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# Supplementary materials for latent Bayesian melding for integrating individual and population models

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**Theorem 1.** *If  $E_{p_S(S)} \left[ \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right] < \infty$ , then a constant  $c_\alpha < \infty$  exists such that  $\int \tilde{p}_{S,\xi}(S, \xi) d\xi dS = 1$ , for any fixed  $\alpha \in [0, 1]$ .*

*Proof.* If  $\alpha = 1$ , then  $c_\alpha = 1$ . If  $\alpha = 0$ , then,

$$\begin{aligned} \int p_S(S) \left( \frac{p_\tau(f(S)|\xi)p(\xi)}{p_\tau^*(f(S))} \right) d\xi dS &= \int p_S(S) \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \frac{p_\tau(f(S)|\xi)p(\xi)}{p_\tau(f(S))} d\xi dS \\ &= \int p_S(S) \frac{p_\tau(f(S))}{p_\tau^*(f(S))} dS < \infty \end{aligned}$$

Now we look at  $\alpha \in (0, 1)$ . Firstly, for  $x > 0$ , if  $\alpha \in (0, 1)$ , then  $g(x) = x^{1-\alpha}$  is a concave function, because  $g(x)'' = -\alpha(1-\alpha)x^{-\alpha-1} < 0$ . Similarly,  $x^\alpha$  is also a concave function. Then we have

$$\begin{aligned} &\int p_S(S) \left( \frac{p_\tau(f(S)|\xi)p(\xi)}{p_\tau^*(f(S))} \right)^{1-\alpha} d\xi dS \\ &= \int p_S(S) \left( \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right)^{1-\alpha} E_{p(\xi|f(S))} \left[ \left( \frac{p_\tau(f(S))}{p_\tau(f(S)|\xi)p(\xi)} \right)^\alpha \right] dS \\ &\leq \int p_S(S) \left( \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right)^{1-\alpha} \left[ E_{p(\xi|f(S))} \left( \frac{p_\tau(f(S))}{p_\tau(f(S)|\xi)p(\xi)} \right) \right]^\alpha dS \\ &= \int p_S(S) \left( \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right)^{1-\alpha} dS \\ &\leq \left[ E_{p_S(S)} \left( \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right) \right]^{1-\alpha} \\ &< \infty \end{aligned}$$

where Jensen's inequality has been applied twice. Therefore,  $c_\alpha < \infty$  exists, satisfying  $\int \tilde{p}_{\xi,S}(\xi, S) d\xi dS = 1$ . □

**Theorem 2.** If  $\lim_{\delta \rightarrow 0} p_\delta(\tau) = p_\tau^*(\tau)$ , and  $g_\delta(\tau)$  has bounded derivatives in any order, then  $\lim_{\delta \rightarrow 0} \int p_\delta(\tau|S)g_\delta(\tau)d\tau = g(f(S))$ .

*Proof.* Since  $\tau$  is an Uniform distribution on  $[f(S) - \delta, f(S) + \delta]$  conditional on  $S$  and  $\delta$ , we could draw  $N$  samples for  $\tau$  such that  $\tau_i = f(S) + (2u_i - 1)\delta$  where  $u_i$  is a sample drawn from the standard Uniform distribution, where  $i = 1, 2, \dots, N$ . By using Monte Carlo approximation and Taylor's expansion, we have

$$\begin{aligned}
& \lim_{\delta \rightarrow 0} \int p_\delta(\tau|S)g_\delta(\tau)d\tau \\
&= \lim_{\delta \rightarrow 0} \int p(\tau \in (f(S) - \delta, f(S) + \delta))g_\delta(\tau)d\tau \\
&= \lim_{\delta \rightarrow 0} \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N g_\delta(f(S) + (2u_i - 1)\delta) \\
&= \lim_{\delta \rightarrow 0} \lim_{N \rightarrow \infty} \left\{ g_\delta(f(S)) + g'_\delta(f(S))\delta \frac{1}{N} \sum_{i=1}^N (2u_i - 1) + \frac{1}{2!}g''_\delta(f(S))\delta^2 \frac{1}{N} \sum_{i=1}^N (2u_i - 1)^2 + \dots \right\} \\
&= \lim_{\delta \rightarrow 0} g_\delta(f(S)) \\
&= g(f(S)).
\end{aligned}$$

This holds, since  $|2u_i - 1|^k \leq 1$  ( $k = 1, 2, 3, \dots$ ) and  $\lim_{N \rightarrow \infty} \sum_{i=1}^N \frac{1}{N} |2u_i - 1|^k \leq 1$ ,  $\sum_{i=1}^N \frac{1}{N} (2u_i - 1)^k$  converges absolutely when  $N \rightarrow \infty$ .  $\square$