mathbb{E} \exp (\rho \log M_p) \leq \max_{0 \leq m < \ell} \frac{m+1}{\ell} - \frac{m}{\ell},$$

where we have used (7). Therefore $\limsup_{p \rightarrow \infty} \frac{1}{\log M_p} \log \mathbb{E} \exp (\rho \log M_p) = O(1/\ell)$, and letting $\ell \rightarrow \infty$ proves the claim. \square

Proof of Proposition 3. Denote by \mathcal{S}_k the set of k -sparse vectors in \mathbb{R}^p . Note that the cardinality of $\{0, 1/\sqrt{k}\}^p \cap \mathcal{S}_k$ is $\binom{p}{k}$ and the cardinality of $\{-1/\sqrt{k}, 0, 1/\sqrt{k}\}^p \cap \mathcal{S}_k$ is $\binom{p}{k} 2^k$. In the case of the Bernoulli prior, the identification $\mathbf{x} \mapsto x^{\otimes d}$ is a bijection, so M_N for the Bernoulli prior is $\binom{p}{k}$. In the case of the Bernoulli-Rademacher prior, when d is odd the map $\mathbf{x} \mapsto x^{\otimes d}$ is still a bijection, but when d is even, the vectors \mathbf{x} and $-\mathbf{x}$ give rise to the same tensor. Therefore M_N for the Bernoulli-Rademacher prior is either $\binom{p}{k} 2^k$ or $\binom{p}{k} 2^{k-1}$. Nevertheless, using Stirling's approximation, since $k = o(p)$, we have for both the Bernoulli and Bernoulli-Rademacher prior that

$$\log M_N = (1 + o(1))k \log \frac{p}{k}.$$

Now notice that the overlap $\langle \mathbf{X}, \mathbf{X}' \rangle$ in the case that \mathbf{x} is Bernoulli-Rademacher is stochastically dominated by the overlap when \mathbf{x} is Bernoulli. To prove this, let us consider the natural coupling between the two different priors on \mathbf{x} : we first sample \mathbf{x}_1 from the sparse Bernoulli distribution and then choose uniformly at random the signs for the non-zero values of \mathbf{x}_1 to form a sample \mathbf{x}_2 from the Bernoulli-Rademacher distribution. Notice that by triangle inequality under this coupling it holds almost surely

$$\langle \mathbf{x}_2^{\otimes d}, \mathbf{x}_2'^{\otimes d} \rangle \leq |\langle \mathbf{x}_2^{\otimes d}, \mathbf{x}_2'^{\otimes d} \rangle| \leq \langle \mathbf{x}_1^{\otimes d}, \mathbf{x}_1'^{\otimes d} \rangle.$$

For this reason it suffices to prove our result only in the case the prior \tilde{P}_p is the uniform distribution over $\{0, 1/\sqrt{k}\}^p \cap \mathcal{S}_k$. We therefore focus on this case in the rest of the proof.

Now fix any $t \in [0, 1]$ and notice that by elementary algebra for any $v, v' \in \mathbb{R}^p$ with $\|v\| = \|v'\| = 1$ since $d \geq 2$ it holds $\langle v^{\otimes d}, v'^{\otimes d} \rangle = \langle v, v' \rangle^d \leq \langle v, v' \rangle^2$. Hence as \mathbf{x}, \mathbf{x}' live on the sphere of dimension p ,

$$\begin{aligned} \mathbb{P}_N^{\otimes 2}[\langle \mathbf{X}, \mathbf{X}' \rangle \geq t] &= \tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}^{\otimes d}, \mathbf{x}'^{\otimes d} \rangle \geq t] = \tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle^d \geq t] \\ &\leq \tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle^2 \geq t] \\ &= \tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle \geq \sqrt{t}]. \end{aligned} \quad (8)$$

Since \mathbf{x}, \mathbf{x}' are drawn from the uniform distribution over $\{0, 1/\sqrt{k}\}^p \cap \mathcal{S}_k$, Lemma 6 combined with (8) yields

$$\lim_{N \rightarrow +\infty} \frac{1}{\log M_N} \log \mathbb{P}_N^{\otimes 2}[\langle \mathbf{X}, \mathbf{X}' \rangle \geq t] \leq -\sqrt{t}.$$

The elementary inequality $-\sqrt{t} \leq -\frac{2t}{1+t}$ concludes the proof. \square

Proof of Proposition 4. Let

$$Z(Y) = \frac{Q_{\lambda_N, N}(Y)}{Q_{0, N}(Y)} = \mathbb{E}_{\mathbf{X}' \sim P_N} \exp\left(\sqrt{\lambda_N} \langle Y, \mathbf{X}' \rangle - \frac{\lambda_N}{2}\right)$$

Following *mutatis mutandis* the first two arguments in the proof of [BMV⁺18, Theorem 5] we obtain

$$D(Q_{\lambda_N, N} \parallel Q_{0, N}) \leq D(\tilde{Q}_{\lambda_N, N} \parallel Q_{0, N}) + o(1) \cdot \sqrt{\mathbb{E}_{\mathbf{Y} \sim Q_{\lambda_N, N}}[\log^2 Z(\mathbf{Y})]}. \quad (9)$$

It is straightforward to see that for all Y ,

$$|\log Z(Y)| \leq \sqrt{\lambda_N} \max_{X' \in \text{Support}(P_N)} \langle X', Y \rangle + \frac{\lambda_N}{2}$$

which implies that

$$\mathbb{E}_{\mathbf{Y} \sim Q_{\lambda_N, N}} \log^2 Z(\mathbf{Y}) \leq 2\lambda_N \cdot \mathbb{E}_{\mathbf{Y} \sim Q_{\lambda_N, N}} \max_{X' \in \text{Support}(P_N)} \langle X', \mathbf{Y} \rangle^2 + O(\lambda_N^2). \quad (10)$$

Now recall $\mathbf{Y} = \sqrt{\lambda_N} \mathbf{X} + \mathbf{Z}$ for $\mathbf{Z} \sim Q_{0, N}$ and for all $X' \in \text{Support}(P_N)$ it holds $|\langle \mathbf{X}, X' \rangle| \leq \|\mathbf{X}\| \|X'\| = 1$ almost surely. Hence,

$$\begin{aligned} \mathbb{E}_{\mathbf{Y} \sim Q_{\lambda_N, N}} \max_{X' \in \text{Support}(P_N)} \langle X', \mathbf{Y} \rangle^2 &= \mathbb{E}_{\mathbf{Z} \sim Q_{0, N}} \left(\max_{X' \in \text{Support}(P_N)} |\sqrt{\lambda_N} \langle X', \mathbf{X} \rangle + \langle X', \mathbf{Z} \rangle| \right)^2 \\ &\leq 2\lambda_N + 2\mathbb{E}_{\mathbf{Z} \sim Q_{0, N}} \max_{X' \in \text{Support}(P_N)} \langle X', \mathbf{Z} \rangle^2. \end{aligned}$$

Since $Q_{0, N}$ is simply the law of a vector with i.i.d. standard Gaussian coordinates and the cardinality of the discrete subset of the sphere $\text{Support}(P_N)$ is equal to M_N , by Lemma 5 we have $\mathbb{E}_{\mathbf{Z} \sim Q_{0, N}} \max_{X' \in \text{Support}(P_N)} \langle X', \mathbf{Z} \rangle^2 = O(\log M_N)$. Therefore since $\lambda_N = O(\log M_N)$,

$$\mathbb{E}_{\mathbf{Y} \sim Q_{\lambda_N, N}} \max_{X' \in \text{Support}(P_N)} \langle X', \mathbf{Y} \rangle^2 \leq O(\lambda_N + \log M_N) = O(\log M_N).$$

Combining the last inequality with (10), we conclude that

$$\mathbb{E}_{\mathbf{Y} \sim Q_{\lambda_N, N}} \log^2 Z(\mathbf{Y}) = O(\lambda_N^2) = O(\log^2 M_N).$$

Using (9) completes the proof of the proposition. \square

Proof of Proposition 5. We let C denote an absolute positive constant whose value may change from line to line. Let us write $\mathbf{W} = \langle X, \mathbf{Z} \rangle / \sqrt{\lambda_N}$ and $\mathbf{W}' = \langle X', \mathbf{Z} \rangle / \sqrt{\lambda_N}$. Recall that X, X' lie on the unit sphere with $\langle X, X' \rangle = \rho$.

Then \mathbf{W} and \mathbf{W}' are jointly Gaussian with mean 0 and covariance $\frac{1}{\lambda_N} \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} =: \frac{1}{\lambda_N} \Sigma_\rho$. Under this parametrization, we have

$$\exp(\sqrt{\lambda_N}(\langle X, \mathbf{Z} \rangle + \langle X', \mathbf{Z}' \rangle) - \lambda_N) = \exp(\lambda_N(\mathbf{W} + \mathbf{W}' - 1)).$$

Let us write S for the set $\{(w, w') : |w - 1| \leq \lambda_N^{-1/4}, |w' - 1| \leq \lambda_N^{-1/4}\}$.

We consider three cases:

Case 1: $\rho \leq 0$ Using the moment generating function of the univariate normal distribution yields

$$\mathbb{E} \exp(\lambda_N(\mathbf{W} + \mathbf{W}' - 1)) \mathbb{1}_S(\mathbf{W}, \mathbf{W}') \leq \mathbb{E} \exp(\lambda_N(\mathbf{W} + \mathbf{W}' - 1)) = e^{\lambda_N \rho} \leq 1,$$

so

$$\frac{1}{\lambda_N} \log m_N(\rho) \leq 0 = \left(\frac{\rho}{1 + \rho} \right)_+.$$

Case 2: $\rho \in (0, 1/2]$ Write $\phi_\rho(w, w')$ for the joint density of \mathbf{W} and \mathbf{W}' . Note that on S

$$\begin{aligned} \phi_\rho(w, w') &\leq \frac{\lambda_N}{2\pi(1 - \rho^2)} \exp\left(-\frac{\lambda_N}{2} \mathbf{w}^\top \Sigma_\rho^{-1} \mathbf{w}\right), \quad \mathbf{w} = (w, w') \\ &\leq C e^{-\frac{\lambda_N}{1+\rho} + C\lambda_N^{3/4}}, \end{aligned}$$

where we use that $\lambda_N \rightarrow +\infty$ as $N \rightarrow +\infty$. Hence

$$\begin{aligned} \frac{1}{\lambda_N} \log m_N(\rho) &= \frac{1}{\lambda_N} \log \int_S e^{\lambda_N(w+w'-1)} \phi_\rho(w, w') dw dw' \\ &\leq \frac{1}{\lambda_N} \log \int_S \max_{(w, w') \in S} e^{\lambda_N(w+w'-1)} \cdot \max_{(w, w') \in S} \phi_\rho(w, w') dw dw' \\ &\leq \frac{1}{\lambda_N} \log(\text{vol}(S) \cdot e^{\lambda_N + O(\lambda_N^{3/4})} \cdot C e^{-\frac{\lambda_N}{1+\rho} + C\lambda_N^{3/4}}) \\ &\leq \frac{\rho}{1 + \rho} + \frac{C}{\lambda_N^{1/4}}. \end{aligned}$$

Case 3: $\rho \in (1/2, 1]$ The sum $\mathbf{W} + \mathbf{W}'$ is Gaussian with mean 0 and variance $\frac{2}{\lambda_N}(1 + \rho)$, and if $(w, w') \in S$, then $|w + w' - 2| \leq 2\lambda_N^{-1/4}$.

We obtain

$$m_N(\rho) = \mathbb{E} \exp(\lambda_N(\mathbf{W} + \mathbf{W}' - 1)) \mathbb{1}_S(\mathbf{W}, \mathbf{W}') \leq \mathbb{E} \exp(\lambda_N(\mathbf{W}'' - 1)) \mathbb{1}_{|\mathbf{W}'' - 2| \leq 2\lambda_N^{-1/4}},$$

where $\mathbf{W}'' \sim \mathcal{N}(0, \frac{2}{\lambda_N}(1 + \rho))$. Similar with the analysis in Case 2, the density of \mathbf{W}'' is bounded by $C e^{-\frac{\lambda_N}{1+\rho} + C\lambda_N^{3/4}}$ on the set $T := \{w'' : |w'' - 2| \leq 2\lambda_N^{-1/4}\}$, and we obtain

$$\begin{aligned} \frac{1}{\lambda_N} \log m_N(\rho) &\leq \frac{1}{\lambda_N} \log \int_T \max_{w'' \in T} e^{\lambda_N(w'' - 1)} \cdot C e^{-\frac{\lambda_N}{1+\rho} + C\lambda_N^{3/4}} \\ &\leq \frac{1}{\lambda_N} \log(\text{vol}(T) \cdot e^{\lambda_N + O(\lambda_N^{3/4})} \cdot C e^{-\frac{\lambda_N}{1+\rho} + C\lambda_N^{3/4}}) \\ &\leq \frac{\rho}{1 + \rho} + \frac{C}{\lambda_N^{1/4}}, \end{aligned}$$

as claimed. \square

B Additional lemmas

Lemma 1. For all N and $\lambda > 0$, the function $\beta \mapsto \frac{1}{\lambda} \text{D}(\mathbb{Q}_{\beta\lambda, N} \| \mathbb{Q}_{0, N})$ is nonnegative, nondecreasing, and $1/2$ -Lipschitz.

Proof. Let us fix some N and λ . The nonnegativity follows from the nonnegativity of the KL divergence. By Lemma 2, we have

$$\frac{1}{\lambda} D(Q_{\beta\lambda, N} \| Q_{0, N}) = \frac{\beta}{2} - \frac{1}{\lambda} I_{\beta\lambda, N}(\mathbf{X}; \mathbf{Y}).$$

Differentiating with respect to β and using the I-MMSE theorem [GSV05] we conclude

$$\frac{d}{d\beta} \frac{1}{\lambda} D(Q_{\beta\lambda, N} \| Q_{0, N}) = \frac{1}{2} - \frac{1}{2} \text{MMSE}_N(\beta\lambda). \quad (11)$$

The results that $\beta \mapsto \frac{1}{\lambda} D(\mathbf{Y}_{\beta\lambda} \| \mathbf{Z})$ is nondecreasing and $1/2$ -Lipschitz follow directly from the fact that $\text{MMSE}_N(\beta\lambda) \in [0, 1]$. \square

Lemma 2. Denote by $I_{\lambda, N}(\mathbf{X}; \mathbf{Y})$ the mutual information between \mathbf{X} and \mathbf{Y} in (1), and denote by $Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})}$ their joint law. Then

$$I_{\lambda, N}(\mathbf{X}; \mathbf{Y}) = D(Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})} \| P_N \otimes Q_{\lambda, N}) = \frac{\lambda}{2} - D(Q_{\lambda, N} \| Q_{0, N}).$$

Proof. The first equality is the definition of mutual information. We then have

$$\begin{aligned} D(Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})} \| P_N \otimes Q_{\lambda, N}) &= \mathbb{E}_{Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})}} \log \frac{Q_{\lambda, N}(\mathbf{Y} | \mathbf{X})}{Q_{\lambda, N}(\mathbf{Y})} \\ &= \mathbb{E}_{Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})}} \log \frac{Q_{\lambda, N}(\mathbf{Y} | \mathbf{X})}{Q_{0, N}(\mathbf{Y})} - \mathbb{E}_{Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})}} \log \frac{Q_{\lambda, N}(\mathbf{Y})}{Q_{0, N}(\mathbf{Y})}. \end{aligned}$$

Using the fact that \mathbf{Z} has i.i.d. standard Gaussian entries we have

$$\mathbb{E}_{Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})}} \log \frac{Q_{\lambda, N}(\mathbf{Y} | \mathbf{X})}{Q_{0, N}(\mathbf{Y})} = \mathbb{E}_{Q_{\lambda, N}^{(\mathbf{X}, \mathbf{Y})}} \frac{\|\mathbf{Y}\|_2^2 - \|\mathbf{Y} - \sqrt{\lambda}\mathbf{X}\|_2^2}{2} = \frac{\lambda}{2},$$

and by definition

$$D(Q_{\lambda, N} \| Q_{0, N}) = \mathbb{E}_{Q_{\lambda, N}} \log \frac{Q_{\lambda, N}(\mathbf{Y})}{Q_{0, N}(\mathbf{Y})}.$$

The claim follows. \square

Lemma 3. For all $\lambda \geq 0$,

$$D(Q_{\lambda, N} \| Q_{0, N}) \geq \frac{\lambda}{2} - \log M_N.$$

Proof. Writing explicitly the Kullback-Leibler divergence gives

$$\begin{aligned} D(Q_{\lambda, N} \| Q_{0, N}) &= \mathbb{E} \log \frac{1}{M_N} \sum_{X' \in \text{Support}(P_N)} \exp\left(\sqrt{\lambda}\langle \mathbf{Y}, X' \rangle - \frac{\lambda}{2}\right) \quad \mathbf{Y} \sim Q_{\lambda, N} \\ &\geq \mathbb{E} \log \frac{1}{M_N} \exp\left(\sqrt{\lambda}\langle \mathbf{Z}, \mathbf{X} \rangle + \frac{\lambda}{2}\right) \\ &= \mathbb{E} \left\{ \sqrt{\lambda}\langle \mathbf{Z}, \mathbf{X} \rangle + \frac{\lambda}{2} - \log M_N \right\} = \frac{\lambda}{2} - \log M_N, \end{aligned}$$

where the inequality follows from writing $\mathbf{Y} = \sqrt{\lambda}\mathbf{X} + \mathbf{Z}$ and taking only the $X' = \mathbf{X}$ term in the sum. \square

Lemma 4. Let $\alpha_1 = (\alpha_1)_{N \in \mathbb{N}}$ and $\alpha_2 = (\alpha_2)_{N \in \mathbb{N}}$ be two sequences in $[0, 1]$ such that $\alpha_1 = 1 - o(1)$ and $\alpha_2 = o(1)$ as $N \rightarrow \infty$, and let λ_N be any sequence tending to infinity as $N \rightarrow +\infty$ such that $\frac{1}{\lambda_N} d(\alpha_1 \| \alpha_2)$ is bounded. Then

$$\limsup_{N \rightarrow \infty} \frac{1}{\lambda_N} d(\alpha_1 \| \alpha_2) = \limsup_{N \rightarrow \infty} \frac{1}{\lambda_N} \log \frac{1}{\alpha_2}.$$

Proof. The given asymptotics imply

$$\lim_{N \rightarrow \infty} (1 - \alpha_1) \log \frac{1 - \alpha_1}{1 - \alpha_2} = 0.$$

Moreover, since $\alpha_1 \log \alpha_1$ is bounded, we have

$$\lim_{N \rightarrow \infty} \frac{1}{\lambda_N} \alpha_1 \log \alpha_1 = 0.$$

Combining these facts yields

$$\begin{aligned} \limsup_{N \rightarrow \infty} \frac{1}{\lambda_N} d(\alpha_1 \parallel \alpha_2) &= \limsup_{N \rightarrow \infty} \frac{1}{\lambda_N} \alpha_1 \log \frac{\alpha_1}{\alpha_2} + (1 - \alpha_1) \log \frac{1 - \alpha_1}{1 - \alpha_2} \\ &= \limsup_{N \rightarrow \infty} \frac{1}{\lambda_N} \alpha_1 \log \frac{1}{\alpha_2}. \end{aligned}$$

Since $\frac{1}{\lambda_N} d(\alpha_1 \parallel \alpha_2)$ is bounded, so is the sequence $\frac{1}{\lambda_N} \alpha_1 \log \frac{1}{\alpha_2}$, and since α_1 is bounded away from 0, this implies that $\frac{1}{\lambda_N} \log \frac{1}{\alpha_2}$ is bounded as well. Using that $\lim_{N \rightarrow \infty} \alpha_1 = 1$ therefore yields the claim. \square

Lemma 5. *Let $M, N \in \mathbb{N}$ and let S be a discrete subset of the N -dimensional unit sphere with cardinality M . Then for G the law of the N -dimensional random variable \mathbf{Z} with i.i.d. standard Gaussian coordinates it holds*

$$E_{\mathbf{Z} \sim G} \max_{X' \in S} \langle X', \mathbf{Z} \rangle^2 = O(\log M).$$

Proof. It suffices to show that

$$E_{\mathbf{Z} \sim G} \max_{X' \in S} \langle X', \mathbf{Z} \rangle^2 \mathbb{1} \left(\max_{X' \in S} \langle X', \mathbf{Z} \rangle^2 \geq 2 \log M \right) = O(1).$$

or

$$\int_0^\infty G \left(\max_{X' \in S} \langle X', \mathbf{Z} \rangle^2 \geq 2 \log M + t \right) dt = O(1).$$

Using a union bound argument and the fact that for all $X' \in S$ the quantity $\langle X', \mathbf{Z} \rangle$ follows a standard Gaussian distribution, we have for all $t \geq 0$,

$$G \left(\max_{X' \in S} \langle X', \mathbf{Z} \rangle^2 \geq 2 \log M + t \right) \leq M \exp \left(-\log M - \frac{t}{2} \right) = \exp \left(-\frac{t}{2} \right).$$

Hence

$$\int_0^\infty G \left(\max_{X' \in S} \langle X', \mathbf{Z} \rangle^2 \geq 2 \log M + t \right) dt \leq \int_0^\infty \exp \left(-\frac{t}{2} \right) dt = O(1),$$

as we wanted. \square

Lemma 6. *Suppose that $k = o(p)$ and the prior $\tilde{\mathbb{P}}_p$ is the uniform distribution on all the k -sparse vectors with elements either 0 or $1/\sqrt{k}$. Then for any $t \in [0, 1]$ it holds*

$$\lim_{p \rightarrow +\infty} \frac{1}{k \log \frac{p}{k}} \log \tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle \geq t] = -t.$$

Proof. First note that the claim follows immediately when $t = 1$ as when $k = o(p)$, the distribution $\tilde{\mathbb{P}}_p$ is distribution over a discrete subset of the unit sphere of cardinality $(1 + o(1))k \log \frac{p}{k}$. Similarly, since for all v, v' in the support of $\tilde{\mathbb{P}}_p$ it holds $\langle v, v' \rangle \geq 0$, the claim also follows straightforwardly for $t = 0$. For the rest of the proof we assume $t \in (0, 1)$.

We first show that the limit superior is bounded above by $-t$. The distribution of the rescaled overlap $k\langle \mathbf{x}, \mathbf{x}' \rangle = \langle \sqrt{k}\mathbf{x}, \sqrt{k}\mathbf{x}' \rangle$ follows the Hypergeometric distribution $\text{Hyp}(p, k, k)$ with probability mass function $p(s) = \binom{k}{s} \binom{p-k}{k-s} / \binom{p}{k}$, for $s = 0, \dots, k$. Therefore for a fixed $t \in (0, 1]$,

$$\tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle \geq t] = \sum_{s=\lceil tk \rceil}^k p(s). \quad (12)$$

Now for any $s \geq \lceil tk \rceil$ it holds

$$\frac{p(s+1)}{p(s)} = \frac{\binom{k}{s+1} \binom{p-k}{k-s-1}}{\binom{k}{s} \binom{p-k}{k-s}} = \frac{(k-s)^2}{(s+1)(p-2k+s+1)}.$$

Using that $k = o(p)$ and $s \geq tk$ we conclude that for sufficiently large p and all $s \geq \lceil tk \rceil$ it holds

$$\frac{p(s+1)}{p(s)} \leq 2 \frac{k}{tp} < \frac{1}{2}.$$

or by telescopic product,

$$\frac{p(s)}{p(\lceil tk \rceil)} \leq \frac{1}{2^{s-\lceil tk \rceil}}.$$

Hence, using (12) we have for large enough values of p ,

$$\tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle \geq t] \leq \sum_{s=\lceil tk \rceil}^k p(\lceil tk \rceil) \frac{1}{2^{s-\lceil tk \rceil}} \leq 2p(\lceil tk \rceil). \quad (13)$$

We have

$$p(\lceil tk \rceil) = \binom{k}{\lceil tk \rceil} \binom{p-k}{k-\lceil tk \rceil} / \binom{p}{k}$$

and combining with the elementary bound $\log \binom{m_1}{m_2} = m_2 \log \left(\frac{em_1}{m_2} \right) + O(m_2)$, for $m_1 \leq m_k$, we obtain

$$\begin{aligned} \log p(\lceil tk \rceil) &= tk \log \frac{1}{t} + (1-t)k \log \frac{p-k}{(1-t)k} - k \log \frac{p}{k} + O(k) \\ &= -tk \log \frac{p}{k} + O(k), \end{aligned} \quad (14)$$

where in the second step we have used that, for fixed $t \in (0, 1)$, if $k = o(p)$, then

$$\log \frac{p-k}{(1-t)k} = \log \frac{p}{k} + O(1).$$

We therefore conclude

$$\log \tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle \geq t] \leq \log p(\lceil tk \rceil) = -tk \log \frac{p}{k} + O(k). \quad (15)$$

Using the fact that $k = o(p)$ completes the proof of the upper bound.

We now prove the lower bound. By (12),

$$\tilde{\mathbb{P}}_p^{\otimes 2}[\langle \mathbf{x}, \mathbf{x}' \rangle \geq t] \geq p(\lceil tk \rceil),$$

and combining this with (14) yields the claim. \square