
Supplemental Material of "Triad Constraints for Learning Causal Structure of Latent Variables"

A Proof of the Theorems and Propositions

We begin with the Darmois-Skitovitch Theorem [Kagan *et al.*, 1973].

Darmois-Skitovitch Theorem (D-S Theorem): Define two random variables X_1 and X_2 as linear combinations of independent random variables $n_i (i = 1, \dots, q)$:

$$X_1 = \sum_{i=1}^q \alpha_i n_i, X_2 = \sum_{i=1}^q \beta_i n_i.$$

Then, if X_1 and X_2 are independent, all variables n_j for which $\alpha_j \beta_j \neq 0$ are Gaussian. In other words, if there exists a non-Gaussian n_j for which $\alpha_j \beta_j \neq 0$, X_1 and X_2 are dependent.

A.1 Proof of the Theorem 1

Theorem 1. Let L_a and L_b be two directed connected latent variables without confounders and let $\{X_i\}$ and $\{X_j, X_k\}$ be their children, respectively. Then if $\{X_i, X_j\}$ and X_k violate the Triad constraint, $L_a \rightarrow L_b$ holds. In other words, if the Triad condition is violated and the latent variables have no confounders, then the latent variable of the reference variable is a child of the other latent variable.

Proof. For L_a and L_b , there are two possible causal relations, $L_a \rightarrow L_b$ and $L_b \rightarrow L_a$, corresponding to the causal structure (a) and (b) in Figure 1.

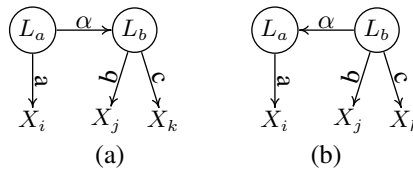


Figure 1: Identification of causal direction between two latent variables based on Triad constraints

In the non-trivial case, the causal strengths α, a, b and c are not equal to 0.

As the variables strictly follow linear assumption, for the structure (a), we obtain

$$\begin{aligned} L_a &= \varepsilon_{L_a}, \\ L_b &= \alpha L_a + \varepsilon_{L_b} = \alpha \varepsilon_{L_a} + \varepsilon_{L_b}, \end{aligned} \tag{1}$$

$$\begin{aligned} X_i &= a L_a + \varepsilon_{X_i} = a \varepsilon_{L_a} + \varepsilon_{X_i}, \\ X_j &= b L_b + \varepsilon_{X_j} = b \alpha \varepsilon_{L_a} + b \varepsilon_{L_b} + \varepsilon_{X_j}, \\ X_k &= c L_b + \varepsilon_{X_k} = c \alpha \varepsilon_{L_a} + c \varepsilon_{L_b} + \varepsilon_{X_k}. \end{aligned} \tag{2}$$

Then $E_{(i,j|k)}$ is as follows:

$$\begin{aligned}
E_{(i,j|k)} &= X_i - \frac{\text{Cov}(X_i, X_k)}{\text{Cov}(X_j, X_k)} \cdot X_j \\
&= a\varepsilon_{L_a} + \varepsilon_{X_i} - t \cdot (b\alpha\varepsilon_{L_a} + b\varepsilon_{L_b} + \varepsilon_{X_j}), \\
&= (a - \alpha tb)\varepsilon_{L_a} + \varepsilon_{X_i} - tb\varepsilon_{L_b} - t\varepsilon_{X_j},
\end{aligned} \tag{3}$$

where $t = \frac{\alpha ac \text{Var}(\varepsilon_{L_a})}{(\alpha)^2 bc \text{Var}(\varepsilon_{L_a}) + bc \text{Var}(\varepsilon_{L_b})} \neq 0$.

For $\{X_i, X_j\}$ and X_k , based on Equation (3) and Equation (2), we can find that they have one common non-Gaussian noise ε_{L_b} and $tb \neq 0, c \neq 0$. Hence, by D-S theorem, we have $E_{(i,j|k)} \not\perp\!\!\!\perp X_k$, i.e., $\{X_i, X_j\}$ and $\{X_k\}$ violate the Triad constraint.

Similarly for the structure (b), we have

$$\begin{aligned}
L_b &= \varepsilon_{L_b}, \\
L_a &= \alpha L_b + \varepsilon_{L_a} = \alpha\varepsilon_{L_b} + \varepsilon_{L_a},
\end{aligned} \tag{4}$$

$$\begin{aligned}
X_i &= aL_a + \varepsilon_{X_i} = a\alpha\varepsilon_{L_b} + a\varepsilon_{L_a} + \varepsilon_{X_i}, \\
X_j &= bL_b + \varepsilon_{X_j} = b\varepsilon_{L_b} + \varepsilon_{X_j}, \\
X_k &= cL_b + \varepsilon_{X_k} = c\varepsilon_{L_b} + \varepsilon_{X_k}.
\end{aligned} \tag{5}$$

Then the pseudo residual $E_{(i,j|k)}$ is as below.

$$\begin{aligned}
E_{(i,j|k)} &= X_i - \frac{\text{Cov}(X_i, X_k)}{\text{Cov}(X_j, X_k)} \cdot X_j \\
&= a\alpha\varepsilon_{L_b} + a\varepsilon_{L_a} + \varepsilon_{X_i} - \frac{a\alpha}{b} \cdot (b\varepsilon_{L_b} + \varepsilon_{X_j}) \\
&= a\varepsilon_{L_a} + \varepsilon_{X_i} - \frac{a\alpha}{b} \varepsilon_{X_j}.
\end{aligned} \tag{6}$$

For $\{X_i, X_j\}$ and X_k , based on Equation (6) and Equation (5), we find that there is no common non-Gaussian, independent component shared by $E_{(i,j|k)}$ and X_k . According to D-S Theorem, we reach the result that $E_{(i,j|k)} \perp\!\!\!\perp X_k$, i.e., $\{X_i, X_j\}$ and $\{X_k\}$ satisfy the Triad constraint. This finishes the proof. \square

A.2 Proof of the Theorem 2

Theorem 2. Let S be a set of correlated variables. If $\forall X_i, X_j \in S$ and $\forall X_k \in \mathbf{X} \setminus S$, $\{X_i, X_j\}$ and X_k satisfy the Triad constraints, then S is a cluster.

Proof. The proof is done by contradiction. Assume S is not a cluster, the elements in S must have at least two different parental latent variables. Without loss of generality, let L_a and L_b be the two latent variables, and let their children be $\{X_i, X_j\}$ and $\{X_k, X_l\}$, respectively. There are two cases to consider.

Case 1). When there is a causal relationship between L_a and L_b , e.g., L_a is the ancestor of L_b . We know that X_k contains the noise ε_{L_b} while X_i and X_j do not. Since S is a correlated variable set, then X_i, X_j , and X_k are related, i.e., $\text{Cov}(X_i, X_j) \neq 0$ and $\text{Cov}(X_k, X_j) \neq 0$. By $E_{(i,k|j)} = X_i - \frac{\text{Cov}(X_i, X_k)}{\text{Cov}(X_k, X_j)} \cdot X_k$, we obtain that $E_{(i,k|j)}$ must contain ε_{L_b} . According to D-S Theorem, $E_{(i,k|j)} \not\perp\!\!\!\perp X_j$, i.e., $\{X_i, X_k\}$ and X_j violate Triad constraint, which contradicts the original assumption.

Case 2). When there is no causal relationship between L_a and L_b . Since S is a correlated variable set, we know $L_a \not\perp\!\!\!\perp L_b$. Therefore, L_a and L_b have at least one common ancestor. This structure shows that $\{X_i, X_j\}$ and $\{X_k, X_l\}$ contain respective latent variable noises, i.e., ε_{L_a} and ε_{L_b} . Therefore, $E_{(i,k|j)}$ and $E_{(i,k|l)}$ both contain ε_{L_a} and ε_{L_b} . Based on the D-S Theorem, $E_{(i,k|j)} \not\perp\!\!\!\perp X_j$ and $E_{(i,k|l)} \not\perp\!\!\!\perp X_l$, i.e., $\{X_i, X_k\}$ and X_j violate Triad constraint, and so do $\{X_i, X_k\}$ and X_l . This contradicts the original assumption.

Based on the above analysis, Theorem 2 holds. \square

A.3 Proof of the Proposition 1

Proposition 1. Let C_1 and C_2 be two clusters. If C_1 and C_2 are overlapping, C_1 and C_2 share a same latent parent.

Proof. Since C_1 and C_2 are two clusters, then the elements in C_1 have only one common latent variable. Without loss of generality, we let L_1 denote the parental latent variable of C_1 . Similarly, L_2 denotes the parental latent variable of C_2 . Since C_1 and C_2 are overlapping, then they have at least one shared element. Let X_i denote the shared element of C_1 . then X_i has two latent parents L_1 and L_2 , which contradicts with Theorem 1. This finishes the proof. \square

A.4 Proof of the Proposition 2

Proposition 2. Given a latent variable L_r and its two children $\{V_i, V_j\}$, L_r is a root latent variable if and only if $E_{(i,j|k)} \perp\!\!\!\perp V_k$ holds for all the V_k , where V_k is a child of any other latent variables.

Proof. (i) " \Rightarrow " Since L_r is a root latent variable, there is no confounder between L_r and another latent variables. Based on Theorem 1, we reach the conclusion that $E_{(i,j|k)} \perp\!\!\!\perp V_k$ holds for all V_k .

(ii) " \Leftarrow " This part of proof is proved by contradiction. Assume L_r is not a root variable, L_r has at least one parent. Considering the following two cases. Case 1: L_r has only one parent. Let L_s denote the parent of L_r and $\{V_k\}$ denote a child of L_s . Based on Theorem 1, $E_{(i,k|j)}$ is not independent of V_j . Case 2: L_r has more than one parent, e.g., L_s and L_t , where their children variables are $\{V_k\}$ and $\{V_l\}$ respectively. If $L_s \perp\!\!\!\perp L_t$, based on the Theorem 1, one can find that $E_{(i,k|j)}$ is not independent of V_j and that $E_{(i,l|j)}$ is not independent of V_j . If $L_s \not\perp\!\!\!\perp L_t$ and $L_s \rightarrow L_t$, due to linear assumption, $E_{(i,p|j)}$ always contains noise ε_{L_t} and V_p also contains this noise variable. According to D-S theorem, $E_{(i,p|j)}$ is not independent of V_p ; Similarly, when $L_s \not\perp\!\!\!\perp L_t$ and $L_s \leftarrow L_t$, $E_{(i,k|j)}$ is not independent of V_k . The results of Case 1 and Case 2 show that the assumption does not hold, and L_r is a root variable. This finishes the proof. \square

B The correctness of Phase 2

We illustrate the correctness of Phase 2, especially the procedure of learning causal order of latent variables recursively, with the example in our paper.

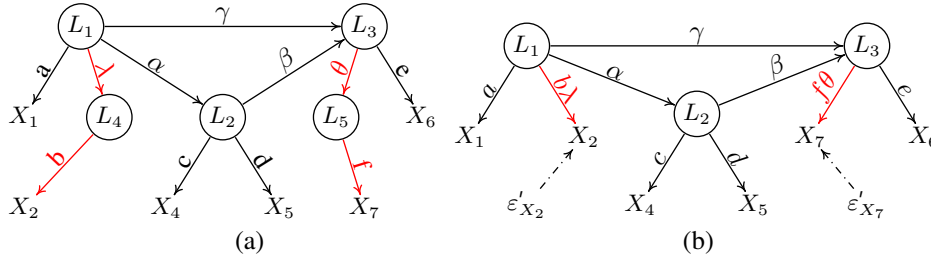


Figure 2: The considered structure in our paper, where (a) is the ground truth graph, (b) is the equivalent graph of (a)

In phase 1, we have get three clusters $\{\{X_1, L_4\}, \{X_4, X_5\}, \{X_6, L_5\}\}$ and their latent variables are L_1, L_2 and L_3 , respectively. Next, we will show the process of Phase 2 step by step.

Note that, although we can not get the observed values of latent variables L_4 and L_5 , we can use the values of their pure child as surrogates (This is because linear causal models are transitive). Here, we use X_2 and X_7 to replace L_4 and L_5 , respectively, where $\varepsilon'_{X_2} = b\lambda\varepsilon_{L_4} + \varepsilon_{X_2}$ and $\varepsilon'_{X_7} = f\varepsilon_{L_5} + \varepsilon_{X_7}$ (See Figure 3).

We obtain,

$$\begin{aligned}
X_1 &= aL_1 + \varepsilon_{X_1}, \\
X_2 &= b\lambda L_1 + \varepsilon'_{X_2}, \\
X_4 &= cL_2 + \varepsilon_{X_4} = c(\alpha L_1 + \varepsilon_{L_2}) + \varepsilon_{X_4} = cL'_2 + c\alpha L_1 + \varepsilon_{X_4}, \\
X_5 &= dL_2 + \varepsilon_{X_5} = d(\alpha L_1 + \varepsilon_{L_2}) + \varepsilon_{X_5} = dL'_2 + d\alpha L_1 + \varepsilon_{X_5},
\end{aligned} \tag{7}$$

$$\begin{aligned}
X_6 &= eL_3 + \varepsilon_{X_6} = e((\gamma + \alpha\beta)L_1 + \beta\varepsilon_{L_2} + \varepsilon_{L_3}) + \varepsilon_{X_6} \\
&= eL'_3 + e(\gamma + \alpha\beta)L_1 + \varepsilon_{X_6}, \\
X_7 &= f\theta L_3 + \varepsilon'_{X_7} = f\theta((\gamma + \alpha\beta)L_1 + \beta\varepsilon_{L_2} + \varepsilon_{L_3}) + \varepsilon'_{X_7} \\
&= f\theta L'_3 + f\theta(\gamma + \alpha\beta)L_1 + \varepsilon'_{X_7},
\end{aligned} \tag{8}$$

We can then learn the causal structure of latent variable step by step in the following way.

- First, according to the Proposition 2, we know that $\{X_1, X_2\}$ correspond to the root latent cause.
- Next, we are ready to find the causal direction between the latent variables L_2 and L_3 . Let us consider the pseudo residual variables $E_{(i,k|l)}$ instead of X_i , where $i \in \{4, 5, 6, 7\}$, $k \in \{1, 2\}$, $l \in \{1, 2\}$ and $k \neq l$. For convenience, we let $L'_2 := \varepsilon_{L_2}$ and $L'_3 := \beta\varepsilon_{L_2} + \varepsilon_{L_3}$. Then, we update the rest of variables by $\{X_1, X_2\}$,

$$E_{(4,1|2)} = X_4 - \frac{\text{Cov}(X_4, X_2)}{\text{Cov}(X_1, X_2)} \cdot X_1 = cL'_2 + \varepsilon_{X_4} - \frac{c\alpha}{a} \cdot \varepsilon_{X_1} \tag{9}$$

$$\begin{aligned}
E_{(6,1|2)} &= e((\gamma + \alpha\beta)L_1 + \beta\varepsilon_{L_2} + \varepsilon_{L_3}) + \varepsilon_{X_6} - \frac{e((\gamma + \alpha\beta))}{a}(aL_1 + \varepsilon_{X_1}) \\
&= eL'_3 + \varepsilon_{X_6} - \frac{e((\gamma + \alpha\beta))}{a}\varepsilon_{X_1}.
\end{aligned} \tag{10}$$

- Finally, combining equations (7-10) and Theorem 1, one can see that $\{E_{(4,1|2)}, E_{(6,1|2)}\}$ and X_5 satisfy the Triad constraint and that $\{E_{(4,1|2)}, E_{(6,1|2)}\}$ and X_7 violate it. Figure 4 gives the graphical representations of the relationships among those variables, from which the above conclusion can be immediately seen. Thanks to this asymmetry, we know that the latent variable L_2 , which generated $\{X_4, X_5\}$, is a cause of L_3 , which generated $\{X_6, X_7\}$.

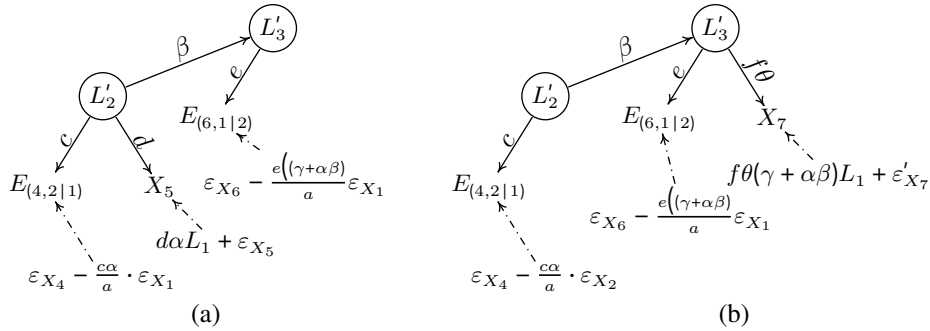


Figure 3: Using the Triad constraint to determine the direction between L_2 and L_3 . (a) $\{E_{(4,1|2)}, E_{(6,1|2)}\}$ and X_5 satisfy the Triad constraint. (b) $\{E_{(4,1|2)}, E_{(6,1|2)}\}$ and X_7 violate the Triad constraint. The influences of noise terms are shown by dashed lines; note that the noise terms are mutually independent in each case.

References

Abram M Kagan, Calyampudi Radhakrishna Rao, and Yuriy Vladimirovich Linnik. Characterization problems in mathematical statistics. 1973.