Tracking Most Significant Shifts in Nonparametric Contextual Bandits

Anonymous Author(s) Affiliation Address email

Abstract

1	We study nonparametric contextual bandits where Lipschitz mean reward functions
2	may change over time. We first establish the minimax dynamic regret rate in
3	this less understood setting in terms of number of changes L and total-variation
4	V, both capturing all changes in distribution over context space, and argue that
5	state-of-the-art procedures are suboptimal in this setting.
6	Next, we tend to the question of an <i>adaptivity</i> for this setting, i.e. achieving the
7	minimax rate without knowledge of L or V. Quite importantly, we posit that the
8	bandit problem, viewed local at a given context X_t , should not be affected by
9	reward changes in other parts of context space \mathcal{X} . We therefore propose a notion of
10	change that better accounts for locality, and thus counts significantly less changes
11	than L and V. Our main result is to show that this more strict notion of change,
12	which we term experienced significant shifts, can in fact be adapted to. As in
13	previous work on non-stationary MAB (Suk and Kpotufe, 2022), not only do our
14	results capture changes only at the experienced contexts x , but also only the most
15	significant in terms of changes in mean rewards (e.g., only count severe best-arm
16	changes at x).

17 **1 Introduction**

¹⁸ Contextual bandits model sequential decision making problems where the reward of a chosen action ¹⁹ depends on an observed context X_t at time t, e.g., a consumer's profile, a medical patient's history. ²⁰ The goal is to maximize the total rewards over time of chosen actions, as informed by seen contexts. ²¹ As such, one suitable measure of performance is that of *dynamic regret*, which compares earned ²² rewards to a time-varying oracle maximizing mean rewards at X_t . While it is often assumed in the ²³ bulk of works in this setting that rewards distributions remain stationary over time, it is understood ²⁴ that in practice, environmental changes induce nontrivial changes in rewards.

In fact, the problem of non-stationary environments has received a surge of attention in the simpler 25 non-contextual Multi-Arm-Bandits (MAB) setting, while the more challenging contextual case 26 remains ill-understood. In particular in the contextual case, some recent works of Wu et al. [2018], 27 Luo et al. [2018], Chen et al. [2019], Wei and Luo [2021] consider parametric settings, i.e. where 28 reward functions belong to fixed parametric family, and show that one may achieve rates adaptive to 29 an unknown number of L of shifts in rewards or to a notion of total-variation V, both acccounting 30 for all changes over time and context space. Instead here, we consider a much larger class of reward 31 functions, namely Lipschitz rewards, corresponding to the natural assumption that closeby contexts 32 33 have similar rewards even as reward distributions change.

As a first result for this nonparametric setting, we establish some minimax lower-bounds as a baseline in terms of either L or V, and argue that state-of-the-art procedures for the parametric case—extended to the class of Lipschitz functions—do not achieve these baselines.

Submitted to 37th Conference on Neural Information Processing Systems (NeurIPS 2023). Do not distribute.

We then turn attention to whether such baselines may be achieved *adaptively*, i.e., without knowledge 37 of L or V. The answer as we show is affirmative, and more importantly, some much weaker notions 38 of change may be adapted to; for intuition, while L or V accounts for any change at any time over 39 the context space (say \mathcal{X}), it may be that all changes are relegated to parts of the space irrelevant 40 to observed contexts X_t at the time they are played. For instance, suppose at time t, we observe 41 $X_t = x_0$, then it may not make sense to count changes that happen at some other x_1 far from x_0 , or 42 changes that happened at x_0 itself but far back in time. 43 We therefore propose a new parameterization of change, termed experienced significant shifts that 44

we difference propose a new parameterization of change, termed experienced significant single that better accounts for the locality of changes in time and space, and as such may register much less changes than either L or V. As a sanity check, we show that an oracle policy which restarts only at experienced significant shifts can attain enhanced regret rates in terms of the number $\tilde{L} =$ $\tilde{L}(X_1, \ldots, X_T)$ of such experienced shifts (Proposition 2), a rate always no worse that the baseline we first established in terms of L and V.

Our main result is to show that *experienced significant shifts* can be adapted to (Theorem 3), i.e., 50 with no prior knowledge of such shifts. Importantly, the result holds in both stochastic environments, 51 and in (oblivious) adversarial ones with no change to our notion, algorithmic approach, nor analysis. 52 Furthermore, similar to recent advances in the non-contextual case [Abbasi-Yadkori et al., 2022, Suk 53 and Kpotufe, 2022], an *experienced shift* is only triggered under *severe changes* such as changes of 54 best arms locally at a context X_t . An added difficulty in the contextual case is that we cannot hope to 55 observe rewards for a given arm (action) repeatedly at X_t as the context may only appear once, and 56 have to rely on careful chosen nearby points to identify unknown shifts in reward at X_t . 57

58 1.1 Other Related Work

Nonparametric Contextual Bandits. The stationary bandits with covariates (where rewards and 59 contexts follow a joint distribution) was first introduced in a one-armed bandit problem [Woodroofe, 60 1979, Sarkar, 1991], with the nonparametric model first studied by Yang et al. [2002]. Minimax 61 regret rates, based on a margin condition, were first established for the two-armed bandit in Rigollet 62 and Zeevi [2010] and generalized to any finite number of arms in Perchet and Rigollet [2013], with 63 further insights thereafter [Qian and Yang, 2016a,b, Reeve et al., 2018, Guan and Jiang, 2018, Gur 64 et al., 2022, Hu et al., 2020, Arya and Yang, 2020, Suk and Kpotufe, 2021, Cai et al., 2022]. However, 65 the mentioned works all assume a stationary distribution of rewards over contexts. Blanchard et al. 66 67 [2023] studies non-stationary nonparametric contextual bandits, but in the much-different context of universal learning, concerning when sublinear regret is achievable asymptotically. 68

⁶⁹ Lipschitz contextual bandits appears as part of studies on broader infinite-armed settings [Lu et al.,

2009, Krishnamurthy et al., 2019]. Related, Slivkins [2014] allows for non-stationary (i.e., obliviously

⁷¹ adversarial) environments, but only studies regret to the (per-context) best arm in hindsight.

72 *Realizable contextual bandits* posits that the regression function capturing mean rewards in contexts

⁷³ lies in some known class of regressors \mathcal{F} , over which one can do empirical risk minimization [Foster

et al., 2018, Foster and Rakhlin, 2020, Simchi-Levi and Xu, 2021]. While this setting can recover
 Lipschitz contextual bandits, the only result on non-stationary guarantees to our knowledge is Wei

⁷⁶ and Luo [2021], which yields suboptimal dynamic regret (see Table 1).

Non-Stationary Bandits and RL. In the simpler non-contextual bandits, changing reward distribu-77 tions (a.k.a. switching bandits) was introduced in Garivier and Moulines [2011] and further explored 78 with various assumptions and formulations [Besbes et al., 2019, Karnin and Anava, 2016, Allesiardo 79 et al., 2017, Liu et al., 2018, Wei and Srivatsva, 2018, Besson et al., 2022, Cao et al., 2019, Mukherjee 80 and Maillard, 2019]. While these earlier works focused on algorithmic design assuming knowledge 81 82 of non-stationarity, such a strong assumption was removed via the *adaptive* procedures of Auer et al. [2019], Chen et al. [2019]. In followup works, Abbasi-Yadkori et al. [2022], Suk and Kpotufe [2022] 83 show that tighter dynamic regret rates are possible, scaling only with severe changes in best arm. 84 The ideas from non-stationary MAB were extended to various contextual bandit settings by Wu et al. 85

Ine ideas from non-stationary MAB were extended to various contextual bandit settings by Wu et al.
 [2018] (for linear mean rewards in contexts), Luo et al. [2018], Chen et al. [2019] (for finite policy)

classes), and Wei and Luo [2021] (for realizable mean reward functions).

⁸⁸ There have also been extensions of these ideas to various reinforcement learning setups [Jaksch

et al., 2010, Gajane et al., 2018, Ortner et al., 2020, Cheung et al., 2020, Fei et al., 2020, Mao et al.,

2021, Zhou et al., 2022, Touati and Vincent, 2020, Domingues et al., 2021, Chi Cheung et al., 2019, 90

- Domingues et al., 2021, Ding and Lavaei, 2023, Wei and Luo, 2021, Lykouris et al., 2021, Wei et al., 91
- 2022, Chen and Luo, 2022]. Among these works, only Domingues et al. [2021] can recover Lipschitz 92
- contextaul bandits, whereupon we find their dynamic regret bounds are suboptimal (see Table 1). 93

Again, the typical aim of aforementioned works on contextual bandits or RL is to minimize a notion 94 of dynamic regret in terms of the number of changes L or total-variation V. As such, regardless of 95

setting, known guarantees in said works do not involve tighter notions of experienced non-stationarity. 96

Problem Formulation 2 97

2.1 Contextual Bandits with Changing Rewards 98

Preliminaries. We assume a finite set of arms $[K] \doteq \{1, 2, ..., K\}$. Let $Y_t \in [0, 1]^K$ denote the vector of rewards for arms $a \in [K]$ at round $t \in [T]$ (horizon T), and X_t the observed context 99 100 at that round, lying in $\mathcal{X} \doteq [0,1]^d$, which have joint distribution $(X_t, Y_t) \sim \mathcal{D}_t$. We let $\mathbf{X}_t \doteq$ 101 $\{X_s\}_{s \le t}, \mathbf{Y}_t \doteq \{Y_s\}_{s \le t}$ denote the observed contexts and (observed and unobserved) rewards from 102 rounds 1 to t. In our setting, an oblivious adversary decides a sequence of (independent) distributions 103 on $\{(X_t, Y_t)\}_{t \in [T]}$ before play. 104

Notation. The reward function $f_t : \mathcal{X} \to [0,1]^K$ is $f_t^a(x) \doteq \mathbb{E}[Y_t^a|X_t = x], a \in [K]$, and captures the mean rewards of arm a at context x and time t. 105 106

A *policy* chooses actions at each round t, based on observed contexts (up to round t) and passed 107 rewards, whereby at each round t only the reward Y_t^a of the chosen action a is revealed. Formally: 108

Definition 1 (Policy). A policy $\pi \doteq {\pi_t}_{t\in\mathbb{N}}$ is a random sequence of functions $\pi_t : \mathcal{X}^t \times [K]^{t-1} \times [K]^{t-1}$ 109

 $[0,1]^{t-1} \to [K]$. In the case of a randomized policy, i.e., where π_t in fact maps to distributions 110

on [K], In an abuse of notation, in the context of a sequence of observations till round t, we'll let 111

 $\pi_t \in [K]$ denote the (possibly random) action chosen at round t. 112

The performance of a policy is evaluated using the dynamic regret, defined as follows: 113

Definition 2. Fix a context sequence X_T . Define the dynamic regret of a policy π , as 114

$$R_T(\pi, \mathbf{X}_T) \doteq \sum_{t=1}^T \max_{a \in [K]} f_t^a(X_t) - f_t^{\pi_t}(X_t).$$

Thus, we seek a policy π that minimizes $\mathbb{E}[R_T(\pi, \mathbf{X}_T)]$ where the expectation is over $\mathbf{X}_T, \mathbf{Y}_T$, and 115 any randomness in π . 116

Notation. As much of our analysis focuses on the gaps in mean rewards between arms at observed 117

118

contexts X_t , the following notation will serve useful. Let $\delta_t(a', a) = f_t^{a'}(X_t) - f_t^a(X_t)$ denote the **relative gap** of arms a to a' at round t at context X_t . Define the **worst gap** of arm a as $\delta_t(a) = \max_{a' \in [K]} \delta_t(a', a)$, corresponding to the instantaneous regret of playing a at round t and 119

120

context X_t . Thus, the dynamic regret can be written as $\sum_{t \in [T]} \mathbb{E}[\delta_t(\pi_t)]$. Additionally, let $\delta_t^{a',a}(x) \doteq \delta_t^{a',a}(x)$ 121

 $f_t^{a'}(x) - f_t^a(x)$ and $\delta_t^a(x) \doteq \max_{a' \in [K]} \delta_t^{a',a}(x)$ be the gap functions mapping $\mathcal{X} \to [0,1]$. 122

2.2 Nonparametric Setting 123

We assume, as in prior work on nonparametric contextual bandits [Rigollet and Zeevi, 2010, Perchet 124 and Rigollet, 2013, Slivkins, 2014, Reeve et al., 2018, Guan and Jiang, 2018, Suk and Kpotufe, 2021], 125 that the reward function is 1-Lipschitz. 126

- 127
 - **Assumption 1** (Lipschitz f_t). For all rounds $t \in \mathbb{N}$, $a \in [K]$ and $x, x' \in \mathcal{X}$,

$$|f_t^a(x) - f_t^a(x')| \le ||x - x'||_{\infty}.$$
(1)

For ease of presentation, we assume the contextual marginal distribution μ_X remains the same across 128

rounds. Furthermore, we make a standard strong density assumption on μ_X , which is typical in this 129

nonparametric setting [Audibert and Tsybakov, 2007, Perchet and Rigollet, 2013, Qian and Yang, 130

- 2016a,b, Gur et al., 2022, Hu et al., 2020, Arya and Yang, 2020, Cai et al., 2022]. This holds, e.g. if 131
- μ_X has a continuous Lebesgue density on $[0,1]^d$, and ensures good coverage of the context space. 132

Assumption 2 (Strong Density Condition). There exist $C_d, c_d > 0$ s.t. $\forall \ell_{\infty}$ balls $B \subset [0, 1]^d$ of diameter $r \in (0, 1]$:

$$C_d \cdot r^d \ge \mu_X(B) \ge c_d \cdot r^d. \tag{2}$$

Remark 1. We can in fact relax the above assumptions on context marginals so that $\mu_{X,t}(\cdot)$ is changing with time t and the above strong density assumption is satisfied with different constants $C_{d,t}, c_{d,t}$. Our procedures in the end will not require knowledge of any $C_{d,t}, c_{d,t}$.

138 2.3 Model Selection

- 139 A common algorithmic approach in nonparametric contextual bandits, starting from earlier work
- [Rigollet and Zeevi, 2010, Perchet and Rigollet, 2013], is to discretize or partition the context space
- 141 \mathcal{X} into *bins* where we can maintain local reward estimates. These bins have a natural hierarchical 142 tree structure which we first elaborate.

Definition 3 (Partition Tree). Let $\mathcal{R} \doteq \{2^{-i} : i \in \mathbb{N} \cup \{0\}\}$, and let $\mathcal{T}_r, r \in \mathcal{R}$ denote a regular partition of $[0, 1]^d$ into hypercubes (which we refer to as **bins**) of side length (a.k.a. bin size) r. We then define the dyadic **tree** $\mathcal{T} \doteq \{\mathcal{T}_r\}_{r \in \mathcal{R}}$, i.e., a hierarchy of nested partitions of $[0, 1]^d$. We will refer to the **level** r of T as the collection of bins in partition \mathcal{T}_r . The **parent** of a bin $B \in \mathcal{T}_r, r < 1$ is the bin $B' \in \mathcal{T}_{2r}$ containing B; **child**, **ancestor** and **descendant** relations follow naturally. The notation $\mathcal{T}_r(x)$ will then refer to the bin at level r containing x.

Note that, while in the above definition, \mathcal{T} has infinite levels $r \in \mathcal{R}$, at any round t in a procedure, we implicitly only operate on the subset of \mathcal{T} containing data.

Key in securing good regret is then finding the optimal level $r \in \mathcal{R}$ of discretization (balancing

regression bias and variance), which over n stationary rounds is known to be $\propto (K/n)^{\frac{1}{2+d}}$ [Rigollet

and Zeevi, 2010]. We introduce the following general notation, useful later in the approaching the

non-stationary problem, for associating the size of a level to an intervals of rounds.

155 Notation 1 (Level). For $n \in \mathbb{N} \cup \{0\}$, let r_n be the largest $2^{-m} \in \mathcal{R}$ such that $(K/n)^{\frac{1}{2+d}} \geq 2^{-m}$.

We use $\mathcal{T}_m, T_m(x)$ as shorthand to denote (respectively) the tree \mathcal{T}_r of level $r = r_m$ and the (unique) bin at level r_m containing x.

158 3 Results Overview

159 3.1 Minimax Lower Bounds Under Global Shifts

As a baseline, we start with some basic lower-bounds under the simplest parametrizations of changes in rewards which have appeared in the literature, namely a *global number of shifts*, and *total variation*.

Definition 4 (Global Number of Shifts). Let $L \doteq \sum_{t=2}^{T} \mathbf{1}\{\exists x \in \mathcal{X}, a \in [K] : f_t^a(x) \neq f_{t-1}^a(x)\}$ be the number of global shifts, *i.e.*, it counts every change in mean-reward overtime and over \mathcal{X} space.

Definition 5 (Total Variation). Define $V_T \doteq \sum_{t=2}^T \|\mathcal{D}_t - \mathcal{D}_{t-1}\|_{TV}$ where recall $\mathcal{D}_t \in \mathcal{X} \times [0, 1]^K$ is the joint distribution on context and rewards at time t.

We have the following initial result (for two-armed bandits) to serve as baseline for this study.

Theorem 1 (Dynamic Regret Lower Bound). Suppose there are K = 2 arms. For $V, L \in [0, T]$, let

- 168 $\mathcal{P}(V,L,T)$ be the family of joint distributions $\mathcal{D} \doteq \{\mathcal{D}_t\}_{t \in [T]}$ with either total variation $V_T \leq V$ or
- at most L global shifts. Then, there exists a constant c > 0 such that:

$$\sup_{\mathcal{D}\in\mathcal{P}(V,L,T)} \mathbb{E}_{\mathcal{D}}[R(\pi, X_T)] \ge c \left(T^{\frac{1+d}{2+d}} + T^{\frac{2+d}{3+d}} \cdot V^{\frac{1}{3+d}} \right) \wedge \left((L+1)^{\frac{1}{2+d}} T^{\frac{1+d}{2+d}} \right).$$
(3)

Remark 2. Note setting d = 0 in Theorem 1 recovers the established non-contextual minimax rate of $(\sqrt{T} + T^{2/3}V_T^{1/3}) \wedge \sqrt{(L+1) \cdot T}$.

Achievability of Miminimax Lower-Bound (3). We are interested in whether the rates of (3) are achievable, with, or without knowledge of relevant parameters. First, we note that no existing algorithm currently guarantees a rate that matches (3). See Table 1 for a rate comparison (details in Appendix A).

- In particular, the prior adaptive works [Chen et al., 2019, Wei and Luo, 2021] both rely on the
- approach of randomly scheduling *replays* of stationary algorithms to detect unknown non-stationarity.
- However, the scheduling rate is designed to safeguard against their parametric $\sqrt{LT} \wedge V_T^{1/3} T^{2/3}$
- regret rates and thus lead to suboptimal dependence on L and V_T .
- However, a simple back of the envelope calculation indicates that the rate in (3) may be attainable, at least given some distributional knowledge: a procedure restarting at each shift will incur regret, over

182 L equally spaced shifts,
$$(L+1) \cdot \left(\frac{T}{L+1}\right)^{\frac{1+d}{2+d}} \approx L^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}$$
.

As it turns out as we will show in the next section, (3) is indeed attainable, even adaptively; in fact, this is shown via a more optimistic problem parametrization as described next.

	Dynamic Regret Upper Bound
ADA-ILTCB [Chen et al., 2019]	$\left(L^{1/2} \cdot T^{\frac{1+d}{2+d}} \right) \wedge \left(V_T^{1/3} \cdot T^{\frac{2+d}{3+d} + \frac{d}{3(2+d)(3+d)}} \right)$
MASTER with FALCON [Wei and Luo, 2021]	$\left(L^{1/2} \cdot T^{\frac{1+d}{2+d}} \right) \wedge \left(V_T^{1/3} \cdot T^{\frac{2+d}{3+d} + \frac{d}{3(2+d)(3+d)}} \right)$
KeRNS [Domingues et al., 2021] (non-adaptive)	$V_T^{1/3} T^{\frac{2+d}{3+d} + O(1/d)}$
Minimax Lower-Bound	$\left(L^{\frac{1}{2+d}}T^{\frac{1+d}{2+d}}\right) \wedge \left(V_T^{\frac{1}{3+d}}T^{\frac{2+d}{3+d}}\right)$

Table 1: Existing dynamic Regret Upper-Bounds appear suboptimal in the Lipschitz setting.

185 3.2 A New Problem Parametrization: Experienced Significant Shifts.

- As discussed in Section 1, typical approaches in our setting discretize the context space \mathcal{X} into bins, each of which is treated as an MAB instance. At a high level, our new measure of non-stationarity
- will trigger an **experienced significant shift** when the observed context X_t arrives in a bin $B \in \mathcal{T}$
- where there has been a severe change in local best arm, *w.r.t. the observed data in that bin.*
- We first define a notion of significant regret for an arm $a \in [K]$ locally within a bin $B \in \mathcal{T}$. We say arm a incurs significant regret in bin B on interval I if:

$$\sum_{s \in I} \delta_s(a) \cdot \mathbf{1}\{X_s \in B\} \ge \sqrt{K \cdot n_B(I)} + r(B) \cdot n_B(I), \tag{(\star)}$$

where $n_B(I) \doteq \sum_{s \in I} \mathbf{1}\{X_s \in I\}$. The intuition for (\star) is as follows: suppose that, over n separate rounds, we observe the same context $X_s = x_0$ in bin B. Then, arm a would be considered unsafe in the local bandit problem at context x_0 if its regret exceeds $\sqrt{K \cdot n}$ (i.e., the first term on the above RHS), which is a safe regret to pay for the non-contextual problem. Our broader notion (\star) extends this over the bin B by also accounting for the bias (i.e., the second term on the above RHS) of observing X_s near a given context $x_0 \in B$.

- We then propose to record an *experienced significant shift* when we experience a context X_t , for which there is no safe arm to play in the sense of (\star).
- **Definition 6.** Fix the context sequence X_1, X_2, \ldots, X_T .
- We say an arm $a \in [K]$ is unsafe at context $x \in \mathcal{X}$ on I if there exists a bin $B \in \mathcal{T}$ containing x such that arm a incurs significant regret (*) in bin B on I.
- 203 We then have the following recursive definition:

• Let $\tau_0 = 1$. We then have the following recursive definition: define the (i + 1)-th experienced significant shift as the earliest time $\tau_{i+1} \in (\tau_i, T]$ such that every arm $a \in [K]$ is unsafe at X_t on some interval $I \subset [\tau_i, \tau_{i+1}]$. We refer to intervals $[\tau_i, \tau_{i+1}), i \ge 0$, as experienced significant phases. The unknown number of such phases (by time T) is denoted \tilde{L} , whereby $[\tau_{\tilde{L}-1}\tau_{\tilde{L}})$, for $\tau_{\tilde{L}} \doteq T + 1$, is the last phase. Remark 3 (Significant Shifts Depend on Contexts). It should be understood that the significant shifts

- 210 τ_i and \tilde{L} depend on X_T and mean rewards $\{f_t^a(X_t)\}_{t\in[T],a\in[K]}$, but not the realized rewards Y_T .
- For simplicity of presentation, we will not make the dependence on X_T explicit in most places where
- 212 τ_i, \tilde{L} are mentioned.

It's clear from Definition 6 and (*) that only changes in the mean rewards $f_t^a(x)$ at experienced contexts $x \in \mathbf{X}_T$ are counted, and that they are only counted when experienced. Furthermore, an experienced significant shift τ_i implies a best-arm change at X_{τ_i} since, by smoothness (Assumption 1), and (*) we have

$$\sum_{s\in I} \delta_s^a(X_{\tau_i}) \cdot \mathbf{1}\{X_s \in B\} \ge \sum_{s\in I} \delta_s(a) \cdot \mathbf{1}\{X_s \in B\} - r(B) \sum_{s\in I} \mathbf{1}\{X_s \in B\} > 0.$$

Thus, $\tilde{L} \leq L + 1$, the global count of shifts.

On the other hand, so long as an experienced significant shift does not occur, there will be arms safe to play at each context X_t . As a result, procedures need not restart exploration so long as unsafe arms can be quickly ruled out.

As a warmup to presenting our main regret bounds and algorithms, we'll first consider an oracle procedure which restarts only at experienced significant shifts.

Definition 7 (Oracle Procedure). For each round t in phase $[\tau_i, \tau_{i+1})$, define a good arm set \mathcal{G}_t as the set of safe arms, i.e., arms which do not yet satisfy (*) in bin $T_r(X_t)$ for $r = r_{\tau_{i+1}-\tau_i}$ (recall from Subsection 2.3 that this is the oracle choice of level over phase $[\tau_i, \tau_{i+1})$).

226 Then, define an oracle procedure π : at each round t, π plays a random arm $a \in \mathcal{G}_t$ w.p. $1/|\mathcal{G}_t|$.

We then claim such an oracle procedure attains an enhanced dynamic regret rate in terms of the significant shifts $\{\tau_i\}_i$ which recovers the minimax lower bound in terms of global number of shifts L and total variation V_T from before.

Proposition 2 (Sanity Check). We have the oracle procedure π of Definition 7 satisfies with probability at least $1 - 1/T^2$ w.r.t. the randomness of X_T : for some C > 0

$$\mathbb{E}_{\pi}[R_{T}(\pi, \mathbf{X}_{T}) \mid \mathbf{X}_{T}] \leq C \log(K) \log(T) \sum_{i=1}^{\tilde{L}(\mathbf{X}_{T})} (\tau_{i}(\mathbf{X}_{T}) - \tau_{i-1}(\mathbf{X}_{T}))^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}.$$

²³² *Proof.* See Appendix C.

By Jensen's inequality on the concave function $z \mapsto z^{\frac{1+d}{2+d}}$, the above regret rate is at most $\tilde{L}(\mathbf{X}_T)^{\frac{1}{2+d}}$. $T^{\frac{1+d}{2+d}} \ll L^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}$. At the same time, the rate is also faster than $V_T^{\frac{1}{3+d}}T^{\frac{1+d}{2+d}}$ (see Corollary 5). We next aim to design an algorithm which can attain the same regret without knowledge of τ_i or \tilde{L} .

236 3.3 Main Results: Adaptive Upper-bounds

Our main result is a dynamic regret upper bound of similar order to Proposition 2 without knowledge of the environment, e.g., the significant shift times, or the number of significant phases. It is stated for our algorithm CMETA (Algorithm 1 of Section 4), which, for simplicity, requires knowledge of the time horizon T (knowledge of T removable using doubling tricks).

Theorem 3. Let π denote the CMETA procedure. Let $\{\tau_i(X_T)\}_{i=0}^{\tilde{L}+1}$ denote the unknown experienced significant shifts (Definition 6). We then have with probability at least $1 - 1/T^2$ w.r.t. the randomness of X_T , for some C > 0:

$$\mathbb{E}[R_T(\pi, X_T) \mid X_T] \le C \log^4(T) \sum_{i=1}^{\tilde{L}(X_T)} (\tau_i(X_T) - \tau_{i-1}(X_T))^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}.$$

By Jensen's inequality, since the function $z \mapsto z^{\frac{1+d}{2+d}}$ is concave, the above regret rate is upper bounded by $\tilde{L}(\mathbf{X}_T)^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}$,

Corollary 4 (Adapting to Experienced Significant Shifts). Under the conditions of Theorem 3, with probability at least $1 - 1/T^2$ w.r.t. the randomness in X_T :

$$\mathbb{E}[R_T(\pi, \mathbf{X}_T) \mid \mathbf{X}_T] \le C \log^4(T) \cdot (K \cdot \tilde{L}(\mathbf{X}_T))^{\frac{1}{2+d}} \cdot T^{\frac{1+a}{2+d}}.$$

. . .

- Note, this is tighter than the earlier mentioned $(K \cdot L)^{\frac{1}{2+d}} T^{\frac{1+d}{2+d}}$ rate. The next corollary asserts that
- ²⁴⁹ Theorem 3 also recovers the optimal rate in terms of total-variation V_T .
- **Corollary 5** (Adapting to Total Variation). Under the conditions of Theorem 3, taking expectation over X_T :

$$\mathbb{E}[R_T(\pi, X_T)] \le C \log^4(T) \left(T^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} + (V_T \cdot K)^{\frac{1}{3+d}} \cdot T^{\frac{2+d}{3+d}} \right)$$

252 **4** Algorithm

Algorithm 1: Contextual Meta-Elimination while Tracking Arms (CMETA)

Input: horizon *T*, set of arms [*K*], tree *T* with levels $r \in \mathcal{R}$. **1 Initialize:** round count $t \leftarrow 1$. **2 Episode Initialization (setting global variables): 3** $t_{\ell} \leftarrow t$. ; // t_{ℓ} indicates start of ℓ -th episode. **4** For each bin $B \in \mathcal{T}$, set $\mathcal{A}_{\text{master}}(B) \leftarrow [K]$; // Initialize master candidate arm sets **5** For each $m = 2, 4, \ldots, 2^{\lceil \log(T) \rceil}$ and $s = t_{\ell} + 1, \ldots, T$: **6** Sample and store $Z_{m,s} \sim \text{Bernoulli} \left(\left(\frac{1}{m}\right)^{\frac{1}{2+d}} \cdot \left(\frac{1}{s-t_{\ell}}\right)^{\frac{1+d}{2+d}} \right)$. ; // Set replay schedule. **7** Run Base-Alg $(t_{\ell}, T + 1 - t_{\ell})$. **8 if** t < T then restart from Line 2 (i.e. start a new episode). ;

Algorithm 2: Base-Alg (t_{start}, m_0) : binned successive elimination with randomized arm-pulls

Input: starting round t_{start} , scheduled duration m_0 . 1 **Initialize**: $t \leftarrow t_{\text{start}}$ For each bin B at any level in \mathcal{T} , set $\mathcal{A}(B) \leftarrow [K]$ 2 while $t \leq t_{\text{start}} + m_0$ do **Choose level in** \mathcal{R} : $r \leftarrow r_{t-t_{\text{start}}}$. Let $\mathcal{A}_t \leftarrow \mathcal{A}(B)$ and let $B \leftarrow T_r(X_t)$. 3 4 Play a random arm $a \in A_t$ selected with probability $1/|A_t|$. 5 Increment $t \leftarrow t + 1$. 6 if $\exists m$ such that $Z_{m,t} > 0$ then 7 Let $m \doteq \max\{m \in \{2, 4, \dots, 2^{\lceil \log(T) \rceil}\} : Z_{m,t} > 0\}$.; // Set maximum replay length. 8 Run Base-Alg (t, m).; // Replay interrupts. 9 Evict bad arms in bin B: 10 $\mathcal{A}(B) \leftarrow \mathcal{A}(B) \setminus \{a \in [K] :$ 11 \exists rounds $[s_1, s_2] \subseteq [t_{\text{start}}, t)$ s.t. (5) holds for bin $T_{s_2-s_1}(X_t)$ }. $\mathcal{A}_{\text{master}}(B) \leftarrow \mathcal{A}_{\text{master}}(B) \setminus \{a \in [K] :$ 12 \exists rounds $[s_1, s_2] \subseteq [t_{\ell}, t)$ s.t. (5) holds for bin $T_{s_2-s_1}(X_t)$. // Discard arms previously discarded in ancestor bins **Refine candidate arms:** ; 13 $\mathcal{A}(B) \leftarrow \cap_{B' \in \mathcal{T}, B \subset B'} \mathcal{A}(B').$ 14 $\mathcal{A}_{\text{master}}(B) \leftarrow \cap_{B' \in \mathcal{T}, B \subseteq B'} \mathcal{A}_{\text{master}}(B').$ 15 **Restart criterion: if** $\mathcal{A}_{\text{master}}(B) = \emptyset$ for some bin *B* then RETURN.; 16 17 RETURN.

We take a similar algorithmic approach to Suk and Kpotufe [2022], with several important modifications for our setting. The high-level strategy is to schedule multiple copies of a *base algorithm* (Algorithm 2) Base-Alg at random times and durations, in order to ensure updated and reliable estimation of the gaps in (\star). This allows fast enough detection of unknown experienced significant shifts.

Overview of Algorithm Hierarchy. Our main algorithm CMETA (Algorithm 1) proceeds in episodes, each of which begins by playing according to an initially scheduled base algorithm of possible duration equal to the number of rounds left till *T*. Base algorithms occasionally activate their own base algorithms of varying durations (Line 9 of Algorithm 2), called *replays*, according to a random schedule (stored in the variable $\{Z_{m,s}\}$). We refer to the base algorithm playing at round t as the *active base algorithm*. This induces a hierarchy of base algorithms, from *parent* to *child* instances of Base-Alg.

Choice of Level. Focusing on a single base algorithm now, each Base-Alg manages its own discretization of the context space $\mathcal{X} = [0, 1]^d$, corresponding to a level $r \in \mathcal{R}$ (see Definition 3). Within each bin $B \in \mathcal{T}_r$ at the level r, candidate arms, maintained in a set $\mathcal{A}(B)$, are evicted according to estimates (4) of local gaps.

As said earlier in Subsection 2.3, key in attaining optimal regret is using the right level $r \in \mathcal{R}$. An immediate difficulty is that the oracle choice of level used in Definition 7 depends on the unknown significant phase length $\tau_{i+1} - \tau_i$. To circumvent this, as in previous works [Perchet and Rigollet, 2013, Slivkins, 2014], we rely on an adaptive time-varying choice of level r_t . Specifically, each base algorithm choose the level $r_{t-t_{start}}$ based on the time elapsed since the time t_{start} it was first activated.

274 **Sharing Information across Base Algorithms.** Instances of Base-Alg and CMETA share informa-275 tion, in the form of *global variables* as listed below:

• All variables defined in CMETA, namely $t_{\ell}, t, \{\mathcal{A}_{\text{master}}(B)\}_{B \in \mathcal{T}}, \{Z_{m,t}\}$ (see Lines 3–6 of Algorithm 1).

• All arms played at any round t, along with observed rewards Y_t^a , and the candidate arm set \mathcal{A}_t which takes the value of the set $\mathcal{A}(B)$ of the active Base-Alg at round t and bin $B = T_r(X_t)$ used.

By sharing these global variables, any Base-Alg can trigger a new episode: every time an arm is evicted from $\mathcal{A}(B)$ a Base-Alg, it is also evicted from $\mathcal{A}_{master}(B)$, which is essentially the candidate arm set for the current episode. A new episode is triggered at time t when $\mathcal{A}_{master}(B)$ becomes empty for some bin B (necessarily a currently experienced bin), i.e., there is no *safe* arm left to play at the context X_t in the sense of Definition 6.

Note that $\mathcal{A}(B)$ are *local variables* internal to each Base-Alg (the owner of which will be clear from context in usage).

To ensure consistent behavior while using a time-varying choice of level, we enforce further regularity in arm evictions across \mathcal{X} : arms evicted from $\mathcal{A}(B')$ are also evicted from child bins $B \subseteq B'$ to ensure $\mathcal{A}(B) \subseteq \mathcal{A}(B')$.

Estimating Aggregate Local Gaps. The quantity $\sum_{s=s_1}^{s_2} \delta_s(a', a) \cdot \mathbf{1}\{X_s \in B\}$ is estimated as $\sum_{s=s_1}^{s_2} \hat{\delta}_s^B(a', a)$, whereby the relative gap $\delta_s(a', a) \cdot \mathbf{1}\{X_s \in B\}$ is estimated by importance weighting as:

$$\hat{\delta}_s^B(a',a) \doteq |\mathcal{A}_t| \cdot \left(Y_t^{a'} \cdot \mathbf{1}\{\pi_t = a'\} - Y_t^a \cdot \mathbf{1}\{\pi_t = a\} \right) \cdot \mathbf{1}\{a \in \mathcal{A}_t\} \cdot \mathbf{1}\{X_s \in B\}.$$
(4)

Note that the above is an unbiased estimate of $\delta_t(a', a) \cdot \mathbf{1}\{X_s \in B\}$ whenever a' and a are both in \mathcal{A}_t at time t, conditional on the contexts X_t . It then follows that, conditional on \mathbf{X}_T , the difference $\sum_{t=s_1}^{s_2} \left(\hat{\delta}_t^B(a', a) \cdot \mathbf{1}\{X_s \in B\} - \delta_t(a', a) \right)$ is a martingale that concentrates at a rate roughly $\sqrt{K \cdot n_B([s_1, s_2])}$, where recall from earlier that $n_B(I) \doteq \sum_{s \in I} \mathbf{1}\{X_s \in I\}$ is the context count in bin B over interval I.

ut An arm *a* is then evicted at round *t* if, for some fixed $C_0 > 0^{-1}$, \exists rounds $s_1 < s_2 \le t$ such that at level $r_{s_2-s_1}$ and (i.e., the bin at level $r_{s_2-s_1}$ containing X_t) letting $B := T_{s_2-s_1}(X_t)$ (i.e., the bin at level $r_{s_2-s_1}$ containing X_t)

$$\max_{a' \in [K]} \sum_{s=s_1}^{s_2} \hat{\delta}_s^B(a', a) > \log(T) \sqrt{C_0 \cdot (Kn_B([s_1, s_2]) \vee K^2)} + r_{s_2 - s_1} \cdot n_B([s_1, s_2]).$$
(5)

 $^{{}^{1}}C_{0} > 0$ needs to be sufficiently large, but is a universal constant free of the horizon T or any distributional parameters.

301 5 Key Technical Highlights of Analysis

While a full analysis is deferred to Appendix D due to space constraints, we highlight some of the key novelties and core points of the analysis.

• Local Safety in Bins implies Safe Total Regret. We first argue that the notion of significant regret (*) within a bin *B* captures the total regret rates $T^{\frac{1+d}{2+d}}$ we wish to compete with. If (*) holds for no intervals $[s_1, s_2]$ in all bins *B*, arm *a* would be safe and incur little regret over any $[s_1, s_2]$. As it turns out, bounding the per-bin regret by (*) implies a total regret of $T^{\frac{1+d}{2+d}}$ as seen from the following rough calculation: via concentration and the strong density assumption (Assumption 2) to conflate $n_B([1,T]) \approx r(B)^d \cdot T$ and the fact that there are $\approx r^{-d}$ bins at level *r*, we have:

$$\sum_{B \in T_r} \sqrt{K \cdot n_B([1,T])} + r \cdot n_B([1,T]) \le K^{1/2} \cdot T^{1/2} \cdot r^{-d/2} + T \cdot r.$$
(6)

In particular taking $r \propto (K/T)^{\frac{1}{2+d}}$ makes the above RHS the desired rate $K^{\frac{1}{2+d}}T^{\frac{1+d}{2+d}}$.

• Significant Regret Threshold is Estimation Error. At the same time, the RHS of the definition of significant regret (\star) is a variance and bias decomposition of the bound on the (conditional on \mathbf{X}_T) error of estimating the cumulative regret $\sum_{s=s_1}^{s_2} \delta_s^a(x) \cdot \mathbf{1}\{X_s \in B\}$ at any context $x \in B$. Thus, intuitively, changes of magnitude above the threshold $\sqrt{K \cdot n_B(I)} + r(B) \cdot n_B(I)$ in (\star) are detectable.

So, the notion of significant regret (*) perfectly balances both (1) detection of unsafe arms and (2) regret minimization of playing safe arms.

• A New Balanced Replay Scheduling. As mentioned earlier in Subsection 3.1, previous adaptive works on contextual bandits fail to attain the optimal regret in this setting due to an inappropriate frequency of scheduling replays. We introduce a novel scheduling (Line 6 of Algorithm 1) which carefully balances exploration and fast detection of significant regret in the sense of (*). The chosen rate $(1/m)^{\frac{1}{2+d}}(1/t)^{\frac{1+d}{2+d}}$ comes from the following intuitive calculation. A scheduled replay of duration m will incur an additional regret of about $m^{\frac{1+d}{2+d}}$. Then, summing over all possible replays, the extra regret incurred due to replays is in total roughly upper bounded by

$$\sum_{t=1}^{T} \sum_{m=2,4,\dots,T} \left(\frac{1}{m}\right)^{\frac{1}{2+d}} \left(\frac{1}{t}\right)^{\frac{1+d}{2+d}} \cdot m^{\frac{1+d}{2+d}} \lesssim \sum_{t=1}^{T} T^{\frac{d}{2+d}} \cdot (1/t)^{\frac{1+d}{2+d}} \lesssim T^{\frac{1+d}{2+d}}.$$

In other words, the cost of replays only incurs extra constants in the regret. Surprisingly, this scheduling rate is also sufficient for detecting significant regret in *any* experienced subregion *B* of the context space \mathcal{X} , i.e. there is no need to do additional exploration on a localized per-bin basis.

Next, a key feature of the analysis is that one need only minimize regret and detect changes at the critical level $r_{s_2-s_1} \propto (K/(s_2-s_1))^{\frac{1}{2+d}}$. In particular, the following two observations play a major role in bounding the regret.

• Suffices to Only Check (\star) at Critical Levels $r_{s_2-s_1}$. At first glance, detecting experienced significant shifts (Definition 6) appears difficult as an arm *a* may incur significant regret over a different bin *B'* from the bin *B* that is currently being used by the algorithm.

This difficulty is further compounded by the fact there may even be missing data problems as arms $a \in \mathcal{A}(B)$ in contention at B may have been evicted from sibling bins of the parent $B' \supset B$, thus preventing reliable estimation of a across B'. We in fact show that we only require detecting significant regret in bins B' at the critical level $r_{s_2-s_1}$ and only for the arms still in contention across all of B'. In other words, changes at other levels are all accounted for by changes at this critical level.

Additionally, we observe that the calculations in (6) would hold if we were just concerned with checking (\star) for intervals $[s_1, s_2]$ and bins $B_{s_2-s_1}$ at level $r_{s_2-s_1} := \left(\frac{K}{s_2-s_1}\right)^{\frac{1}{2+d}}$. Thus, the critical level $r_{s_2-s_1}$ is the key to both regret minimization and experienced significant shift detection

References 342

Yasin Abbasi-Yadkori, András György, and Nevena Lazic. A new look at dynamic regret for 343 non-stationary stochastic bandits. arXiv preprint arXiv:2201.06532, 2022. 344

Alekh Agarwal, Daniel Hsu, Satyen Kale, John Langford, Lihong Li, and Robert Schapire. Taming 345 the monster: A fast and simple algorithm for contextual bandits. volume 32 of Proceedings of 346 Machine Learning Research, pages 1638–1646. PMLR, 22–24 Jun 2014. 347

- Robin Allesiardo, Raphaël Féraud, and Odalric-Ambrym Maillard. The non-stationary stochastic 348
- multi-armed bandit problem. International Journal of Data Science and Analytics, 3(4):267-283, 349 2017. 350
- Sakshi Arya and Yuhong Yang. Randomized allocation with nonparametric estimation for contextual 351 multi-armed bandits with delayed rewards. Statistics & Probability Letters, 164:108818, 2020. 352 ISSN 0167-7152. 353
- Jean-Yves Audibert and Alexander B Tsybakov. Fast learning rates for plug-in classifiers. The Annals 354 of Statistics, 35(2):608-633, 2007. 355
- Peter Auer, Pratik Gajane, and Ronald Ortner. Adaptively tracking the best bandit arm with an 356 357 unknown number of distribution changes. *Conference on Learning Theory*, pages 138–158, 2019.
- Omar Besbes, Yonatan Gur, and Assaf Zeevi. Optimal exploration-exploitation in a multi-armed-358 bandit problem with non-stationary rewards. Stochastic Systems, 9(4):319–337, 2019. 359
- Lilian Besson, Emilie Kaufmann, Odalric-Ambrym Maillard, and Julien Seznec. Efficient change-360 point detection for tackling piecewise-stationary bandits. Journal of Machine Learning Research, 361 23(77):1-40, 2022. 362
- Alina Beygelzimer, John Langford, Lihong Li, Lev Reyzin, and Robert E. Schapire. Contextual 363 bandit algorithms with supervised learning guarantees. AISTATS, 2011. 364
- Moise Blanchard, Steve Hanneke, and Patrick Jaillet. Non-stationary contextual bandits and universal 365 learning. arXiv preprint arXiv:2302.07186, 2023. 366
- Changxiao Cai, T. Tony Cai, and Hongzhe Li. Transfer learning for contextual multi-armed bandits. 367 arxiv preprint arXiv:2211.12612, 2022. 368
- Yang Cao, Zheng Wen, Branislav Kveton, and Yao Xie. Nearly optimal adaptive procedure with 369 370 change detection for piecewise-stationary bandit. Proceedings of the 22nd International Conference on Artificial Intelligence and Statistics (AISTATS), 2019. 371
- Livu Chen and Haipeng Luo. Near-optimal goal-oriented reinforcement learning in non-stationary 372 environments. Advances in Neural Information Processing Systems, 2022. 373
- Yifang Chen, Chung-Wei Lee, Haipeng Luo, and Chen-Yu Wei. A new algorithm for non-stationary 374 contextual bandits: efficient, optimal, and parameter-free. In 32nd Annual Conference on Learning 375 Theory, 2019. 376
- Wang Chi Cheung, David Simchi-Levi, and Ruihao Zhu. Reinforcement learning for non-stationary 377 Markov decision processes: The blessing of (more) optimism. In International Conference on 378 Machine Learning, pages 1843–1854. PMLR, 2020. 379
- Wang Chi Cheung, David Simchi-Levi, and Ruihao Zhu. Hedging the drift: learning to optimize under 380 non-stationarity. In Proceedings of the 22nd International Conference on Artificial Intelligence 381 and Statistics, 2019. 382
- Yuhao Ding and Javad Lavaei. Provably efficient primal-dual reinforcement learning for CMDPs 383 with non-stationary objectives and constraints. AAAI Conference on Artificial Intelligence (AAAI), 384 2023. 385
- Omar Darwiche Domingues, Pierre Ménard, Matteo Pirotta, Emilie Kaufmann, and Michal Valko. A 386 kernel-based approach to non-stationary reinforcement learning in metric spaces. In International 387
- Conference on Artificial Intelligence and Statistics, pages 3538–3546. PMLR, 2021. 388

Miroslav Dudik, Daniel Hsu, Satyen Kale, Nikos Karampatziakis, John Langford, Lev Reyzin, and
 Tong Zhang. Efficient optimal learning for contextual bandits. In *Proceedings of the Twenty-Seventh*

Conference on Uncertainty in Artificial Intelligence, page 169–178. AUAI Press, 2011.

- Yingjie Fei, Zhuoran Yang, Zhaoran Wang, and Qiaomin Xie. Dynamic regret of policy optimization
 in non-stationary environments. *Advances in Neural Information Processing Systems*, 33:6743–
 6754, 2020.
- Dylan Foster and Alexander Rakhlin. Beyond ucb: Optimal and efficient contextual bandits with
 regression oracles. *Proceedings of the 37th International Conference on Machine Learning*, 119:
 3199–3210, 2020.
- Dylan Foster, Alekh Agarwal, Miroslav Dudik, Haipeng Luo, and Robert Schapire. Practical
 contextual bandits with regression oracles. In *Proceedings of the 35th International Conference on Machine Learning*, volume 80 of *Proceedings of Machine Learning Research*, pages 1539–1548.
 PMLR, 10–15 Jul 2018.
- Pratik Gajane, Ronald Ortner, and Peter Auer. A sliding-window algorithm for Markov decision
 processes with arbitrarily changing rewards and transitions. *arXiv preprint arXiv:1805.10066*,
 2018.
- Aurélien Garivier and Eric Moulines. On upper-confidence bound policies for switching bandit
 problems. In *Proceedings of the 22nd International Conference on Algorithmic Learning Theory*,
 pages 174–188. ALT 2011, Springer, 2011.
- ⁴⁰⁸ Melody Y Guan and Heinrich Jiang. Nonparametric stochastic contextual bandits. AAAI, 2018.
- Yonatan Gur, Ahmadreza Momeni, and Stefan Wager. Smoothness-adaptive contextual bandits.
 Operations Research, 70(6):3198–3216, 2022.
- Yichun Hu, Nathan Kallus, and Xiaojie Mao. Smooth contextual bandits: Bridging the parametric
 and non-differentiable regret regimes. *Conference on Learning Theory*, 2020.
- ⁴¹³ Thomas Jaksch, Ronald Ortner, and Peter Auer. Near-optimal regret bounds for reinforcement ⁴¹⁴ learning. *Journal of Machine Learning Research*, 11:1563–1600, 2010.
- Zohar S Karnin and Oren Anava. Multi-armed bandits: Competing with optimal sequences. In
 Advances in Neural Information Processing Systems, pages 199–207, 2016.
- Akshay Krishnamurthy, John Langford, Aleksandrs Slivkins, and Chicheng Zhang. Contextual
 bandits with continuous actions: Smoothing, zooming, and adapting. In *Proceedings of the Thirty- Second Conference on Learning Theory*, volume 99 of *Proceedings of Machine Learning Research*,
 pages 2025–2027. PMLR, 25–28 Jun 2019.
- John Langford and Tong Zhang. The epoch-greedy algorithm for multi-armed bandits with side information. In *Advances in neural information processing systems*, pages 817–824, 2008.
- Fang Liu, Joohyun Lee, and Ness Shroff. A change-detection based framework for piecewisestationary multi-armed bandit problem. *Proceedings of the AAAI Conference on Artificial Intelligence*, 2018.
- Tyler Lu, Dávid Pál, and Martin Pál. Showing relevant ads via context multi-armed bandits. In
 Proceedings of AISTATS, 2009.
- Haipeng Luo, Chen-Yu Wei, Alekh Agarwal, and John Langford. Efficient contextual bandits in
 non-stationary worlds. In *Conference On Learning Theory*, pages 1739–1776. PMLR, 2018.
- Thodoris Lykouris, Max Simchowitz, Alex Slivkins, and Wen Sun. Corruption-robust exploration in
 episodic reinforcement learning. In *Conference on Learning Theory*, pages 3242–3245. PMLR,
 2021.
- Weichao Mao, Kaiqing Zhang, Ruihao Zhu, David Simchi-Levi, and Tamer Basar. Near-optimal
 model-free reinforcement learning in non-stationary episodic mdps. In *International Conference on Machine Learning*, pages 7447–7458. PMLR, 2021.

- Subhojyoti Mukherjee and Odalric-Ambrym Maillard. Distribution-dependent and time-uniform
 bounds for piecewise i.i.d bandits. *Reinforcement Learning for Real Life (RL4RealLife) Workshop*
- *in the 36th International Conference on Mearning Learning*, 2019.
- Ronald Ortner, Pratik Gajane, and Peter Auer. Variational regret bounds for reinforcement learning.
 In *Proceedings of The 35th Uncertainty in Artificial Intelligence Conference*, volume 115 of
 Proceedings of Machine Learning Research, pages 81–90. PMLR, 22–25 Jul 2020.
- Vianney Perchet and Philippe Rigollet. The multi-armed bandit problem with covariates. *The Annals of Statistics*, 41(2):693–721, 2013.
- Yury Polyanskiy and Yihong Wu. *Information Theory: From Coding to Learning*. Cambridge
 University Press, 2022.
- Wei Qian and Yuhong Yang. Kernel estimation and model combination in a bandit problem with covariates. *Journal of Machine Learning Research*, 17(149):1–37, 2016a.
- Wei Qian and Yuhong Yang. Randomized allocation with arm elimination in a bandit problem with covariates. *Electronic Journal of Statistics*, 10(1):242 – 270, 2016b.
- Henry Reeve, Joe Mellor, and Gavin Brown. The k-nearest neighbour ucb algorithm for multi-armed
 bandits with covariates. In *Proceedings of Algorithmic Learning Theory*, volume 83 of *Proceedings of Machine Learning Research*, pages 725–752. PMLR, 07–09 Apr 2018.
- 453 Phillipe Rigollet and Assaf Zeevi. Nonparametric bandits with covariates. COLT, 2010.
- 454 Jyotirmoy Sarkar. One-armed bandit problems with covariates. *The Annals of Statistics*, pages 455 1978–2002, 1991.
- David Simchi-Levi and Yunzong Xu. Bypassing the monster: a faster and simpler optimal algorithm
 for contextual bandits under realizability. *Mathematics of Operations Research*, 47(3):1904–1931,
 2021.
- Aleksandrs Slivkins. Contextual bandits with similarity information. *The Journal of Machine Learning Research*, 15(1):2533–2568, 2014.
- Joe Suk and Samory Kpotufe. Tracking most significant arm switches in bandits. In *Proceedings of Thirty Fifth Conference on Learning Theory*, volume 178 of *Proceedings of Machine Learning Research*, pages 2160–2182. PMLR, 02–05 Jul 2022.
- Joseph Suk and Samory Kpotufe. Self-tuning bandits over unknown covariate-shifts. *International Conference on Algorithmic Learning Theory*, 2021.
- Ahmed Touati and Pascal Vincent. Efficient learning in non-stationary linear Markov decision
 processes. *arXiv preprint arXiv:2010.12870*, 2020.
- Chen-Yu Wei and Haipeng Luo. Non-stationary reinforcement learning without prior knowledge:
 An optimal black-box approach. *Proceedings of the 32nd International Conference on Learning Theory*, 2021.
- Chen-Yu Wei, Christoph Dann, and Julian Zimmert. A model selection approach for corruption
 robust reinforcement learning. In *International Conference on Algorithmic Learning Theory*, pages
 1043–1096. PMLR, 2022.
- Lai Wei and Vaihbav Srivatsva. On abruptly-changing and slowly-varying multiarmed bandit problems. *Annual American Control Conference (ACC)*, 2018.
- 476 Michael Woodroofe. A one-armed bandit problem with a concomitant variable. *Journal of the* 477 *American Statistical Association*, 74(368):799–806, 1979.
- 478 Qingyun Wu, Naveen Iyer, and Hongning Wang. Learning contextual bandits in a non-stationary
 environment. In *The 41st International ACM SIGIR Conference on Research & Development in* Information Retrieval, pages 495–504, 2018.

- Yuhong Yang, Dan Zhu, et al. Randomized allocation with nonparametric estimation for a multi armed bandit problem with covariates. *The Annals of Statistics*, 30(1):100–121, 2002.
- 483 Huozhi Zhou, Jinglin Chen, Lav R. Varshney, and Ashish Jagmohan. Nonstationary reinforcement
- learning with linear function approximation. *Transactions on Machine Learning Research*, 2022.
 ISSN 2835-8856.

A Details for Specializing Previous Contextual Bandit Results to Lipschitz Contextual Bandits

488 A.1 Finite Policy Class Contextual Bandits

In the finite policy class setting², one is given access to a known finite class Π of policies $\pi : \mathcal{X} \to [K]$, and in the non-stationary variant, seeks to minimize regret to the time-varying benchmark of best policies $\pi_t^* := \operatorname{argmax}_{\pi \in \Pi} \mathbb{E}_{(X,Y) \in \mathcal{D}_t}[Y(\pi(X))]$. In other words, the "dynamic regret" in this setting is defined by (for chosen policies $\{\hat{\pi}_t\}_t$)

$$\mathbb{E}\left[\sum_{t=1}^{T}\max_{\pi\in\Pi}\mathbb{E}_{(X,Y)\in\mathcal{D}_{t}}[Y(\pi(X))] - \sum_{t=1}^{T}Y_{t}(\hat{\pi}_{t})\right].$$
(7)

We can in fact recover the nonparametric setting and relate the above to our notion of dynamic regret (Definition 2). To do so, we let Π be the class of policies which uses a level $r \in \mathcal{R}$ and discretizes decision-making across individual bins $B \in \mathcal{T}_r$. Then, we claim there is an *oracle sequence of policies* $\{\pi_t^{\text{oracle}}\}_t$ which attains the minimax regret rate of Theorem 1. So, it remains to bound the regret to the sequence $\{\pi_t^{\text{oracle}}\}_t$ in the sense above.

• Parametrizing in Terms of Global Number L of Shifts. Suppose there are L + 1 stationary phases of length T/(L+1). Then, we first claim there is an oracle sequence of policies π_t^{oracle} which attains reget $(L+1)^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}$.

First, recall from Subsection 2.3 the *oracle choice of level* r_n for a stationary period of n rounds, or

the level $r_n \propto (K/n)^{\frac{1}{2+d}}$. Now, define $\{\pi_t^{\text{oracle}}\}_t$ as follows: at each round t, π_t^{oracle} uses the oracle

level $r := r_{T/(L+1)} \propto \left(\frac{K(L+1)}{T}\right)^{\frac{1}{2+d}}$ and plays in each bin $B \in \mathcal{T}_r$, the arm maximizing the average reward in that bin $\mathbb{E}[f_t^a(X_t) \mid X_t \in B]$. As this is a biased version of the actual bandit problem $\{f_t^a(X_t)\}_{a \in [K]}$ at context X_t , it will follow that π_t^{oracle} incurs regret of order the bias of estimation in B which is r.

⁵⁰⁷ Concretely, suppose X_t falls in bin B at level r, and let $\pi_t^{\text{oracle}}(B)$ be the arm selected at round t by ⁵⁰⁸ π_t^{oracle} in bin B. Then, mean rewards are Lipschitz, each policy π_t^{oracle} suffers regret:

$$\max_{a \in [K]} f_t^a(X_t) - f_t^{\pi_t^{\text{cracle}}(B)}(X_t) \le \max_{a \in [K]} \mathbb{E}[f_t^a(X_t) - f_t^{\pi_t^{\text{cracle}}(B)}(X_t) \mid X_t \in B] + r = r.$$

Thus, the sequence of policies $\{\pi_t^{\text{pracle}}\}_t$ achieves dynamic regret (in the sense of Definition 2)

$$\mathbb{E}\left[\sum_{t=1}^{T}\max_{a\in[K]}f_t^a(X_t) - f_t^{\pi_t^{\text{oracle}}(X_t)}(X_t)\right] \lesssim (L+1)\cdot \left(\frac{T}{L+1}\right)\cdot \left(\frac{K}{(L+1)T}\right)^{\frac{1}{2+d}} \propto L^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}.$$

Thus, it suffices to minimize dynamic regret in the sense of (7) to this oracle policy π_t^{oracle} . The state-of-the-art guarantee in this setting is that of the ADA-ILTCB algorithm of Chen et al. [2019], which achieves a dynamic regret of $\sqrt{KLT \log(|\Pi|)}$. It then remains to compute $|\Pi|$.

As we need only consider levels in \mathcal{R} of size at least $(K/T)^{\frac{1}{2+d}}$, the size of the policy class II is $|\Pi| = \sum_{r \in \mathcal{R}} K^{r^{-d}} \propto K^{(T/K)^{\frac{d}{2+d}}} \implies \log(|\Pi|) = \left(\frac{T}{K}\right)^{\frac{d}{2+d}} \log(K)$. Plugging this into $\sqrt{KLT \log(|\Pi|)}$ gives a regret rate of $K^{\frac{1}{2+d}} \cdot L^{1/2}T^{\frac{1+d}{2+d}}$, which has a worse dependence on the global number of shifts L than the minimax optimal rate of $L^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}$ (see Theorem 1).

• **Parametrizing in Terms of Total-Variation** V_T . Fix any positive real number $V \in [T^{-\frac{3+d}{2+d}}, T]$. Then, the lower bound construction of Theorem 1 reveals that there exists an environment with

⁵¹⁹ $L+1 = T/\Delta$ stationary phases of length $\Delta \doteq \left[\left(\frac{T}{V}\right)^{\frac{2+d}{3+d}} \right]$ and total-variation of order V.

²While there are matters of efficiency and what offline learning guarantees may be assumed in this broader *agnostic* setting, we do not discuss these here, and readers are deferred to Langford and Zhang [2008], Dudik et al. [2011], Agarwal et al. [2014].

Then, the earlier defined oracle sequence of policies $\{\pi_t^{\text{oracle}}\}_t$ attains optimal dynamic regret in terms of V_T :

$$(L+1)^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}} \propto T^{\frac{2+d}{3+d}} \cdot V^{\frac{1+d}{3+d}}$$

Meanwhile, the state-of-the-art regret guarantee in this parametrization is Theorem 2 of Chen et al. [2019], where ADA-ILTCB's regret bound becomes:

$$(K \cdot \log(|\Pi|) \cdot V)^{1/3} T^{2/3} + \sqrt{K \log(|\Pi|) \cdot T} \propto K^{\frac{2}{3(2+d)}} \cdot V^{\frac{1}{3}} \cdot T^{\frac{2+d}{3+d} + \frac{d}{3(2+d)(3+d)}} + K^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}$$

We claim this rate is worse than our rate in Corollary 5, in fact in all parameters V, K, T. For $K \ge T$, both rates imply linear regret. Assume K < T. Then, note by elementary calculations that for all $d \in \mathbb{N} \cup \{0\}$:

$$\frac{2}{3} + \frac{d}{3(2+d)} = \frac{2+d}{3+d} + \frac{1}{3+d} - \frac{2}{3(2+d)}$$

⁵²⁷ Then, it follows that rate of Corollary 5 is smaller using the fact that K < T:

 $K^{\frac{2}{3(2+d)}} \cdot V^{1/3} \cdot T^{\frac{2}{3} + \frac{d}{3(2+d)}} \ge K^{\frac{2}{3(2+d)}} \cdot V^{\frac{1}{3+d}} \cdot T^{\frac{2+d}{3+d}} \cdot K^{\frac{1}{3+d} - \frac{2}{3(2+d)}} \ge (KV)^{\frac{1}{3+d}} \cdot T^{\frac{2+d}{3+d}}.$

528 A.2 Realizable Contextual Bandits

Lipschitz contextual bandits is also a special case of contextual bandits with *realizability*. In this broader setting, the learner is given a function class Φ which contains the true regression function $\phi_t^* : \mathcal{X} \times [K] \to [0, 1]$ describing mean rewards of context-arm pairs at round t. The goal is to compete with the time-varying benchmark of policies $\pi_{\phi_t^*}(x) := \operatorname{argmax}_{a \in [K]} \phi_t^*(x, a)$, using calls to a regression oracle over Φ .

⁵³⁴ While the natural choice for Φ is the infinite class of all Lipschitz functions from $\mathcal{X} \times [K] \rightarrow [0, 1]$, ⁵³⁵ the state-of-the-art non-stationary algorithm only provides guarantees for finite Φ [Wei and Luo, ⁵³⁶ 2021, Appendix I.7].

However, it is still possible to recover the Lipschitz contextual bandit setting, by defining Φ similarly to how we defined the finite class of policies Π above. Let Φ be the class of all piecewise constant functions which depends on a level $r \in \mathcal{R}$, and are constant on bins $B \in \mathcal{T}_r$ at level r, taking values which are multiples of $T^{-\frac{1}{2+d}}$ (there are O(T) many such values in [0, 1]). Note this is quite similar to how we defined the policy-class Π above.

For this specification of Φ , the realizability assumption is false. Rather, this is a mildly misspecified regression class which is allowed by the stationary guarantees of FALCON [Simchi-Levi and Xu, 2021, Section 3.2]. In particular, by smoothness, at each round $t \in [T]$ there is a function $\phi_t^* \in \Phi$ such that

$$\sup_{x \in \mathcal{X}, a \in [K]} |\phi_t(x, a) - f_t^a(x)| \lesssim \left(\frac{1}{T}\right)^{\frac{1}{2+d}}.$$

This introduces an additive term in the regret bound of FALCON of order $T^{\frac{1+d}{2+d}}$ which is of the right order in our setting.

Then, the MASTER black-box algorithm using FALCON Simchi-Levi and Xu [2021] as a base algorithm can obtain dynamic regret upper bounded by [see Wei and Luo, 2021, Theorem 2]:

$$\min\left\{\sqrt{\log(|\Phi|) \cdot L \cdot T}, \log^{1/3}(|\Phi|) \cdot \Delta^{1/3} \cdot T^{2/3} + \sqrt{\log(|\Phi|) \cdot T}\right\}$$

As Φ is essentially the same size as the policy class Π defined in the previous section, the above regret bound specializes to similar rates as those of ADA-ILTCB derived above.

552 **B** Useful Lemmas

Throughout the appendix, c_1, c_2, \ldots will denote universal positive constants not depending on T, Kor any of the significant shifts $\{\tau_i(\mathbf{X}_T)\}_i$.

555 B.1 Concentration of Aggregate Gap over an Interval within a Bin

We first recall a Freedman's inequality, which will help us establish concentration of our gap estimators(Proposition 7).

Lemma 6 (Theorem 1 of Beygelzimer et al. [2011]). Let $X_1, \ldots, X_n \in \mathbb{R}$ be a martingale difference sequence with respect to some filtration $\mathcal{F}_0, \mathcal{F}_1, \ldots$ Assume for all t that $X_t \leq R$ a.s.. Then for any

560 $\delta \in (0, 1)$, with probability at least $1 - \delta$, we have:

$$\sum_{i=1}^{n} X_{i} \le (e-1) \left(\sqrt{\log(1/\delta) \sum_{i=1}^{n} \mathbb{E}[X_{i}^{2} | \mathcal{F}_{t-1}]} + R \log(1/\delta) \right).$$
(8)

⁵⁶¹ Recall from Section 4 that for round t,

$$\hat{\delta}_t^{a',a}(B) \doteq |\mathcal{A}_t| \cdot (Y_t(a') \cdot \mathbf{1}\{\pi_t = a'\} - Y_t(a) \cdot \mathbf{1}\{\pi_t = a\}) \cdot \mathbf{1}\{a \in \mathcal{A}_t\} \cdot \mathbf{1}\{X_t \in B\}.$$

- ⁵⁶² We next apply Lemma 6 to our aggregate estimator from Section 4.
- **Proposition 7.** With probability at least $1 1/T^2$ w.r.t. the randomness of Y_T , $\{\pi_t\}_t \mid X_T$, we have for all bins $B \in \mathcal{T}$ and rounds $s_1 < s_2$ and all arms $a \in [K]$ that for large enough $c_1 > 0$:

$$\left|\sum_{s=s_1}^{s_2} \hat{\delta}_s^{i_t,a}(B) - \sum_{s=s_1}^{s_2} \mathbb{E}[\hat{\delta}_s^{a',a}(B)|\mathcal{F}_{s-1}]\right| \le c_1 \log(T) \left(\sqrt{K \cdot n_B([s_1,s_2])} + K\right), \quad (9)$$

where $\mathcal{F} \doteq {\mathcal{F}_t}_{t=1}^T$ is the filtration with \mathcal{F}_t generated by ${\pi_s, Y_s^{\pi_s}}_{s=1}^t$.

- 566 *Proof.* The proof is similar to the proof of Proposition 3 in Suk and Kpotufe [2022].
- The martingale difference $\hat{\delta}_s^{a',a}(B) \mathbb{E}[\hat{\delta}_s^{a',a}(B) | \mathcal{F}_{s-1}]$ is clearly bounded above by 2K for all bins B, rounds s, and all arms a, a'. We also have a cumulative variance bound:

$$\begin{split} \sum_{s=s_1}^{s_2} \mathbb{E}[(\hat{\delta}_s^{a',a}(B))^2 \mid \mathcal{F}_{s-1}] &\leq \sum_{s=s_1}^{s_2} \mathbf{1}\{X_s \in B\} \cdot |\mathcal{A}_s|^2 \cdot \mathbb{E}[\mathbf{1}\{\pi_s = a \text{ or } a'\} | \mathcal{F}_{s-1}] \\ &\leq \sum_{s=s_1}^{s_2} \mathbf{1}\{X_s \in B\} \cdot 2|\mathcal{A}_s| \\ &\leq 2K \cdot n_B([s_1, s_2]). \end{split}$$

Then, the result follows from (8), and taking union bounds over bins B (at most T levels and at most T bins per level), arms a, a', and rounds s_1, s_2 .

571 Since the error probability of Proposition 7 is negligible with respect to regret, we assume going

forward in the analysis that (9) holds for all arms $a, a' \in [K]$ and rounds s_1, s_2 . Specifically, let \mathcal{E}_1

be the good event over which the bounds of Proposition 7 hold for all all arms and intervals $[s_1, s_2]$.

574 B.2 Concentration of Covariate Counts

Notation. To ease notation throughout, we'll henceforth use $\mu(\cdot)$ to refer to the context marginal distribution $\mu_X(\cdot)$.

Lemma 8. Let $\{i_t\}_{t=1}^T$ be a random sequence of arms whose distribution depends on X_T . With probability at least $1 - 1/T^2$ w.r.t. the randomness of X_T , we have for all bins $B \in \mathcal{T}$, all arms $a', a \in [K]$, and rounds $s_1 < s_2$, for some large enough $c_2 > 0$ the following inequalities hold:

$$|n_B([s_1, s_2]) - (s_2 - s_1 + 1) \cdot \mu(B)| \le c_2 \left(\log(T) + \sqrt{\log(T)\mu(B) \cdot (s_2 - s_1 + 1)} \right)$$
(10)

$$\left| \sum_{s=s_1}^{s_2} \delta_s(i_s, a) \cdot (\mathbf{1}\{X_s \in B\} - \mu_s(B)) \right| \le c_2 \left(\log(T) + \sqrt{\log(T)\mu(B) \cdot (s_2 - s_1 + 1)} \right)$$
(11)

$$\left|\sum_{s=s_1}^{s_2} \delta_s(a) \cdot (\mathbf{1}\{X_s \in B\} - \mu_s(B))\right| \le c_2 \left(\log(T) + \sqrt{\log(T)\mu(B) \cdot (s_2 - s_1 + 1)}\right)$$
(12)

Proof. The first inequality (10) follow from Lemma 6 since $\sum_{s=s_1}^{s_2} \mathbf{1}\{X_s \in B\} - \mu(B)$ is a martingale, which has conditional variance at most $(s_2 - s_1 + 1) \cdot \mu(B)$. 580 581

The other two inequalities are trickier since $\delta_s(a)$ depends on X_s (so that the summand may not be a 582 martingale difference) while $\delta_s(i_s, a)$ may not even be adapted to the canonical filtration generated 583 by \mathbf{X}_T (i.e., i_t may depend on X_s for s > t). Nevertheless, we observe that for any random variable 584 $W_s = W_s(\mathbf{X}_T) \in [-1, 1]:$ 585

$$-(\mathbf{1}\{X_t \in B\} - \mu(B)) \le W_t \cdot (\mathbf{1}\{X_t \in B\} - \mu(B)) \le \mathbf{1}\{X_t \in B\} - \mu(B).$$

The upper and lower bounds above are both martingale differences with respect to the canonical 586 filtration of X_T and thus, summing the above over t we have via Lemma 6: 587

$$\left| \sum_{s=s_1}^{s_2} W_s \cdot (\mathbf{1}\{X_t \in B\} - \mu(B)) \right| \le \left| \sum_{s=s_1}^{s_2} \mathbf{1}\{X_s \in B\} - \mu(B) \right| \\\le c_2 \left(\log(T) + \sqrt{\log(T)\mu(B) \cdot (s_2 - s_1 + 1)} \right).$$

Then, taking union bounds over rounds s_1, s_2 , bins $B \in \mathcal{T}$, and arms $a \in [K]$ gives the result. 588

Notation 2 (good event). Let \mathcal{E}_1 be the good event over which the bounds of Proposition 7 hold 589 for all rounds $s_1, s_2 \in [T]$ and arms $a', a \in [K]$. Thus, on \mathcal{E}_1 , our estimated gaps in each bin will 590 concentrate. 591

- Let \mathcal{E}_2 be the good event on which bounds of Lemma 8 holds for all bins B, arms $a \in [K]$, rounds 592
- $s_1, s_2 \in [T]$. Thus, on \mathcal{E}_2 , our covariate counts $n_B([s_1, s_2])$ will concentrate and we will be able to relate the empirical quantities $\sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1}\{X_s \in B\}$ with their expectations. 593
- 594

Next, we establish a lemma which allow us to relate significant regret (*) and thus our eviction 595 criterion (5) between different bins and levels. 596

Lemma 9 (Relating Aggregate Gaps Between Levels). On event \mathcal{E}_2 , if for rounds $s_1 < s_2$, bin B' at 597 level $r_{s_2-s_1}$ and arm a, for some $c_3 > 0$: 598

$$\sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1}\{X_s \in B'\} \le c_3 \left(\sqrt{K \cdot n_{B'}([s_1, s_2]) \vee K^2} + r(B') \cdot n_{B'}([s_1, s_2])\right),$$

then for any bin $B \subseteq B'$ and some $c_4 > 0$: 599

$$\sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1}\{X_s \in B\} \le c_4 \left(\log^{1/2}(T) \cdot r(B)^d \cdot K^{\frac{1}{2+d}} \cdot (s_2 - s_1)^{\frac{1+d}{2+d}} + K \log(T) + \sqrt{\log(T)\mu(B)(s_2 - s_1 + 1)} \right).$$

- The same applies for $\delta_s(a)$ replaced with $\delta_s(a', a)$ with any other fixed arm a'. 600
- *Proof.* We have using (12) and the strong density assumption (Assumption 2): 601

$$\sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1}\{X_s \in B\} \le \sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mu(B) + c_2 \left(\log(T) + \sqrt{\log(T)(s_2 - s_1 + 1) \cdot \mu(B)}\right)$$
$$\le \frac{r(B)^d}{r(B')^d} \sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mu(B') + c_2 \left(\log(T) + \sqrt{\log(T)(s_2 - s_1 + 1) \cdot \mu(B)}\right)$$
(13)

Again using (12)602

$$\begin{split} \sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mu_s(B') &\leq \sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1} \{ X_s \in B' \} + c_2 \left(\log(T) + \sqrt{\log(T)(s_2 - s_1 + 1) \cdot \mu(B')} \right) \\ &\leq c_5 \left(\sqrt{K \cdot n_{B'}([s_1, s_2]) \vee K^2} + r(B') \cdot n_{B'}([s_1, s_2]) \right) \\ &\quad + \log(T) + \sqrt{\log(T)(s_2 - s_1 + 1) \cdot \mu(B')} \end{split}$$

Next, applying (10) to $n_{B'}([s_1, s_2])$ and using the strong density assumption (Assumption 2) to bound the mass $\mu(B')$ above by $C_d \cdot r(B')^d$, the above R.H.S. is further upper bounded by

$$c_6 \left(\log^{1/2}(T) K^{\frac{1+d}{2+d}} \cdot (s_2 - s_1)^{\frac{1}{2+d}} + K \log(T) \right).$$
(14)

Finally, plugging (14) into (13) and using the fact that $(r(B')/2)^d \ge (K/(s_2 - s_1))^{\frac{d}{2+d}}$, we have that (13) is of the desired order. The proof of the same inequalities with $\delta_s(a', a)$ is analogous. \Box

The following lemma relating the bias and variance terms in the notion of significant regret (\star) will serve useful many places in the analysis. They all follow from concentration and similar calculations via the strong density assumption (Assumption 2) as done previously.

Lemma 10 (bias-variance bound and strong density). Let $r = r_{s_2-s_1}$. Then, for any bin $B \in T_r$:

$$\begin{aligned} c_7(s_2 - s_1)^{\frac{1}{2+d}} \cdot K^{\frac{d/2}{2+d}} &\leq \sqrt{(s_2 - s_1 + 1) \cdot \mu(B)} \leq c_8(s_2 - s_1)^{\frac{1}{2+d}} \cdot K^{\frac{d/2}{2+d}} \\ &\sqrt{n_B([s_1, s_2])} \leq c_9(s_2 - s_1)^{\frac{1}{2+d}} \cdot K^{\frac{d/2}{2+d}} \\ c_{10}(s_2 - s_1)^{\frac{1}{2+d}} \cdot K^{\frac{1+d}{2+d}} \leq n_B([s_1, s_2]) \cdot r \leq c_{11}(s_2 - s_1)^{\frac{1}{2+d}} \cdot K^{\frac{1+d}{2+d}} \end{aligned}$$

611 **B.3** Useful Facts about Levels $r \in \mathcal{R}$ and Blocks $[s_{\ell}(r), e_{\ell}(r)]$

The following basic facts about the level selection procedure on Line 2 of Algorithm 2 will be useful as we decompose the analysis into the blocks, or different periods of rounds, where different levels are used. The proofs all follow from Notation 1 and basic calculations.

Fact 1 (relating level to interval length). The level $r_{s_2-s_1} = 2^{-m}$ satisfies for $s_2 - s_1 \ge K$:

$$2^{-(m-1)} > \left(\frac{K}{s_2 - s_1}\right)^{\frac{1}{2+d}} \ge 2^{-m},$$

616 and hence

$$K \cdot 2^{(m-1)(2+d)} < s_2 - s_1 \le K \cdot 2^{m(2+d)}.$$

Fact 2 (the first block). *The first block* $[s_{\ell}(1), e_{\ell}(1)]$ *consists of rounds* $[t_{\ell}, t_{\ell} + K]$.

Fact 3 (start and end times of a block). For r < 1, the start time or first round $s_{\ell}(r)$ of the block corresponding to level r in episode $[t_{\ell}, t_{\ell+1})$ is $s_{\ell}(r) = t_{\ell} + \lceil K \cdot (2r)^{-(2+d)} \rceil$ and the anticipated end time or last round of the block is $e_{\ell}(r) = t_{\ell} + \lceil K \cdot r^{-(2+d)} \rceil - 1$.

Fact 4 (length of a block). Each block $[s_{\ell}(r), e_{\ell}(r)]$ is at least K rounds long. For the first block $[s_{\ell}(1), e_{\ell}(1)]$, this is already clear. Otherwise, suppose r < 1 in which case:

$$e_{\ell}(r) - s_{\ell}(r) + 1 = \left\lceil K \cdot r^{-(2+d)} \right\rceil - \left\lceil K \cdot (2r)^{-(2+d)} \right\rceil \ge K \cdot r^{-(2+d)} (1 - 2^{-(2+d)}) - 1 \ge K.$$

623 We also have the above implies

$$2 \cdot (e_{\ell}(r) - s_{\ell}(r)) \ge \frac{K \cdot r^{-(2+d)} \cdot (1 - 2^{-(2+d)})}{2}$$

Rearranging, this becomes for some constant c_{12} depending only on d:

$$c_{12}^{-1} \cdot r \le \left(\frac{K}{e_{\ell}(r) - s_{\ell}(r)}\right)^{\frac{1}{2+d}} < c_{12} \cdot r.$$

- Note we can make c_{12} large enough so that the above also holds for level r = 1.
- The above implies that the block length $e_{\ell}(r) s_{\ell}(r)$ and the episode length $e_{\ell}(r) t_{\ell}(r)$ up to the end of block $[s_{\ell}(r), e_{\ell}(r)]$ can be conflated up to constants

$$c_{13}^{-1} \cdot (e_{\ell}(r) - s_{\ell}(r)) \le e_{\ell}(r) - t_{\ell} \le c_{13} \cdot (e_{\ell}(r) - s_{\ell}(r)).$$

С Proof of Oracle Regret Bound (Proposition 2) 628

Recall that \mathcal{E}_2 is the good event on which our covariate counts concentrate by Lemma 8. It suffices to 629 show our desired regret bound for any fixed \mathbf{X}_T on this event. 630

631

Fix a phase $[\tau_i, \tau_{i+1})$ and let $r = r_{\tau_{i+1}-\tau_i}$. Fix a bin $B \in \mathcal{T}_r$ and let τ_i^a be the last round $t \in [\tau_i, \tau_{i+1})$ such that $X_t \in B$ and arm a is included in \mathcal{G}_t . If a is never excluded from \mathcal{G}_t for all such t, let $\tau_i^a \doteq \tau_{i+1} - 1$. WLOG suppose $\tau_i^1 \le \tau_i^2 \le \cdots \le \tau_i^K$. Then, letting B' be the bin at level $r_{\tau_i^a - \tau_i}$ containing covariate $X_{\tau_i^a}$, we have by (\star) that: 632 633

634

$$\sum_{t=\tau_i}^{\tau_i^a} \delta_t(a) \cdot \mathbf{1}\{X_t \in B'\} \le \sqrt{K \cdot n_{B'}([\tau_i, \tau_i^a])} + r(B') \cdot n_{B'}([\tau_i, \tau_i^a]).$$

From Lemma 9, we conclude 635

$$\sum_{t=\tau_{i}}^{\tau_{i}^{a}} \frac{\delta_{t}(a) \cdot \mathbf{1}\{X_{t} \in B\}}{|\mathcal{G}_{t}|} \leq \frac{c_{4} \left(\log^{1/2}(T) \cdot r^{d} \cdot K^{\frac{1}{2+d}} \cdot (\tau_{i+1}^{a} - \tau_{i})^{\frac{1+d}{2+d}} + K \log(T) + \sqrt{\log(T)(\tau_{i}^{a} - \tau_{i} + 1) \cdot \mu(B)} \right)}{K + 1 - a}$$
(15)

- where we use the fact that $|\mathcal{G}_t| \ge K + 1 a$ for $t \le \tau_i^a$ such that $X_t \in B$. Summing over arms $a \in [K]$ with $\sum_{a \in [K]} \frac{1}{K+1-a} \le \log(K)$, we obtain: 636
- 637

$$\sum_{a \in [K]} \sum_{t=\tau_i}^{\tau_i^*} \frac{\delta_t(a) \cdot \mathbf{1}\{X_t \in B\}}{|\mathcal{G}_t|} \le c_4 \log(K) \left(\log^{1/2}(T) r^d K^{\frac{1}{2+d}} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} + K \log(T) + \sqrt{\log(T)(\tau_i^a - \tau_i + 1) \cdot \mu(B)} \right)$$
(16)

Next, we claim that each significant phase $[\tau_i, \tau_{i+1})$ is at least K rounds long or $K \le \tau_{i+1} - \tau_i$. This follows from the definition of significant regret (\star) since for $[s_1, s_2] \subseteq [\tau_i, \tau_{i+1})$: 638

639

$$n_B([s_2, s_2]) \ge \sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1}\{X_s \in B\} \ge \sqrt{K \cdot n_B([s_1, s_2])} \implies \tau_{i+1} - \tau_i \ge n_B([s_1, s_2]) \ge K.$$

Then $K \leq \tau_{i+1} - \tau_i$ implies (via Fact 1 about the level $r_{\tau_{i+1} - \tau_i}$) 640

$$\sum_{B \in \mathcal{T}_r} K \log(T) \le K \log(T) \cdot r^{-d} \le c_{14} \log(T) K^{\frac{2}{2+d}} (\tau_{i+1} - \tau_i)^{\frac{d}{2+d}} \le c_{14} \log(T) K^{\frac{1}{2+d}} \cdot (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}}$$

Additionally, we have by Lemma 10: 641

$$\sqrt{(\tau_i^a - \tau_i) \cdot \mu(B)} \le \sqrt{(\tau_{i+1} - \tau_i) \cdot \mu(B)} \le c_8(\tau_{i+1} - \tau_i)^{\frac{1}{2+d}} K^{\frac{d/2}{2+d}} \le c_8 K^{\frac{1}{2+d}} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} K^{\frac{d}{2+d}} \le c_8 K^{\frac{1}{2+d}} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} K^{\frac{d}{2+d}}$$

Then, plugging the above into (16) and summing over bins B at level r, we have the regret in episode 642

 $[\tau_i, \tau_{i+1})$ is with probability at least $1 - 1/T^2$ w.r.t. the distribution of \mathbf{X}_T : 643

$$\mathbb{E}\left[\sum_{t=\tau_i}^{\tau_{i+1}-1} \delta_t(\pi_t) \middle| \mathbf{X}_T\right] = \mathbb{E}\left[\sum_{B \in \mathcal{T}_r} \sum_{t=\tau_i}^{\tau_{i+1}-1} \sum_{a \in \mathcal{G}_t} \frac{\delta_t(a) \cdot \mathbf{1}\{X_t \in B\}}{|\mathcal{G}_t|} \middle| \mathbf{X}_T\right]$$
$$= \mathbb{E}\left[\sum_{B \in \mathcal{T}_r} \sum_{a \in [K]} \sum_{t=\tau_i}^{\tau_i^a} \frac{\delta_t(a) \cdot \mathbf{1}\{X_t \in B\}}{|\mathcal{G}_t|} \middle| \mathbf{X}_T\right]$$
$$\leq c_{15} \log(K) \sum_{B \in \mathcal{T}_r} \log^{1/2}(T) r^d (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}} + K \log(T)$$
$$\leq c_{16} \log(K) \log(T) \cdot (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}},$$

where we use the strong density assumption to bound $\sum_{B \in \mathcal{T}_r} r^d \leq \sum_{B \in \mathcal{T}_r} c_d^{-1} \cdot \mu(B) \leq c_d^{-1}$ in the last inequality. Summing the regret over all phases $[\tau_i, \tau_{i+1})$ gives the desired result. 644 645

D **Proof of CMETA Regret Upper Bound (Theorem 3)** 646

Recall from Line 3 of Algorithm 1 that t_{ℓ} is the first round of the ℓ -th episode. WLOG, there are T 647 total episodes and, by convention, we let $t_{\ell} \doteq T + 1$ if only $\ell - 1$ episodes occurred by round T. 648

We first quickly handle the simple case of T < K. In this case, the regret bound of Theorem 3 is 649 vacuous since by the sub-additivity of $x \mapsto x^{\frac{1+d}{2+d}}$: 650

$$\sum_{i=0}^{L} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} \ge (\tau_{\tilde{L}+1} - \tau_0)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} \ge T^{\frac{1+d}{2+d}} \cdot T^{\frac{1}{2+d}} = T.$$

- Thus, it remains to show Theorem 3 for $T \ge K$. 651
- We first transform the expected regret into a more suitable form. 652

D.1 Decomposing the Regret 653

It suffices to bound $\mathbb{E}[R_T(\pi, \mathbf{X}_T) \mid \mathbf{X}_T]$ on the good event $\mathcal{E}_1 \cap \mathcal{E}_2$ where the bounds of Lemmas 8 654 and 9 hold. Going forward in the rest of the analysis, we will assume said bounds hold wherever 655 convenient. 656

We first transform the regret into a more convenient form. Let $\mathcal{F} \doteq {\mathcal{F}_t}_{t=1}^T$ be the filtration with \mathcal{F}_t 657 generated by $\{\pi_s, Y_s^{\pi_s}\}_{s=1}^{t}$. Then, 658

$$\mathbb{E}[R_T(\pi, \mathbf{X}_T) \cdot \mathbf{1}\{\mathcal{E}_1 \cap \mathcal{E}_2\} \mid \mathbf{X}_T] = \sum_{t=1}^T \mathbb{E}[\mathbb{E}[\delta_t(\pi_t) \cdot \mathbf{1}\{\mathcal{E}_1 \cap \mathcal{E}_2\} \mid \mathcal{F}_{t-1}] \mid \mathbf{X}_T$$
$$= \sum_{t=1}^T \mathbb{E}\left[\sum_{a \in \mathcal{A}_t} \frac{\delta_t(\pi_t)}{|\mathcal{A}_t|} \cdot \mathbf{1}\{\mathcal{E}_1 \cap \mathcal{E}_2\} \middle| \mathbf{X}_T\right]$$
$$= \mathbb{E}\left[\sum_{t=1}^T \sum_{a \in \mathcal{A}_t} \frac{\delta_t(a)}{|\mathcal{A}_t|} \cdot \mathbf{1}\{\mathcal{E}_1 \cap \mathcal{E}_2\} \middle| \mathbf{X}_T\right].$$

Next, as alluded to in the oracle procedure (Definition 7), until the end of a significant phase $[\tau_i, \tau_{i+1})$, 659 there is a safe arm in each bin \bar{B} at level $r_{\tau_{i+1}-\tau_i}$ which is experienced. 660

Definition 8 (local last safe arm in each phase a_t^{\sharp}). For a round $t \in [\tau_i, \tau_{i+1})$, let B be the bin at level $r_{\tau_{i+1}-\tau_i}$ which contains X_t and let $t_i(B)$ be the last round in $[\tau_i, \tau_{i+1})$ such that $X_{t_i(B)} \in B$. 661 662 Then, by Definition 6, there is a last safe arm a_t^{\sharp} which does not yet incur significant regret in bin B in the following sense: for all $[s_1, s_2] \subseteq [\tau_i, t_i(B)]$ letting $r = r_{s_2-s_1}$ and $B' \in \mathcal{T}_r$ such that 663

664 $B' \supset B$ we have: 665

$$\sum_{s=s_1}^{s_2} \delta_s(a_t^{\sharp}) \cdot \mathbf{1}\{X_s \in B'\} < \sqrt{K \cdot n_{B'}([s_1, s_2])} + r \cdot n_{B'}([s_1, s_2]).$$

Remark 4. The last safe arms $\{a_t^{\sharp}\}_t$ only depend on the distribution of X_T and **not** on the realized 666 rewards Y_T . In particular, conditional on X_T , they are fixed. 667

We first decompose the regret at round t as (a) the regret of a_t^{\dagger} and (b) the regret of arm a to the last 668 safe arm. In other words, it suffices to bound: 669

$$\mathbb{E}\left[\sum_{t=1}^{T}\sum_{a\in\mathcal{A}_{t}}\frac{\delta_{t}(a)}{|\mathcal{A}_{t}|}\cdot\mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\}\middle|\mathbf{X}_{T}\right] = \sum_{t=1}^{T}\delta_{t}(a_{t}^{\sharp})\cdot\mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\} + \mathbb{E}\left[\sum_{t=1}^{T}\sum_{a\in\mathcal{A}_{t}}\frac{\delta_{t}(a_{t}^{\sharp},a)}{|\mathcal{A}_{t}|}\cdot\mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\}\middle|\mathbf{X}_{T}\right]$$

Note that the expectation on the first sum disappears since a_t^{\sharp} is only a function of \mathbf{X}_T and the mean 670 reward functions $\{f_t^a(\cdot)\}_{t,a}$. 671

D.2 Bounding the Regret of the Last Safe Arm 672

Bounding $\sum_{t=1}^{T} \delta_t(a_t^{\sharp})$ will be similar to the proof of Proposition 2. We essentially show that the 673 oracle procedure could have also just played arm a_t^{\sharp} every round. 674

Fix a phase $[\tau_i, \tau_{i+1})$ and let $r = r_{\tau_{i+1}-\tau_i}$. Fix a bin $B \in \mathcal{T}_r$ and let $a_i(B)$ be the last safe arm a_t^{\sharp} of 675

the last round $t \in [\tau_i, \tau_{i+1})$ such that $X_t \in B$. Then, $a_t^{\sharp} = a_i(B)$ for every round $t \in [\tau_i, \tau_{i+1})$ such that $X_t \in B$. Then, we have by Definition 6 that for bin $B' \supseteq B$ at level $r_{t-\tau_i}$: 676

677

$$\sum_{s=\tau_i}^{\iota} \delta_s(a_i(B)) \cdot \mathbf{1}\{X_s \in B'\} \le \sqrt{K \cdot n_{B'}([\tau_i, t])} + r(B') \cdot n_{B'}([\tau_i, t]) \cdot n_{B'}([$$

Then, by Lemma 9, we have: 678

τ

$$\sum_{s=\tau_i}^t \delta_s(a_i(B)) \cdot \mathbf{1}\{X_s \in B\} \le c_4 \left(\log^{1/2}(T) \cdot r^d \cdot (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} + K \log(T) + \sqrt{\log(T)(t - \tau_i + 1) \cdot \mu(B)} \right)$$
(17)

Then, summing the above over bins in the same fashion as the proof of Proposition 2 gives: 679

$$\sum_{t=\tau_i}^{\tau_{i+1}-1} \delta_t(a_t^{\sharp}) = \sum_{B \in \mathcal{T}_r} \sum_{s=\tau_i}^{\tau_{i+1}-1} \delta_s(a_i(B)) \cdot \mathbf{1}\{X_s \in B\} \le c_3 \log(T) \cdot (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}.$$

Finally, summing over phases $[\tau_i, \tau_{i+1})$ we have $\sum_{t=1}^T \delta_t(a_t^{\sharp})$ is of the right order. 680

D.3 Relating Episodes to Significant Phases 681

We next show that w.h.p. a restart occurs (i.e., a new episode begins) only if a significant shift has 682 occurred sometime within the episode. Recall from Definition 6 that $\tau_1, \tau_2, \ldots, \tau_{\tilde{L}}$ are the times of 683 the significant shifts and that t_1, \ldots, t_T are the episode start times. 684

Lemma 11 (**Restart Implies Significant Shift**). On event \mathcal{E}_1 , for each episode $[t_{\ell}, t_{\ell+1})$ with $t_{\ell+1} \leq t_{\ell+1}$ 685 T (i.e., an episode which concludes with a restart), there exists a significant shift $\tau_i \in [t_{\ell}, t_{\ell+1})$. 686

Proof. Fix an episode $[t_{\ell}, t_{\ell+1}]$. Then, by Line 11 of Algorithm 1, there is a bin B such that every 687 arm $a \in [K]$ was evicted from B at some round in the episode, i.e. (5) is true for each arm a on 688 some interval $[s_1, s_2] \subseteq [t_\ell, t_{\ell+1})$. It suffices to show that this implies a significant shift has occurred 689 between rounds t_{ℓ} and $t_{\ell+1}$. 690

Suppose (5) first triggers the eviction of arm a at time t in $B' \supseteq B$ over interval $[s_1, s_2]$ where 691 $r(B') = r_{s_2-s_1}$. By concentration (9) and our eviction criteria (5), we have that there is an arm $a' \neq a$ such that (using the notation of Proposition 7) for large enough $C_0 > 0$ and some $c_{17} > 0$: 692 693

$$\sum_{s=s_1}^{s_2} \mathbb{E}\left[\hat{\delta}_s^B(a',a) \mid \mathcal{F}_{s-1}\right] \ge c_{17}\log(T) \left(\sqrt{K \cdot n_{B'}([s_1,s_2]) + K^2} + r(B') \cdot n_{B'}([s_1,s_2])\right).$$
(18)

Next, if arm a is evicted from $\mathcal{A}(B')$ at round t, then we have by the definition of $\hat{\delta}_s^{B'}(a', a)$ (4): 694

$$\mathbb{E}[\hat{\delta}_{s}^{B'}(a',a) \mid \mathcal{F}_{s-1}] = \begin{cases} \delta_{s}(a',a) \cdot \mathbf{1}\{X_{s} \in B'\} & a,a' \in \mathcal{A}_{s} \\ -f_{s}^{a}(X_{s}) \cdot \mathbf{1}\{X_{s} \in B\} & a \in \mathcal{A}_{s}, a' \notin \mathcal{A}_{s} \\ 0 & a \notin \mathcal{A}_{s} \end{cases}$$

In any case, the above L.H.S. conditional expectation is bounded above by $\delta_s(a) \cdot \mathbf{1}\{X_s \in B'\}$. Thus, 695 (18) implies arm a incurs significant regret (\star) in B' on $[s_1, s_2]$: 696

$$\sum_{s=s_1}^{s_2} \delta_s(a) \cdot \mathbf{1}\{X_s \in B'\} \ge \sqrt{K \cdot n_{B'}([s_2, s_2])} + r(B') \cdot n_{B'}([s_1, s_2]).$$

Then, since every arm a is evicted in bin B by round t, a significant shift must have occurred between 697 rounds t_{ℓ} and $t_{\ell+1}$. 698

D.4 Regret of CMETA to the Last Safe Arm 699

It remains to bound $\mathbb{E}[\sum_{t=1}^{T} \sum_{a \in \mathcal{A}_t} \delta_t(a_t^{\sharp}, a) / |\mathcal{A}_t| | \mathbf{X}_t]$. We further decompose this sum over t into episodes and then *blocks* where a particular choice of level is used within the episode. The following 700 701 notation will be useful. 702

Definition 9. Let $s_{\ell}(r)$ and $e_{\ell}(r)$ denote the first and last rounds when level r is used by the master 703 Base-Alg in episode $[t_{\ell}, t_{\ell+1})$, i.e. rounds $t \in [t_{\ell}, t_{\ell+1})$ such that $r_{t-t_{\ell}} = r$. We call $[s_{\ell}(r), e_{\ell}(r)]$ a 704 **block**. Let $PHASES(\ell, r) \doteq \{i \in [\tilde{L}] : [\tau_i, \tau_{i+1}) \cap [s_\ell(r), e_\ell(r)] \neq \emptyset\}$ be the phases which intersect block $[s_\ell(r), e_\ell(r)]$, let $T(i, r, \ell) \doteq |[\tau_i, \tau_{i+1}) \cap [s_\ell(r), e_\ell(r)]|$ be the effective length of the phase as 705 706 observed in block $[s_{\ell}(r), e_{\ell}(r)]$. 707

Similarly, define PHASES $(\ell) \doteq \{i \in [\tilde{L}] : [\tau_i, \tau_{i+1}) \cap [t_\ell, t_{\ell+1}) \neq \emptyset\}$ be the phases which intersect 708 episode $[t_{\ell}, t_{\ell+1})$. 709

It will in fact suffice to show w.h.p. w.r.t. the distribution of \mathbf{X}_T , for each episode $[t_{\ell}, t_{\ell+1})$, each 710 block $[s_{\ell}(r), e_{\ell}(r)]$ in $[t_{\ell}, t_{\ell+1})$, and each bin $B \in \mathcal{T}_r$: 711

$$\mathbb{E}\left[\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)} \sum_{a\in\mathcal{A}_{t}} \frac{\delta_{t}(a_{t}^{\sharp},a)}{|\mathcal{A}_{t}|} \cdot \mathbf{1}\{X_{t}\in B\} \cdot \mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\} \left| \mathbf{X}_{T} \right] \\ \leq c_{18}\log^{3}(T)\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\} \left(\log(T) + \sum_{i\in\mathsf{PHASES}(\ell,r)} r(B)^{d} \cdot T(i,r,\ell)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}\right) \left| \mathbf{X}_{T} \right] \right]$$

$$(19)$$

D.5 Summing the Per-(Bin,Block,Episode) Regret over Bins, Blocks, and Episodes. 712

Admitting (19), we show that the total dynamic regret over T rounds is of the desired order. 713

Recall from earlier that there are WLOG T total episodes with the convention that $t_{\ell} \doteq T + 1$ if only 714 ℓ episodes occur by round T. Then, summing our per-bin regret bound (19) over all the bins at level r gives (using strong density to bound $\sum_{B \in r} r^d \leq \frac{C_d}{c_d}$): 715

716

$$\mathbb{E}\left[\sum_{B\in\mathcal{T}_{r}}\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)}\sum_{a\in\mathcal{A}_{t}}\frac{\delta_{t}(a_{t}^{\sharp},a)}{|\mathcal{A}_{t}|}\cdot\mathbf{1}\{X_{t}\in B\}\cdot\mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\}\middle|\mathbf{X}_{T}\right] \\
\leq c_{18}\log^{3}(T)\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_{1}\cap\mathcal{E}_{2}\}\left(\sum_{B\in\mathcal{T}_{r}}\log(T)+\sum_{i\in\mathsf{PHASES}(\ell,r)}T(i,r,\ell)^{\frac{1+d}{2+d}}\cdot K^{\frac{1}{2+d}}\right)\middle|\mathbf{X}_{T}\right]. \quad (20)$$

Next, summing over the different levels r (of which there are at most $\log(T)$ used in any episode), 717 we obtain by Jensen's inequality: 718

$$\begin{split} \sum_{r \in \mathcal{R}} \sum_{i \in \mathsf{PHASES}(\ell, r)} T(i, r, \ell)^{\frac{1+d}{2+d}} &= \sum_{i \in \mathsf{PHASES}(\ell)} \sum_{r \in \mathcal{R}: i \in \mathsf{PHASES}(\ell, r)} T(i, r, \ell)^{\frac{1+d}{2+d}} \\ &\leq \sum_{i \in \mathsf{PHASES}(\ell)} \left(\log(T) \sum_{r \in \mathcal{R}: i \in \mathsf{PHASES}(\ell, r)} T(i, r, \ell) \right)^{\frac{1+d}{2+d}}. \end{split}$$

Now, we have 719

$$\sum_{r \in \mathcal{R}: i \in \mathsf{PHASES}(\ell, r)} T(i, r, \ell) = \sum_{r \in \mathcal{R}: i \in \mathsf{PHASES}(\ell, r)} |[\tau_i, \tau_{i+1}) \cap [s_\ell(r), e_\ell(r)]| = \tau_{i+1} - \tau_i + 1.$$

We also have (via Fact 1 about level $r_{t_{\ell+1}-t_{\ell}}$ which is the smallest level used in episode $[t_{\ell}, t_{\ell+1})$). 720

$$\sum_{r \in \mathcal{R}} \sum_{B \in \mathcal{T}_r} \log(T) \leq \sum_{r \in \mathcal{R}} r^{-d} \cdot \log(T)$$
$$\leq c_{19} \log^2(T) \left(\frac{t_{\ell+1} - t_{\ell}}{K}\right)^{\frac{d}{2+d}}$$
$$\leq c_{20} \log^2(T) \sum_{i \in \mathsf{PHASES}(\ell)} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}.$$

Thus, combining the above inequalities with (20), we obtain overall bound: 721

$$c_{18}\log^4(T)\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_1 \cap \mathcal{E}_2\}\sum_{i \in \mathsf{PHASES}(\ell)} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}\right].$$

Recall now that \mathcal{E}_1 is the good event over which the concentration bounds of Proposition 7 hold. Then, 722 using the fact that, on event \mathcal{E}_1 , each phase $[\tau_i, \tau_{i+1})$ intersects at most two episodes (Lemma 11), 723 summing the above R.H.S over episodes $\ell \in [T]$ gives us (since at most $\log(T)$ blocks per episode) 724 order 725

$$2\log^4(T)\sum_{i=1}^{\bar{L}}(\tau_{i+1}-\tau_i)^{\frac{1+d}{2+d}}\cdot K^{\frac{1}{2+d}}.$$

It then remains to show (19). 726

D.6 Bounding the Per-Bin Per-Block Regret to the Last Safe Arm 727

To show (19), we first fix a block $[s_{\ell}(r), e_{\ell}(r)]$ and a bin $B \in \mathcal{T}_r$. We then further decompose 728 $\delta_t(a_t^{\sharp}, a)$ in two parts: 729

(a) The regret of a to the last local arm, denoted by $a_r(B)$, to be evicted from $\mathcal{A}_{master}(B)$ in block 730 $[s_{\ell}(r), e_{\ell}(r)]$ (ties are broken arbitrarily). 731

- (b) The regret of the last local arm $a_r(B)$ to the last safe arm a_t^{\sharp} . 732
- In other words, the L.H.S. of (19) is decomposed as: 733

$$\underbrace{\mathbb{E}\left[\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)}\sum_{a\in\mathcal{A}_{t}}\frac{\delta_{t}(a_{r}(B),a)}{|\mathcal{A}_{t}|}\cdot\mathbf{1}\{X_{t}\in B\}\middle|\mathbf{X}_{T}\right]}_{(a)}+\underbrace{\mathbb{E}\left[\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)}\delta_{t}(a_{t}^{\sharp},a_{r}(B))\cdot\mathbf{1}\{X_{t}\in B\}\middle|\mathbf{X}_{T}\right]}_{(b)}.$$

We will show both (a) and (b) are of order (19). 734

744

• Bounding the Regret of Other Arms to the Last Local Arm $a_r(B)$. We start by partitioning 735 the rounds t such that $X_t \in B$ and $a \in A_t$ in (a) according to before or after they are evicted from 736 $\mathcal{A}_{\text{master}}(B)$. Suppose arm a is evicted from $\mathcal{A}_{\text{master}}(B)$ at round $t_r^a \in [s_\ell(r), e_\ell(r)]$ (formally, we let 737 $t_r^a := e_\ell(r)$ if a is not evicted in block $[s_\ell(r), e_\ell(r)]$). Then, it suffices to bound: 738

$$\mathbb{E}\left[\sum_{a=1}^{K}\sum_{t=s_{\ell}(r)}^{t_{r}^{a}-1}\frac{\delta_{t}(a_{r}(B),a)}{|\mathcal{A}_{t}|}\cdot\mathbf{1}\{X_{t}\in B\}+\sum_{a=1}^{K}\sum_{t=t_{r}^{a}}^{e_{\ell}(r)}\frac{\delta_{t}(a_{r}(B),a)}{|\mathcal{A}_{t}|}\cdot\mathbf{1}\{a\in\mathcal{A}_{t}\}\cdot\mathbf{1}\{X_{t}\in B\}\Big|\mathbf{X}_{T}\right]$$
(21)

Suppose WLOG that $t_r^1 \le t_r^2 \le \cdots \le t_r^K$. Then, for each round $t < t_r^a$ all arms $a' \ge a$ are retained in 739 $\mathcal{A}_{\text{master}}(B)$ and thus retained in the candidate arm set \mathcal{A}_t for all rounds t where $X_t \in B$. Importantly, 740 at each round t a level of at least r is used since a child Base-Alg can only use a higher level than the 741 master Base-Alg. Thus, $|\mathcal{A}_t| \geq K + 1 - a$ for all $t \leq t_r^a$. 742

Next, we bound the first double sum in (21), i.e. the regret of playing a to $a_r(B)$ from $s_\ell(r)$ to 743 t_r^a . Applying our concentration bounds (Proposition 7), since arm a is not evicted from $\mathcal{A}(B)$ till

- round t_r^a , on event \mathcal{E}_1 we have for some $c_5 > 0$ and any other arm $a' \in \mathcal{A}(B)$ through round $t_r^a 1$
- (i.e., $a' \in \mathcal{A}_t$ for all $t \in [t_\ell, t_r^a)$ such that $X_t \in B$ since we always use level at least r at such a round t): for bin $B' \supseteq B$ at level $r_{t_r^a 1 s_\ell(r)}$: on event \mathcal{E}_1 (note that we necessarily always have
- 748 $\mathcal{A}(B') \supseteq \mathcal{A}(B)$ for $B' \supseteq B$):

$$\sum_{t=s_{\ell}(r)}^{t_{r}^{*}-1} \mathbb{E}[\hat{\delta}_{s}^{B'}(a',a) \mid \mathcal{F}_{t-1}] \leq c_{5} \log(T) \sqrt{K \cdot n_{B'}([s_{\ell}(r),t_{r}^{a})) \vee K^{2}} + r(B') \cdot n_{B'}([s_{\ell}(r),t_{r}^{a})).$$

Next, since $a, a' \in \mathcal{A}_t$ for each $t \in [s_\ell(r), t_r^a - 1)$ such that $X_t \in B$, we have:

$$\forall t \in [s_{\ell}(r), t_r^a), X_t \in B : \mathbb{E}[\hat{\delta}_t^B(a', a) \mid \mathcal{F}_{t-1}] = \delta_t(a', a).$$

750 Thus, we conclude

$$\sum_{t=s_{\ell}(r)}^{t_r^a - 1} \delta_t(a', a) \cdot \mathbf{1}\{X_t \in B\} \le c_5 \log(T) \sqrt{K \cdot n_{B'}([s_{\ell}(r), t_r^a)) \vee K^2} + r(B') \cdot n_{B'}([s_{\ell}(r), t_r^a)).$$

Thus, by Lemma 9, and since $B' \supseteq B$, we conclude for any such a' on event \mathcal{E}_1 :

$$\sum_{t=s_{\ell}(r)}^{t_{r}^{a}-1} \frac{\delta_{t}(a',a)}{|\mathcal{A}_{t}|} \cdot \mathbf{1}\{X_{t} \in B\} \leq \frac{c_{5}\left(\log^{1/2}(T)r^{d} \cdot K^{\frac{1}{2+d}} \cdot (t_{r}^{a} - s_{\ell}(r))^{\frac{1+d}{2+d}} + K\log(T) + \sqrt{\log(T)(\tau_{i}^{1} - \tau_{i} + 1) \cdot \mu(B)}\right)}{K+1-a}$$
(22)

where we use the fact that $|\mathcal{A}_t| \geq K + 1 - a$ for all $t \in [s_\ell(r), t_r^a)$. Since this last bound holds

- uniformly for all $a' \in \mathcal{A}(B)$ through round $t_r^a 1$, it must hold for the last master arm $a_r(B)$.
- Then, summing over all arms a, we have on event \mathcal{E}_1 :

$$\sum_{a=1}^{K} \sum_{t=s_{\ell}(r)}^{t_{r}^{a}-1} \frac{\delta_{t}(a_{r}(B),a)}{|\mathcal{A}_{t}|} \cdot \mathbf{1}\{X_{t} \in B\} \leq c_{5}\log(K) \left(\log^{1/2}(T) \cdot r^{d} \cdot (e_{\ell}(r) - s_{\ell}(r))^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} + K^{\frac{1}$$

755 Note that by Lemma 10:

$$\sqrt{(t_r^a - s_\ell(r)) \cdot \mu(B)} \le \sqrt{(e_\ell(r) - s_\ell(r)) \cdot \mu(B)} \le c_8 K^{\frac{d/2}{2+d}} (e_\ell(r) - s_\ell(r))^{\frac{1}{2+d}} \le c_8 K^{\frac{1}{2+d}} \cdot r^d \cdot (e_\ell(r) - s_\ell(r))^{\frac{1+d}{2+d}}$$

Thus, it suffices to consider the RHS above as our bound

Thus, it suffices to consider the RHS above as our bound.

Next, we handle the second double sum in (21). We first observe that if arm *a* is played in bin *B* after round t_r^a , then it must due to an active replay. The difficulty here is that replays may interrupt each other and so care must be taken in managing the contribution of $\sum_t \delta_t(a_r(B), a)$ (which may be negative) by different overlapping replays.

Our strategy, identical to that of Section B.1 in Suk and Kpotufe [2022], is to partition the rounds when a is played by a replay after round t_r^a according to which replay is active and not accounted for by another replay. This involves carefully designating a subclass of replays whose durations while playing a in B span all the rounds where a is played in B after t_r^a . Then, we cover the times when ais played by a collection of intervals corresponding to the schedules of this subclass of replays, on each of which we can employ the eviction criterion (5) and concentration like before.

For this purpose, we define the following terminology (which is all w.r.t. a fixed arm a):

768 Definition 10.

- (i) For each scheduled and activated Base-Alg (s, m), let the round M(s, m) be the minimum of two quantities: (a) the last round in [s, s + m] when arm a is retained in $\mathcal{A}(B)$ by Base-Alg (s, m) and all of its children, and (b) the last round that Base-Alg (s, m) is active and not permanently interrupted. Call the interval [s, M(s, m)] the active interval of Base-Alg (s, m).
- (ii) Call a replay Base-Alg (s, m) proper if there is no other scheduled replay Base-Alg (s', m')such that $[s, s + m] \subset (s', s' + m')$ where Base-Alg (s', m') will become active again after round s + m. In other words, a proper replay is not scheduled inside the scheduled range of rounds of another replay. Let PROPER $(s_{\ell}(r), e_{\ell}(r))$ be the set of proper replays scheduled to start in the block $[s_{\ell}(r), e_{\ell}(r)]$.



subproper replays reintroduce arm a to \mathcal{A}_t

Figure 1: Shown are replay scheduled durations (in gray) with dots marking when arm a is reintroduced to A_t . Black segments indicate the period [s, M(s, m)] for proper and subproper replays. Note that the rounds where $a \in A_t$ in the left unlabeled replay's duration are accounted for by the larger proper replay.

779	(iii)	Call a scheduled replay Base-Alg (s, m) subproper if it is non-proper and if each of its
780		ancestor replays (i.e., previously scheduled replays whose durations have not concluded)
781		Base-Alg (s', m') satisfies $M(s', m') < s$. In other words, a subproper replay either
782		permanently interrupts its parent or does not, but is scheduled after its parent (and all
783		its ancestors) stops playing arm a in B. Let SUBPROPER $(s_{\ell}(r), s_{\ell}(r))$ be the set of all
784		subproper replays scheduled before round $t_{\ell+1}$.

Equipped with this language, we now show some basic claims which essentially reduce analyzing the complicated hierarchy of replays to analyzing the active intervals of replays in PROPER $(s_{\ell}(r), e_{\ell}(r)) \cup$ SUBPROPER $(s_{\ell}(r), s_{\ell}(r))$.

788 **Proposition 12.** The active intervals

$$\{[s, M(s, m)] : \mathsf{Base-Alg}\,(s, m) \in \mathsf{Proper}(s_\ell(r), e_\ell(r)) \cup \mathsf{SubProper}(s_\ell(r), s_\ell(r))\}, \\$$

789 are mutually disjoint.

Proof. Clearly, the classes of replays $PROPER(t_{\ell}, t_{\ell+1})$ and $SUBPROPER(s_{\ell}(r), s_{\ell}(r))$ are disjoint. Next, we show the respective active intervals [s, M(s, m)] and [s', M(s', m')] of any two Base-Alg (s, m) and Base-Alg $(s', m') \in PROPER(s_{\ell}(r), e_{\ell}(r)) \cup SUBPROPER(s_{\ell}(r), s_{\ell}(r))$ are disjoint.

- 1. Proper replay vs. subproper replay: a subproper replay can only be scheduled after the round M(s,m) of the most recent proper replay Base-Alg (s,m) (which is necessarily an ancestor). Thus, the active intervals of proper replays and subproper replays.
- 797 2. Two distinct proper replays: two such replays can only permanently interrupt each other, 798 and since M(s,m) always occurs before the permanent interruption of Base-Alg (s,m), we 799 have the active intervals of two such replays are disjoint.
- 3. Two distinct subproper replays: consider two non-proper replays 800 Base-Alg (s, m), Base-Alg $(s', m') \in$ SUBPROPER $(s_{\ell}(r), s_{\ell}(r))$ with s' > s. The only 801 way their active intervals intersect is if Base-Alg (s, m) is an ancestor of Base-Alg (s', m'). 802 Then, if Base-Alg (s', m') is subproper, we must have s' > M(s, m), which means that 803 [s', M(s', m')] and [s, M(s, m)] are disjoint. 804

805

Next, we claim that the active intervals [s, M(s, m)] for Base-Alg $(s, m) \in \text{PROPER}(t_{\ell}, t_{\ell+1}) \cup$ SUBPROPER $(s_{\ell}(r), s_{\ell}(r))$ contain all the rounds where a is played in B after being evicted from $\mathcal{A}_{\text{master}}(B)$. To show this, we first observe that for each round t when a replay is active, there is a unique proper replay associated to t, namely the proper replay scheduled most recently. Next, note that any round $t > t_r^a$ where $X_t \in B$ and where arm $a \in \mathcal{A}_t$ must belong to the active interval [s, M(s, m)] of the unique proper replay Base-Alg (s, m) associated to round t, or else satisfies t > M(s, m) in which case a unique subproper replay Base-Alg $(s', m') \in \text{SUBPROPER}(s_{\ell}(r), s_{\ell}(r))$

- was active and not yet permanently interrupted by round t. Thus, it must be the case that $t \in [s', M(s', m')]$.
- Overloading notation, we'll let $\mathcal{A}_t(B)$ be the value of $\mathcal{A}(B)$ for the Base-Alg active at round t. Next,
- note that every round $t \in [s, M(s, m)]$ for a proper or subproper Base-Alg (s, m) is clearly a round
- where $a \in \mathcal{A}_t(B)$ and no such round is accounted for twice by Proposition 12. Thus,

$$\{t \in (t_r^a, e_\ell(r)] : a \in \mathcal{A}_t(B)\} = \bigsqcup_{\mathsf{Base-Alg}\,(s,m) \in \mathsf{Proper}(s_\ell(r), e_\ell(r)) \cup \mathsf{SubProper}(s_\ell(r), s_\ell(r))} [s, M(s, m)].$$

818 Then, we can rewrite the second double sum in (21) as:

$$\sum_{a=1}^{K} \sum_{\text{Base-Alg}(s,m)\in \text{PROPER}(s_{\ell}(r),e_{\ell}(r))\cup \text{SUBPROPER}(s_{\ell}(r),s_{\ell}(r))} Z_{m,s} \cdot \sum_{t=s \lor t_r^a}^{M(s,m)} \frac{\delta_t(a_r(B),a)}{|\mathcal{A}_t|} \cdot \mathbf{1}\{X_t \in B\}.$$

Recall in the above that the Bernoulli $Z_{m,s}$ (see Line 6 of Algorithm 1) decides whether Base-Alg(s,m) is scheduled.

Further bounding the sum over t above by its positive part, we can expand the sum over Base-Alg $(s,m) \in \text{PROPER}(t_{\ell}, t_{\ell+1}) \cup \text{SUBPROPER}(s_{\ell}(r), s_{\ell}(r))$ to be over all Base-Alg (s,m),

823 or obtain:

$$\sum_{a=1}^{K} \sum_{\text{Base-Alg}\,(s,m)} Z_{m,s} \cdot \left(\sum_{t=s \lor t_r^a}^{M(s,m)} \frac{\delta_t(a_r(B),a)}{|\mathcal{A}_t|} \cdot \mathbf{1}\{X_t \in B\} \right)_+,$$

where the sum is over all replays Base-Alg (s, m), i.e. $s \in \{t_{\ell} + 1, \ldots, t_{\ell+1} - 1\}$ and $m \in \{2, 4, \ldots, 2^{\lceil \log(T) \rceil}\}$. It then remains to bound the contributed relative regret of each Base-Alg (s, m) in the interval $[s \lor t_x^a, M(s, m)]$, which will follow similarly to the previous steps.

We first have using similar arguments as before (now overloading the notation M(s,m) as M(s,m,a)

for clarity), i.e. combining our concentration bound (9) with the eviction criterion (5) and applying Lemma 9:

$$\sum_{t=s\vee t_r^a}^{M(s,m)} \frac{\delta_t(a_r(B),a)}{|\mathcal{A}_t|} \cdot \mathbf{1}\{X_t \in B\} \le \frac{c_5\left(\log^{1/2}(T) \cdot r^d \cdot K^{\frac{1}{2+d}} \cdot m^{\frac{1+d}{2+d}} + K\log(T) + \sqrt{\log(T)(M(s,m) - s)\mu(B)}\right)}{\min_{t \in [s,M(s,m,a)]} |\mathcal{A}_t|}$$

830 Thus, it remains to bound

$$\sum_{a=1}^{K} \sum_{\text{Base-Alg}\,(s,m)} Z_{m,s} \cdot \left(\frac{c_5 \left(\log^{1/2}(T) \cdot r^d \cdot K^{\frac{1}{2+d}} \cdot m^{\frac{1+d}{2+d}} + K \log(T) + \sqrt{\log(T)(M(s,m) - s) \cdot \mu(B)} \right)}{\min_{t \in [s,M(s,m,a)]} |\mathcal{A}_t|} \right)$$

Swapping the outer two sums and recognizing that $\sum_{a=1}^{K} \frac{1}{\min_{t \in [s, M(s,m,a)]} |\mathcal{A}_t|} \le \log(K)$ by similar arguments to before by summing over arms in the order they are evicted by Base-Alg (s, m), we have

arguments to before by summing over arms in the order they are evicted by Base-Alg (s, m), we that it remains to bound

$$\sum_{\text{Base-Alg}\,(s,m)} Z_{m,s} \cdot c_5\left(\log^{1/2}(T) \cdot r^d \cdot K^{\frac{1}{2+d}} \cdot \tilde{m}^{\frac{1+d}{2+d}} + K\log(T) + \sqrt{\log(T)(m-s)\mu(B)}\right),\tag{23}$$

where $\tilde{m} \doteq m \land (e_{\ell}(r) - s_{\ell}(r))$ (note we may restrict attention to the part of replays in the current block $[s_{\ell}(r), e_{\ell}(r)]$). Let

$$R(m,B) \doteq \left(c_5 \left(\log^{1/2}(T) \cdot r^d \cdot K^{\frac{1}{2+d}} \cdot \tilde{m}^{\frac{1+d}{2+d}} + K \log(T) + \sqrt{\log(T) \cdot \tilde{m} \cdot r^d} \right) \right) \wedge n_B([s,s+m]),$$

Then, in light of the previous calculations, R(m, B) is an upper bound on the within-bin B regret

contributed by a replay of total duration m (note we can always coarsely upper bound this regret by $n_B([s, s + m])$).

Then, plugging R(m, B) into (23) gives via tower law (we remove the "conditional on X_T " part for ease of presentation):

$$\mathbb{E}\left[\mathbb{E}\left[\left|\sum_{\mathsf{Base-Alg}\,(s,m)} Z_{m,s} \cdot R(m,B) \middle| s_{\ell}(r)\right]\right] = \mathbb{E}\left[\left|\sum_{s=s_{\ell}(r)}^{T} \sum_{m} \mathbb{E}[Z_{m,s} \cdot \mathbf{1}\{s \le e_{\ell}(r)\} \mid s_{\ell}(r)] \cdot R(m,B)\right]\right]$$

Next, we observe that $Z_{m,s}$ and $\mathbf{1}\{s \le e_{\ell}(r)\}$ are independent conditional on t_{ℓ} since $\mathbf{1}\{s \le e_{\ell}(r)\}$ only depends on the scheduling and observations of base algorithms scheduled before round s. Thus,

recalling that
$$\mathbb{P}(Z_{m,s}=1) = \mathbb{E}[Z_{m,s} \mid t_{\ell}] = \left(\frac{1}{m}\right)^{\frac{1}{2+d}} \cdot \left(\frac{1}{s-t_{\ell}}\right)^{\frac{1}{2+d}},$$

 $\mathbb{E}[Z_{m,s} \cdot \mathbf{1}\{s \le e_{\ell}(r)\} \mid t_{\ell}] = \mathbb{E}[Z_{m,s} \mid t_{\ell}] \cdot \mathbb{E}[\mathbf{1}\{s \le e_{\ell}(r)\} \mid s_{\ell}(r)]$
 $= \left(\frac{1}{m}\right)^{\frac{1}{2+d}} \cdot \left(\frac{1}{s-t_{\ell}}\right)^{\frac{1+d}{2+d}} \cdot \mathbb{E}[\mathbf{1}\{s \le e_{\ell}(r)\} \mid s_{\ell}(r)].$

⁸⁴⁴ Plugging this into our expectation from before and unconditioning, we obtain:

$$\mathbb{E}\left[\sum_{s=s_{\ell}(r)}^{\lceil \log(T)\rceil} \sum_{n=1}^{\lceil \log(T)\rceil} \left(\frac{1}{2^n}\right)^{\frac{1}{2+d}} \left(\frac{1}{s-t_{\ell}}\right)^{\frac{1+d}{2+d}} \cdot R(2^n, B)\right]$$
(24)

845 We first evaluate the inner sum over n. Note that

843

$$\sum_{n=1}^{\lceil \log(T) \rceil} \left(\frac{1}{2^n}\right)^{\frac{1}{2+d}} \cdot \left(2^n \wedge (e_{\ell}(r) - s_{\ell}(r))^{\frac{1+d}{2+d}} \le \log(T) \cdot (e_{\ell}(r) - s_{\ell}(r))^{\frac{d}{2+d}}$$

$$\sum_{n=1}^{\lceil \log(T) \rceil} \left(\frac{1}{2^n}\right)^{\frac{1}{2+d}} \sqrt{2^n \wedge (e_{\ell}(r) - s_{\ell}(r))} \le (e_{\ell}(r) - s_{\ell}(r))^{\frac{d/2}{2+d}}$$

$$\sum_{n=1}^{\lceil \log(T) \rceil} \left(\frac{1}{2^n}\right)^{\frac{1}{2+d}} (K \wedge 2^n) \le \log(T) \cdot K^{\frac{1+d}{2+d}}.$$

Multiplying by $(s - t_{\ell})^{-\frac{1+d}{2+d}}$ and taking a further sum over $s \in [s_{\ell}(r), e_{\ell}(r)]$ in the above display, (24) becomes

$$(e_{\ell}(r) - t_{\ell})^{\frac{1}{2+d}} \left((e_{\ell}(r) - s_{\ell}(r))^{\frac{d}{2+d}} K^{\frac{1}{2+d}} \cdot r^{d} + (e_{\ell}(r) - s_{\ell}(r))^{\frac{d/2}{2+d}} \sqrt{\log(T) \cdot r^{d}} + K^{\frac{1+d}{2+d}} \log(T) \right).$$

We have the first term inside the parametheses above inside dominates the second term as long as $K \ge \log(T)$.

Next, note from Fact 4 that $e_{\ell}(r) - t_{\ell} \leq c_{13}(e_{\ell}(r) - s_{\ell}(r))$ and so the above is at most:

$$r^{d} \cdot (e_{\ell}(r) - s_{\ell}(r))^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}} + \log(T) K^{\frac{1+d}{2+d}} \cdot (e_{\ell}(r) - s_{\ell}(r))^{\frac{1}{2+d}}.$$
 (25)

We next recall from Fact 4 that each block $[s_{\ell}(r), e_{\ell}(r)]$ is at least K rounds long. Thus,

$$C_d \cdot r^d \cdot (e_\ell(r) - s_\ell(r))^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} \ge c_{21} \cdot (e_\ell(r) - s_\ell(r))^{\frac{1}{2+d}} \cdot K^{\frac{1+d}{2+d}}$$

- Thus, the second term of (25) is at most $\log(T)$ times the first term.
- Showing (a) is order (19) then follows from upper bounding $e_{\ell}(r) s_{\ell}(r)$ by the combined length of
- all phases $[\tau_i, \tau_{i+1})$ intersecting block $[s_\ell(r), e_\ell(r)]$, and using the sub-additivity of $x \mapsto x^{\frac{1+d}{2+d}}$.

• Bounding the Regret of the Last Master Arm $a_r(B)$ to the Last Safe Arm a_t^{\sharp} . Before we proceed, we first convert $\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)} \delta_t(a_t^{\sharp}, a_r(B)) \cdot \mathbf{1}\{X_t \in B\}$ into a more convenient form in terms of the bin-masses $\mu(B)$. By concentration (11) of Proposition 7, we have

$$\sum_{t} \delta_t(a_t^{\sharp}, a_r(B)) \cdot \mathbf{1}\{X_t \in B\} \le \sum_{t} \delta_t(a_t^{\sharp}, a_r(B)) \cdot \mu(B) + c_1 \left(\log(T) + \sqrt{\log(T)(e_\ell(r) - s_\ell(r)) \cdot \mu(B)}\right).$$

858 By Lemma 10, we have

$$\sqrt{(e_{\ell}(r) - s_{\ell}(r)) \cdot \mu(B)} \le r^{d} \cdot (e_{\ell}(r) - s_{\ell}(r))^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}}.$$

Additionally, $\log(T)$ is of the right order with respect to (20). Thus, the concentration error terms

from Proposition 7 above are negligible.

Moving forward, by the strong density assumption and in light of (19), it suffices to show 861

$$\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)} \delta_t(a_t^{\sharp}, a_r(B)) \lesssim \sum_{i \in \mathsf{PHASES}(\ell, r)} (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}}$$

This is the most difficult quantity to bound since arm a_t^{\sharp} may have been evicted from $\mathcal{A}_{\text{master}}(B)$ 862 before round t and, thus, we rely on our replay scheduling to bound the regret incurred while waiting 863

to detect a large aggregate value of $\delta_t(a_t^{\sharp}, a_r(B))$. 864

For each phase $[\tau_i, \tau_{i+1})$ which intersects the remaining rounds $[s_\ell(r), e_\ell(r)]$ (in an abuse of notation, 865 we'll conflate $e_{\ell}(r)$ with the *anticipated block end time* based on $s_{\ell}(r)$; that is, the end block time if 866 no episode restart occurs within the block). 867

Then, our strategy will be to map out in time the *local bad segments* or subintervals of $[\tau_i, \tau_{i+1})$ where 868 arm $a_r(B)$ incurs significant regret to arm a_t^{\sharp} in bin B, roughly in the sense of (\star) . The argument 869 will conclude by arguing that a well-timed replay is scheduled to detect some local bad segment in B, 870 before too many elapse. 871

As mentioned above, the difficulty here is that $a_r(B)$ is a random variable which depends on all the 872 randomness up to time $e_{\ell}(r)$. However, conditional on just the block start time $s_{\ell}(r)$, we define the 873 bad segments for a fixed arm a and then argue that if too many bad segments w.r.t. a elapse in the 874 block, arm a will be evicted in bin B. Crucially, this will hold uniformly over all arms a and thus for 875 arm $a = a_r(B)$, which bounds the regret of $a_r(B)$ in block $[s_\ell(r), e_\ell(r)]$. 876

Notation. Going forward, we will drop the dependence on the bin B, level r, block $[s_{\ell}(r), e_{\ell}(r)]$, 877 and episode $[t_{\ell}, t_{\ell+1})$ in certain definitions as they are fixed in the remainder of the analysis. We will 878 let $a_i^{\sharp}(B)$ denote the last safe of bin B in phase $[\tau_i, \tau_{i+1})$ (see Definition 8). 879

Definition 11. Fix an arm a and $s_{\ell}(r)$, and let $[\tau_i, \tau_{i+1})$ be any phase intersecting $[s_{\ell}(r), e_{\ell}(r)]$. 880 Define rounds $s_{i,0}(a), s_{i,1}(a), s_{i,2}(a) \dots \in [t_{\ell} \vee \tau_i, \tau_{i+1})$ recursively as follows: let $s_{i,0}(a) \doteq t_{\ell} \vee \tau_i$ 881 and define $s_{i,j}(a)$ as the smallest round in $(s_{i,j-1}(a), \tau_{i+1} \wedge e_{\ell}(r))$ such that arm a satisfies for some 882 fixed $c_{21} > 0$: 883

$$\sum_{i=s_{i,j-1}(a)}^{s_{i,j}(a)} \delta_t(a_i^{\sharp}(B), a) \ge c_{21} \log(T) \cdot (s_{i,j}(a) - s_{i,j-1}(a))^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}}.$$
 (26)

where $B' \supseteq B$ is the bin at level $r_{s_{i,j}(a)-s_{i,j-1}(a)}$, if such a round $s_{i,j}(a)$ exists. Otherwise, we 884 let the $s_{i,j}(a) \doteq \tau_{i+1} - 1$. We refer to the interval $[s_{i,j-1}(a), s_{i,j}(a))$ as a bad segment. We call 885 $[s_{i,j-1}(a), s_{i,j}(a))$ a proper bad segment if (26) above holds. 886

It will in fact suffice to constrain our attention to proper bad segments, since non-proper bad segments 887 $[s_{i,i-1}(a), s_{i,i}(a)]$ (where $s_{i,i}(a) = \tau_{i+1} - 1$ and (26) is reversed) will be negligible in the regret 888 analysis since there is at most one non-proper bad segment per phase $[\tau_i, \tau_{i+1})$ (i.e., the regret of 889 such non-proper bad segments is at most (19)). In what follows, we let $B' \supseteq B$ be the bin at level 890 $r_{s_{i,j}(a)-s_{i,j-1}(a)}$ where $[s_{i,j-1}(a), s_{i,j}(a))$ will be some proper bad segment, known from context. 891

Lemma 13. Any proper bad segment is at least K rounds long. 892

893 *Proof.* We have

t=

$$n_{B'}([s_{i,j}(a), s_{i,j+1}(a)]) \ge \sum_{s=s_{i,j}(a)}^{s_{i,j+1}(a)} \delta_t(a_i^{\sharp}(B), a) \cdot \mathbf{1}\{X_t \in B'\}$$

$$\ge \sum_{s=s_{i,j}(a)}^{s_{i,j+1}(a)} \delta_t(a_i^{\sharp}(B), a) \cdot \mu(B') - c_2 \left(\log(T) + \sqrt{\log(T)(s_{i,j+1}(a) - s_{i,j}(a)) \cdot \mu(B')}\right)$$

$$\ge c_{21} \log(T)(s_{i,j+1}(a) - s_{i,j}(a))^{\frac{1}{2+d}} \cdot K^{\frac{1+d}{2+d}} - c_2 \left(\log(T) + \sqrt{\log(T)(s_{i,j+1}(a) - s_{i,j}(a))\mu(B')}\right)$$

$$\ge \sqrt{K \cdot n_{B'}([s_{i,j}(a), s_{i,j+1}(a)])},$$

where the last inequality follows from Lemma 10 and choosing c_{21} large enough.

First, we relate our concentration bound (9) to (26), giving us control of the behavior of CMETA on proper bad segments. But, even before this, we establish an elementary lemma.

Lemma 14. Let $[s_{i,j}(a), s_{i,j+1}(a))$ be a proper bad segment defined w.r.t. arm a. Let $m \in \mathbb{N}$ be such that $r_{s_{i,j+1}(a)-s_{i,j}(a)} = 2^{-m}$. Then, for some $c_{22} = c_{22}(d) > 0$ depending on the dimension d:

$$\sum_{t=s_{i,j+1}(a)-K2^{(m-2)(2+d)-1}}^{s_{i,j+1}(a)} \delta_t(a_i^{\sharp}(B), a) \ge c_{22}\log(T) \cdot K^{\frac{1}{2+d}} \left(s_{i,j+1}(a) - s_{i,j}(a)\right)^{\frac{1+d}{2+d}}.$$
 (27)

Proof. First, we may assume $s_{i,j+1}(a) - s_{i,j}(a) \ge 4 \cdot K$ by choosing c_4 in (26) large enough (this will make m - 1 sensible).

First, observe $K2^{(m-1)(2+d)} \leq s_{i,j+1}(a) - s_{i,j}(a) < K2^{m(2+d)}$. Let $\tilde{s} = s_{i,j+1}(a) - S_{i,j}(a) + K2^{(m-2)(2+d)-1}$. Then, we have by (26) in the construction of the $s_{i,j}(a)$'s (Definition 11) that:

$$\sum_{t=\tilde{s}}^{s_{i,j+1}(a)} \delta_t(a_i^{\sharp}(B), a) = \sum_{t=s_{i,j}(a)}^{s_{i,j+1}(a)} \delta_t(a_i^{\sharp}(B), a) - \sum_{t=s_{i,j}(a)}^{\tilde{s}} \delta_t(a_i^{\sharp}(B), a)$$
$$\geq c_{21} \log(T) K^{\frac{1}{2+d}} \left((s_{i,j+1}(a) - s_{i,j}(a))^{\frac{1+d}{2+d}} - (\tilde{s} - s_{i,j}(a))^{\frac{1+d}{2+d}} \right)$$

903 Let $m_{i,j}(a) \doteq s_{i,j+1}(a) - s_{i,j}(a)$. Then, we have

$$m_{i,j}(a) \le K2^{m(2+d)} \implies \tilde{s} - s_{i,j}(a) = m_{i,j}(a) - K2^{(m-2)(2+d)-1} \le m_{i,j}(a)(1 - 2^{-2(2+d)-1}).$$

904 Plugging this into our earlier bound the constants become

$$1 - \left(1 - \frac{1}{2^{2(2+d)+1}}\right)^{\frac{1+d}{2+d}} > 0.$$

Note this last term is positive and only depends on d.

Lemma 15 (Bin-Count Dominates Concentration Error on Bad Segment). On event \mathcal{E}_1 , letting sor $\tilde{s} = s_{i,j+1}(a) - K2^{(m-2)(2+d)-1}$, we have for bin $B' \supseteq B$ at level $r_{s_{i,j+1}(a)-\tilde{s}}$:

$$n_{B'}([\tilde{s}, s_{i,j+1}(a)]) \ge 2c_1 \left(\log(T) + \sqrt{\log(T)(s_{i,j+1}(a) - \tilde{s})\mu(B')} \right).$$

Proof. Let $W = s_{i,j+1}(a) - \tilde{s}$. We first claim that $W \ge 2^{-2(2+d)} \cdot (s_{i,j+1}(a) - s_{i,j}(a))$. this follows from $s_{i,j+1}(a) - s_{i,j}(a) \le K \cdot 2^{m(2+d)}$ and

$$s_{i,j+1}(a) - \tilde{s} = K \cdot 2^{(m-2)(2+d)-1} = 2^{-2(2+d)-1} \cdot (K \cdot 2^{m(2+d)}) \ge 2^{-2(2+d)-1} \cdot (s_{i,j+1}(a) - s_{i,j}(a)).$$

- This will allow us to conflate W and $s_{i,j+1}(a) s_{i,j}(a)$ up to constants.
- Since $\overline{\delta}_t^B(a_i^{\sharp}(B), a) \leq 1$, we have that (27) of the previous lemma and concentration (namely, (11) of

Proposition 7; note that although $a_i^{\sharp}(B)$ is a random variable, it is a fixed and unchanging arm within [τ_i, τ_{i+1}) and hence $[\tilde{s}, s_{i,j}(a)]$) on $n_{B'}([\tilde{s}, s_{i,j+1}(a)])$ gives

$$n_{B'}([\tilde{s}, s_{i,j+1}(a)]) \ge \sum_{t=\tilde{s}}^{s_{i,j+1}(a)} \delta_t(a_i^{\sharp}(B), a) \cdot \mathbf{1}\{X_t \in B'\}$$

$$\ge c_4 \log(T) \cdot K^{\frac{1+d}{2+d}} (s_{i,j+1}(a) - s_{i,j}(a))^{\frac{1}{2+d}}$$

$$\ge c_4 \left(\log(T) + \sqrt{\log(T) \cdot W \cdot (K/W)^{\frac{d}{2+d}}}\right)$$

$$\ge c_1 \left(\log(T) + \sqrt{\log(T)(s_{i,j+1}(a) - \tilde{s})\mu(B')}\right)$$

where the last inequality follows from the strong density assumption (Assumption 2).

Now, we define a well-timed or perfect replay which, if scheduled, will be able to detect the badness of arm a in bin B over a proper bad segment $[s_{i,j}(a), s_{i,j+1}(a))$.

917 **Definition 12** (Perfect Replay). For a fixed proper bad segment $[s_{i,j}(a), s_{i,j+1}(a))$, define a 918 perfect replay as a Base-Alg (t_{start}, m) with $t_{\text{start}} \in [s_{i,j+1}(a) - K2^{(m-2)(2+d)} + 1, s_{i,j+1}(a) - K2^{(m-2)($

919 $K2^{(m-2)(2+d)-1}$ and $t_{\text{start}} + m \ge s_{i,j+1}(a)$.

The following proposition analyzes the behavior of a perfect replay and shows it will in fact evict arm a from $\mathcal{A}(B)$ within a proper bad segment $[s_{i,j}(a), s_{i,j+1}(a))$.

Proposition 16. Suppose event \mathcal{E}_1 holds. Let $[s_{i,j}(a), s_{i,j+1}(a))$ be a proper bad segment defined with respect to arm a. Let Base-Alg (t_{start}, m) be a perfect replay as defined above which becomes active at t_{start} (i.e., $Z_{t_{\text{start}},m} = 1$). Fix an integer $m \ge s_{i,j+1}(a) - s_{i,j}(a)$. Then:

(*i*) Base-Alg (t_{start}, m) will not evict arm $a_i^{\sharp}(B)$ from $\mathcal{A}(B)$ before round $s_{i,j+1}(a) + 1$ while active.

927 (ii) If
$$a \in A_t$$
 for all rounds $t \in [\tilde{s}, s_{i,j+1}(a))$ where $X_t \in B$, where $\tilde{s} = s_{i,j+1}(a) - K2^{(m-2)(2+d)-1}$, then arm a will be excluded from $\mathcal{A}(B)$ by round $s_{i,j+1}(a)$.

Proof. Suppose event \mathcal{E}_1 (i.e., our concentration bound (9) holds). For (i), if $a_i^{\sharp}(B)$ is evicted over [s_1, s_2] \subseteq [$s_{i,j}(a), s_{i,j+1}(a)$] from $\mathcal{A}(B')$ for bin $B' \supseteq B$ at level $r_{s_2-s_1}$ by Line 11 of Algorithm 2, then $a_i^{\sharp}(B)$ incurs significant regret in bin B' over [s_1, s_2] (following same reasoning as in Lemma 11). This is a contradiction to the definition of the last safe arm $a_i^{\sharp}(B)$ (Definition 8). This shows (i).

- For (ii), we first observe $\mathbb{E}[\hat{\delta}_t^B(a_i^{\sharp}(B), a) \mid \mathcal{F}_{t-1}] = \delta_t(a_i^{\sharp}(B), a)$ for any round $t \in [\tilde{s}, s_{i,j+1}(a)]$
- such that $X_t \in B$ if $a_i^{\sharp}(B), a \in \mathcal{A}_t$. Let $B' \supseteq B$ be the bin at level $r_{s_{i,j+1}(a)-\tilde{s}}$.
- Let $W = s_{i,j+1} \tilde{s}$. Then, by Lemma 14, we have by smoothness that:

$$\sum_{t=\tilde{s}}^{i,j+1} \delta_t(a_i^{\sharp}(B),a) \cdot \mathbf{1}\{X_t \in B'\} \ge c_4 \log(T) K^{\frac{1+d}{2+d}} \cdot W^{\frac{1}{2+d}} - n_{B'}([\tilde{s},s_{i,j+1}(a)]) \cdot r(B')$$

936 Next, note that

$$\log(T) \sqrt{K \sum_{s=\tilde{s}}^{s_{i,j+1}} \mu_s(B')} + \log(T) \cdot r(B') \sum_{s=\tilde{s}}^{s_{i,j+1}} \mu_s(B'),$$
(28)

937 is bounded above by the same order.

Next, we bound (28) below by an empirical analogue. Applying concentration on $n_{B'}([\tilde{s}, s_{i,j+1}(a)])$ which dominates the Bernstein error by the previous lemma, the above is further lower bounded by

$$\log(T) \left(\sqrt{K \cdot n_{B'}([\tilde{s}, s_{i,j+1}(a)])} + r(B') \cdot n_{B'}([\tilde{s}, s_{i,j+1}(a)]) \right)$$

meaning arm a will be evicted in B' over $[\tilde{s}, s_{i,j+1}(a)]$.

Furthermore, within Base-Alg (t_{start}, m) 's play, arms a and $a_i^{\sharp}(B)$ will **not** be evicted in any child of B' before round $s_{i,j+1}(a)$ because such an eviction can only happen through a child base algorithm of Base-Alg (t_{start}, m) which will necessarily use a level at least r_W . This is because of the way perfect replays are defined. By definition, the t_{start} is 'close enough' to the critical round $s_{i,j+1}(a) - K2^{(m-2)(2+d)-1}$ so that it will not use a different level than the perfect replay which starts exactly at this critical round.

Formally, we have that the maximum level a perfect replay is $s_{i,j+1}(a) - t_{\text{start}} \le K \cdot 2^{(m-2)(2+d)} - 1$ and so

$$\left(\frac{K}{s_{i,j+1}(a) - t_{\text{start}}}\right)^{\frac{1}{2+d}} \ge \left(\frac{K}{K \cdot 2^{(m-2)(2+d)} - 1}\right)^{\frac{1}{2+d}} \ge 2^{-(m-2)}.$$

949 On the other hand,

$$\left(\frac{K}{s_{i,j+1}(a)-\tilde{s}}\right)^{\frac{1}{2+d}} = \frac{1}{2^{m-2-\frac{1}{2+d}}} \in [2^{-(m-2)}, 2^{-(m-3)}).$$

Next, we show for any arm a (in particular, $a = a_r(B)$), a perfect replay characterized by Definition 12 is scheduled with high probability if too many bad segments w.r.t. a elapse, thus bounding the regret of a to $a_i^{\sharp}(B)$ over the phases $[\tau_i, \tau_{i+1})$ intersecting block $[s_{\ell}(r), e_{\ell}(r)]$.

954 **D.7** Bounding the Regret of the Last Master Arm $a_r(B)$ to the Last Safe Arm a_t^{\sharp}

Next, we bound the the regret of a fixed arm a to $a_i^{\sharp}(B)$ over the bad segments w.r.t. a in B. it should be understood that in what follows, we condition on $s_{\ell}(r)$. First, fix an arm a and define the *bad round* $s(a) > s_{\ell}(r)$ as the smallest round which satisfies, for some fixed $c_{23} > 0$:

$$\sum_{(i,j)} (s_{i,j+1}(a) - s_{i,j}(a))^{\frac{1+d}{2+d}} > c_{23} \log(T) (s(a) - t_{\ell})^{\frac{1+d}{2+d}}$$
(29)

where the above sum is over all pairs of indices $(i, j) \in \mathbb{N} \times \mathbb{N}$ such that $[s_{i,j}(a), s_{i,j+1}(a))$ is a proper bad segment with $s_{i,j+1}(a) < s(a)$. We will show that arm a is evicted within episode ℓ with high probability by the time the bad round s(a) occurs.

For each proper bad segment $[s_{i,j}(a), s_{i,j+1}(a))$, let $\tilde{s}_{i,j}(a) \doteq s_{i,j+1}(a) - K2^{(m-2)(2+d)-1}$ denote the special point of the bad segment and also let $m_{i,j} \doteq 2^n$ where $n \in \mathbb{N}$ satisfies:

$$2^{n} \ge s_{i,j+1}(a) - s_{i,j}(a) > 2^{n-1}.$$

Next, recall that the Bernoulli $Z_{m,t}$ decides whether Base-Alg (t,m) activates at round t (see Line 6 of Algorithm 1). If for some $t \in [\hat{s}_{i,j}(a), \tilde{s}_{i,j}(a)]$ where $\hat{s}_{i,j}(a) := s_{i,j+1}(a) - K2^{(m-2)(2+d)} + 1$, $Z_{m_{i,j},t} = 1$, i.e. a perfect replay is scheduled, then a will be evicted from $\mathcal{A}(B)$ by round $s_{i,j+1}(a)$ (Proposition 16). We will show this happens with high probability via concentration on the sum $\sum_{(i,j)} \sum_t Z_{m_{i,j},t}$ where j, i, t run through all $t \in [\hat{s}_{i,j}(a), \tilde{s}_{i,j}(a)]$ and all proper bad segments $[s_{i,j}(a), s_{i,j+1}(a)]$ with $s_{i,j+1}(a) < s(a)$. Note that these random variables only depend on the fixed arm a, the block start time $s_{\ell}(r)$, and the randomness of scheduling replays on Line 6. In particular, the $Z_{m_{i,j},t}$ are independent conditional on t_{ℓ} .

Then, a Chernoff bound over the randomization of CMETA on Line 6 of Algorithm 1 conditional on t_{ℓ} yields

$$\mathbb{P}\left(\sum_{(i,j)}\sum_{t} Z_{m_{i,j},t} \le \frac{\mathbb{E}\left[\sum_{(i,j)}\sum_{t} Z_{m_{i,j},t} \mid s_{\ell}(r)\right]}{2} \middle| s_{\ell}(r)\right) \le \exp\left(-\frac{\mathbb{E}\left[\sum_{(i,j)}\sum_{t} Z_{m_{i,j},t} \mid s_{\ell}(r)\right]}{8}\right)$$

⁹⁷³ We claim the error probability on the R.H.S. above is at most $1/T^3$. To this end, we compute:

$$\mathbb{E}\left[\sum_{(i,j)}\sum_{t} Z_{m_{i,j},t} \left| s_{\ell}(r) \right] \ge \sum_{(i,j)}\sum_{t=\hat{s}_{i,j}(a)}^{\tilde{s}_{i,j}(a)} \left(\frac{1}{m_{i,j}}\right)^{\frac{1}{2+d}} \left(\frac{1}{t-t_{\ell}}\right)^{\frac{1+d}{2+d}} \ge \frac{1}{4}\sum_{(i,j)} m_{i,j}^{\frac{1+d}{2+d}} \left(\frac{1}{s(a)-t_{\ell}}\right)^{\frac{1+d}{2+d}} \ge \frac{c_{7}}{4}\log(T),$$

where the last inequality follows from (29). The R.H.S. above is larger than $24 \log(T)$ for c_{23} large enough, showing that the error probability is small. Taking a further union bound over the choice of arm $a \in [K]$ gives us that $\sum_{(i,j)} \sum_t Z_{m_{i,j},t} > 1$ for all choices of arm a (define this as the good event $\mathcal{E}_3(s_\ell(r))$) with probability at least $1 - K/T^3$.

Recall on the event
$$\mathcal{E}_1$$
 the concentration bounds of Proposition 7 hold. Then, on $\mathcal{E}_1 \cap \mathcal{E}_3(s_\ell(r))$.

we must have $e_{\ell}(r) \leq s(a_r(B))$ since otherwise $a_r(B)$ would have been evicted in $\mathcal{A}(B)$ by some

perfect replay before the end of the block $e_{\ell}(r)$ by virtue of $\sum_{(i,j)} \sum_{t} Z_{m_{i,j},t} > 1$ for arm $a_r(B)$.

Thus, by the definition of the bad round $s(a_r(B))$ (29), we must have:

 $[s_i$

$$\sum_{a,j(a_r(B)), s_{i,j+1}(a_r(B))): s_{i,j+1}(a_r(B)) < e_\ell(r)} (s_{i,j+1}(a_r(B)) - s_{i,j}(a_r(B)))^{\frac{1+d}{2+d}} \le c_{23} \log(T) (e_\ell(r) - t_\ell)^{\frac{1+d}{2+d}}$$
(30)

31

Thus, by (26) in Definition 11, over the proper bad segments $[s_{i,j}(a_r(B)), s_{i,j+1}(a_r(B))]$ which elapse before the end of the block $e_{\ell}(r)$ in phase $[\tau_i, \tau_{i+1})$: the regret is at most

$$\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)} \delta_t(a_t^{\sharp}, a_r(B)) \le \sum_{(i,j)} \log(T) \cdot K^{\frac{1}{2+d}} m_{i,j}^{\frac{1+d}{2+d}} \le \log^2(T) \cdot K^{\frac{1}{2+d}} \cdot (e_{\ell}(r) - t_{\ell})^{\frac{1+d}{2+d}}$$

Over each non-proper bad segment $[s_{i,j}(a_r(B)), s_{i,j-1}(a_r(B)))$ and the last segment $[s_{i,j}(a_r(B)), e_{\ell}(r)]$, the regret of playing arm $a_r(B)$ to a_i^{\sharp} is at most $\log(T) \cdot r(B)^d \cdot K^{\frac{1}{2+d}} m_{i,j}^{\frac{1+d}{2+d}}$ by a similar series of calcuations and since there is at most one non-proper bad segment per phase $[\tau_i, \tau_{i+1})$ (see (26) in Definition 11).

So, we conclude that on event $\mathcal{E}_1 \cap \mathcal{E}_3(s_\ell(r))$:

t

$$\sum_{s_{\ell}(r)}^{e_{\ell}(r)} \delta_t(a_t^{\sharp}, a_r(B)) \le 2c_{23} \log^2(T) \sum_{i \in \mathsf{PHASES}(r,\ell)} K^{\frac{1}{2+d}} \cdot (\tau_{i+1} - \tau_i)^{\frac{1+d}{2+d}} \cdot (\tau_i - \tau_i)^{\frac{1+d}{2+d}} \cdot (\tau_$$

Taking expectation (all expectations below are conditional on X_T and the good event \mathcal{E}_2 over which we have concentration of covariate counts), we have by conditioning first on $s_\ell(r)$ and then on event

991 $\mathcal{E}_1 \cap \mathcal{E}_3(s_\ell(r))$:

$$\begin{split} & \mathbb{E}\left[\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)} \delta_{t}(a_{t}^{\sharp}, a_{r}(B))\right] \\ & \leq \mathbb{E}_{s_{\ell}(r)}\left[\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_{1} \cap \mathcal{E}_{3}(s_{\ell}(r))\}\sum_{t=s_{\ell}(r)}^{e_{\ell}(r)} \delta_{t}(a_{t}^{\sharp}, a_{r}(B))\Big|s_{\ell}(r)\right]\right] + T \cdot \mathbb{E}_{t_{\ell}}\left[\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_{1}^{c} \cup \mathcal{E}_{2}^{c}(s_{\ell}(r))\}\Big|s_{\ell}(r)\right]\right] \\ & \leq 2c_{23}\log^{2}(T)\mathbb{E}_{s_{\ell}(r)}\left[\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_{1} \cap \mathcal{E}_{3}(t_{\ell})\}\sum_{i \in \mathsf{PHASES}(\ell, r)} K^{\frac{1}{2+d}}(\tau_{i+1} - \tau_{i})^{\frac{1+d}{2+d}}\Big|s_{\ell}(r)\right]\right] + \frac{K}{T^{2}} \\ & \leq 2c_{23}\log^{2}(T)\mathbb{E}\left[\mathbf{1}\{\mathcal{E}_{1}\}\sum_{i \in \mathsf{PHASES}(\ell, r)}(\tau_{i+1} - \tau_{i})^{\frac{1+d}{2+d}}K^{\frac{1}{2+d}}\right] + \frac{1}{T}, \end{split}$$

where in the last step we bound $\mathbf{1}\{\mathcal{E}_1 \cap \mathcal{E}_3(t_\ell)\} \leq \mathbf{1}\{\mathcal{E}_1\}$ and apply tower law again. Plugging this into our earlier concentration bound on $\sum_{t=s_\ell(r)}^{e_\ell(r)} \delta_t(a_t^{\sharp}, a_r(B)) \cdot \mathbf{1}\{X_t \in B\}$, we conclude this part.

995 E Proof of Corollary 5

The proof of Corollary 5 will follow in a similar fashion to the proof of Corollary 2 in Suk and Kpotufe [2022], which relates the total-variation rates to significant shifts in the non-stationary MAB setting. A novel difficulty here is that our notion of significant shift $\tau_i(\mathbf{X}_T)$, $\tilde{L}(\mathbf{X}_T)$ (Definition 6) depends on the full context sequence \mathbf{X}_T , and so it is not clear how the (random) significant phases [$\tau_i(\mathbf{X}_T)$, $\tau_{i+1}(\mathbf{X}_T)$) relate to the total-variation V_T , which is a deterministic quantity.

Our strategy will be to first convert the regret rate of Theorem 3 into one which depends on a weaker *worst-case notion of significant shift* which does not depend on the observed \mathbf{X}_T . Although this notion of shift is weaker, it will be easier to relate to the total-variation quantity V_T .

1004 Let $\delta_t^a(x) := \max_{a' \in [K]} f_t^{a'}(x) - f_t^a(x)$ be the gap in mean rewards at the fixed context $x \in \mathcal{X}$.

Definition 13 (worst-case sig shift). Let $\tau_0 = 1$. Then, recursively for $i \ge 0$, the (i + 1)-th worstcase significant shift is recorded at time $\tilde{\tau}_{i+1}$, which denotes the earliest time $\tilde{\tau} \in (\tilde{\tau}_i, T]$ such that there exists $x \in \mathcal{X}$ such that for every arm $a \in [K]$, there exists round $s \in [\tilde{\tau}_i, \tilde{\tau}]$, such that

1008
$$\delta^a_s(x) \ge \left(\frac{K}{t-\tilde{\tau}_i}\right)^{\frac{1}{2+d}}$$

We will refer to intervals $[\tilde{\tau}_i, \tilde{\tau}_{i+1}), i \ge 0$, as worst-case (significant) phases. The unknown number of such phases (by time T) is denoted $\tilde{L}_{pop} + 1$, whereby $[\tilde{\tau}_{\tilde{L}_{pop}}, \tilde{\tau}_{\tilde{L}_{pop}+1})$, for $\tau_{\tilde{L}_{pop}+1} \doteq T + 1$, denotes the last phase.

1012 We next claim that

$$\mathbb{E}_{\mathbf{X}_{T}}\left[\sum_{i=0}^{\tilde{L}(\mathbf{X}_{T})} (\tau_{i+1}(\mathbf{X}_{T}) - \tau_{i}(\mathbf{X}_{T}))^{\frac{1+d}{2+d}}\right] \le c_{24} \sum_{i=0}^{\tilde{L}_{pop}} (\tilde{\tau}_{i+1} - \tilde{\tau}_{i})^{\frac{1+d}{2+d}}.$$

This follows since the empirical significant phases $[\tau_i(\mathbf{X}_T), \tau_{i+1}(\mathbf{X}_T))$ interleave the population analogues $[\tilde{\tau}_i, \tilde{\tau}_{i+1})$ in the following sense: at each significant shift $\tau_{i+1}(\mathbf{X}_T)$, for each arm $a \in [K]$, there is around $s \in [\tau_i(\mathbf{X}_T), \tau_{i+1}(\mathbf{X}_T)]$ such that for $\delta_s(X_{\tau_{i+1}}) > \left(\frac{K}{\tau_{i+1}-\tau_i}\right)^{\frac{1}{2+d}}$. This means there must be a worst-case significant shift $\tilde{\tau}_j$ in the interval $[\tau_i(\mathbf{X}_T), \tau_{i+1}(\mathbf{X}_T)]$ since the criterion of Definition 13 is triggered at $x = X_{\tau_{i+1}}$. Thus, by the sub-additivity of the function $x \mapsto x^{\frac{1+d}{2+d}}$. This also allows us to conclude that each worst-case significant phase $[\tilde{\tau}_i, \tilde{\tau}_{i+1})$ can intersect at most two significant phases $[\tau_i(\mathbf{X}_T), \tau_{i+1}(\mathbf{X}_T))$.

1020 Thus,

$$\begin{split} \sum_{i=0}^{\tilde{L}(\mathbf{X}_T)} (\tau_{i+1}(\mathbf{X}_T) - \tau_i(\mathbf{X}_T))^{\frac{1+d}{2+d}} &\leq \sum_{i=0}^{\tilde{L}(\mathbf{X}_T)} \sum_{j:[\tilde{\tau}_j, \tilde{\tau}_{j+1}) \cap [\tau_i(\mathbf{X}_T), \tau_{i+1}(\mathbf{X}_T)) \neq \emptyset} |[\tilde{\tau}_j, \tilde{\tau}_{j+1}) \cap [\tau_i(\mathbf{X}_T), \tau_{i+1}(\mathbf{X}_T))|^{\frac{1+d}{2+d}} \\ &\leq c_{24} \sum_{j=0}^{\tilde{L}_{pop}} (\tilde{\tau}_{j+1} - \tilde{\tau}_j)^{\frac{1+d}{2+d}}, \end{split}$$

where we use Jensen's inequality for $a^p + b^p \leq 2^{1-p}(a+b)^p$ for $p \in (0,1)$ and $a, b \geq 0$ in the last step to re-combine the subintervals of each worst-case significant phase $[\tilde{\tau}_j, \tilde{\tau}_{j+1})$.

1023 Then, it suffices to show

$$\sum_{j=0}^{\tilde{L}_{pop}} (\tilde{\tau}_{j+1} - \tilde{\tau}_j)^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}} \lesssim T^{\frac{1+d}{2+d}} \cdot K^{\frac{1}{2+d}} + (V_T \cdot K)^{\frac{1}{3+d}} \cdot T^{\frac{2+d}{3+d}}.$$
(31)

We first transform the total variation into a more flexible quantity depending on the reward functions $f_t^a(\cdot)$ and the full sequence \mathbf{X}_T .

Lemma 17. Let $G_t : \mathcal{X} \times [0,1]^K \to [-1,1]$ be any measurable function which takes the mean reward vector $f_t : \mathcal{X} \to [0,1]^K$ at round t as input, and outputs a real number in [-1,1]. Then, recalling \mathcal{D}_t is the joint distribution of X_t and Y_t , we have for $t = 2, \ldots, T$:

$$\|\mathcal{D}_t - \mathcal{D}_{t-1}\|_{\mathrm{TV}} \ge \frac{1}{2} \left(G_t(f_t) - G_t(f_{t-1}) \right).$$

Proof. This follows from the variational representation of the total variation distance [Polyanskiy and Wu, 2022, Theorem 7], which says for any measurable function $H : \mathcal{X} \times [0, 1]^K \to [-1, 1]$,

$$\|\mathcal{D}_{t} - \mathcal{D}_{t-1}\|_{\mathrm{TV}} \geq \frac{1}{2} \left(\mathbb{E}_{(X_{t}, Y_{t}) \sim \mathcal{D}_{t}} [H(X_{t}, Y_{t})] - \mathbb{E}_{(X_{t-1}, Y_{t-1}) \sim \mathcal{D}_{t-1}} [H(X_{t-1}, Y_{t-1})] \right).$$
(32)

In particular, we can take H to only depend on the mean reward functions.

Now, fix a worst-case significant phase $[\tilde{\tau}_i, \tilde{\tau}_{i+1})$ such that $\tau_{i+1} < T + 1$. By Definition 13, there exists a context $x_i \in \mathcal{X}$ such that for arm $a_i \in \operatorname{argmax}_{a \in [K]} f^a_{\tilde{\tau}_{i+1}}(x_i)$ we have there exists a round $t_i \in [\tau_i, \tau_{i+1}]$ such that:

$$\delta_{t_i}^{a_i}(x_i) > \left(\frac{K}{\tilde{\tau}_{i+1} - \tilde{\tau}_i}\right)^{\frac{1}{2+d}}.$$

1035 On the other hand, $\delta_{\tilde{\tau}_{i+1}}^{a_i}(x_i) = 0$ by the definition of arm a_i being the best at x_i at round $\tilde{\tau}_{i+1}$. Thus,

$$\left(\frac{K}{\tilde{\tau}_{i+1} - \tilde{\tau}_i}\right)^{\frac{1}{2+d}} < \delta_{t_i}^{a_i}(x_i) - \delta_{\tilde{\tau}_{i+1}}^{a_i}(x_i) = \sum_{t=t_i+1}^{\tau_{i+1}} \delta_t(a_i, x_i) - \delta_{t-1}(a_i, x_i).$$

For each round t = 2, ..., T, let $G_t(f_t) := \delta_t(a_i, x_i)$, where x_i is the associated context of the unique worst-case significant shift $\tilde{\tau}_{i+1}$ such that $t \in [\tilde{\tau}_i, \tilde{\tau}_{i+1})$ and where a_i is defined as above. Then, G_t only depends on the mean reward function $f_t : \mathcal{X} \to [0, 1]^K$ at round t and *not* on the observed contexts \mathbf{X}_T . Then, since $G_t(\cdot)$ satisfies the condition of Lemma 17, we must have

$$\sum_{i=1}^{L_{\text{pop}}} \left(\frac{K}{\tilde{\tau}_{i+1} - \tilde{\tau}_i}\right)^{\frac{1}{2+d}} < \sum_{t=2}^{T} G_t(f_t) - G_{t-1}(f_{t-1}) \le \sum_{t=2}^{T} \|\mathcal{D}_t - \mathcal{D}_{t-1}\|_{\text{TV}}.$$
(33)

Now, by Hölder's inequality for $p \in (0, 1)$ and $q \in \left(0, \frac{1+d}{2+d}\right)$:

$$\sum_{i=1}^{\tilde{L}_{pop}} (\tilde{\tau}_{i+1} - \tilde{\tau}_i)^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}} \le T^{\frac{1+d}{2+d}} K^{\frac{1}{2+d}} + \left(\sum_i K^{\frac{1}{2+d}} (\tilde{\tau}_{i+1} - \tilde{\tau}_i)^{-q/p}\right)^p \left(\sum_i K^{\frac{1}{2+d}} (\tilde{\tau}_{i+1} - \tilde{\tau}_i)^{\left(\frac{1+d}{2+d} + q\right) \cdot \frac{1}{1-p}}\right)^{1-p} K^{\frac{1}{2+d}} K^{\frac{1}{2+$$

In particular, letting $p = \frac{1}{3+d}$ and $q = \frac{1}{(2+d)(3+d)}$ and plugging in our earlier bound (33) makes the above RHS

$$V_T^{\frac{1}{3+d}} \cdot K^{\frac{1}{3+d}} \cdot T^{\frac{2+d}{3+d}}$$

1043

1044 F Proof of Theorem 1

We first note that it suffices to show (3) for integer $L \in [0, T] \cap \mathbb{N}$ as lower bounds for all other L follow via approximation and modifying the constant c > 0 in (3). Thus, going forward, fix $V \in [0, T]$ and $L \in \mathbb{Z} \cap [0, T]$.

At a high level, our construction will repeat L + 1 a hard environment for stationary contextual bandits. In particular, within each stationary phase of length T/(L+1) one is forced to pay a regret of $\left(\frac{T}{L+1}\right)^{\frac{1+d}{2+d}}$, summing to a total regret lower bound of $(L+1) \cdot \left(\frac{T}{L+1}\right)^{\frac{1+d}{2+d}} \approx L^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}}$.

¹⁰⁵¹ To get the rate in terms of V in (3), we will choose $L \propto V^{\frac{2+d}{3+d}} \cdot T^{\frac{1}{3+d}}$ appropriately and argue that ¹⁰⁵² the total-variation V_T is less than V, so that our constructed environment indeed lies in the family ¹⁰⁵³ $\mathcal{P}(V, L, T)$. This is similar to the arguments of the analogous lower bound [Besbes et al., 2019, ¹⁰⁵⁴ Theorem 1] for the non-contextual non-stationary bandit problem.

We start by establishing a lower bound for stationary Lipschitz context bandits. The construction is identical to that of Rigollet and Zeevi [2010, Theorem 4.1]. We only highlight a minor novelty in circumventing the reliance of the cited result on a positive "margin parameter" $\alpha > 0$.

Proposition 18. Suppose there are K = 2 arms. Then, there exists a stationary Lipschitz contextual bandit environment $\mathcal{E}(n)$ over n rounds such that for any algorithm π taking as input random variable U, independent of $\mathcal{E}(n)$, we have for some constant c > 0:

$$\mathbb{E}_{\mathcal{E}(n),U}[R(\pi, X_T)] \ge c \cdot n^{\frac{1+d}{2+d}}.$$

Proof. Let the covariates X_t be uniformly distributed on $[0,1]^d$ at each round $t \in [n]$, so that $\mu_X \equiv \text{Unif}\{[0,1]^d\}$. For ease of presentation, let us reparametrize the two arms as +1 and -1.

At each round $t \in [n]$, let arm -1 have reward $Y_t^{-1} \sim \text{Ber}(1/2)$ and let arm 1 have reward Y_t^{1064} $Y_t^1 \sim \text{Ber}(f(X_t))$ where $f : \mathcal{X} \to [0, 1]$ is some function to be defined. Let

$$M := \left\lceil \left(\frac{n}{8e}\right)^{\frac{1}{2+d}} \right\rceil$$

We next partition $\mathcal{X} = [0, 1]^d$ into a regular grid $\mathcal{Q} = \{q_1, \dots, q_{M^d}\}$, where q_k denotes the center of bin $B_k, k = 1, \dots, M^d$. Specifically, for each index $\mathbf{k} = (k_1, \dots, k_d) \in \{1, \dots, M\}^d$, we define the bin B_k as:

$$B_k = \left\{ x \in \mathcal{X} : \frac{k_\ell - 1}{M} \le x_\ell \le \frac{k_\ell}{M}, \ell = 1, \dots, d \right\}.$$

1068 Define $C_{\phi} \doteq 1/4$. Then, let $\phi : \mathbb{R}^d \to \mathbb{R}_+$ be a smooth function defined by:

$$\phi(x) = \begin{cases} 1 - \|x\|_{\infty} & 0 \le \|x\|_{\infty} \le 1\\ 0 & \|x\|_{\infty} > 1 \end{cases}$$

1069 It's straightforward to verify ϕ is 1-Lipschitz over \mathbb{R}^d .

- Now, define the integer $m = \lfloor \mu \cdot M^d \rfloor$ where $\mu \in (0, 1)$ is chosen small enough to ensure $m \le M^d$.
- 1071 Define $\Sigma_m = \{-1, 1\}^m$ and for any $\omega \in \Omega_m$, define the function f_ω on $[0, 1]^d$ via

$$f_{\omega}(x) = 1/2 + \sum_{j=1}^{m} \omega_j \cdot \phi_j(x),$$

where $\phi_j(x) \doteq M^{-1} \cdot C_{\phi} \cdot \phi(M \cdot (x - q_j)) \cdot \mathbf{1}\{x \in B_j\}$. Then, the optimal arm at context $x \in \mathcal{X}$ in this environment is given by $\pi_f^*(x) \doteq \operatorname{sgn}(f(x) - 1/2)$.

1074 Then, define the family C of environments induced by f_{ω} for $\omega \in \Omega_m$. Next, let $\text{Int}(B_k)$ be the ℓ_{∞} 1075 ball centered at q_k of radius $\frac{1}{2M}$. Then, we have for any $x \in \text{Int}(B_k)$,

$$|f_{\omega}(x) - 1/2| \ge M^{-1} \cdot C_{\phi}/2$$

1076 Then, the worst-case regret over the family of environments in C is at least

$$\sup_{f \in \mathcal{C}} \mathbb{E} \sum_{t=1}^{n} |f^{(1)}(X_t) - f^{(2)}(X_t)| \cdot \mathbf{1}\{\pi_t(X_t) \neq \pi^*(X_t)\}$$

$$\geq \frac{C_{\phi}}{2M} \sup_{f \in \mathcal{C}} \mathbb{E} \sum_{t=1}^{n} \sum_{j=1}^{m} \mathbf{1}\{\pi_t(X_t) \neq \pi^*(X_t), X_t \in \mathrm{Int}(B_j)\}.$$

Lower bounding the remaining supremum on the above RHS display by $\Omega(n)$ follows the same steps as the proof of Theorem 4.1 in Rigollet and Zeevi [2010]. In particular, the algorithm π may depend on additional randomness U, independent of the environment, which is ignorable in the KL calculations by use of chain rule. Plugging in the earlier choice of M this makes the above RHS at least $\Omega(n^{\frac{1+d}{2+d}})$.

1082

Given Proposition 18, the $(L+1) \cdot \left(\frac{T}{L+1}\right)^{\frac{1+d}{2+d}}$ lower bound immediately follows by constructing a random environment which consists of L+1 independent repetitions of the stationary environment $\mathcal{E}(T/(L+1))$. Any such constructed environment clearly has at most L global shifts. Note that the regret over a given stationary phase of length $\frac{T}{L+1}$ is lower bounded by $\left(\frac{T}{L+1}\right)^{\frac{1+d}{2+d}}$ regardless of the information learned prior to that phase, as such information can be formalized as exogeneous randomness U in Proposition 18 w.r.t. the fixed stationary phase.

Next, we tackle the lower bound $V^{\frac{1}{3+d}} \cdot T^{\frac{2+d}{3+d}}$ in terms of total-variation budget V. First, if $V < \frac{1}{2\tau}$

$$\left(T^{\frac{1+d}{2+d}} + T^{\frac{2+d}{3+d}} \cdot V^{\frac{1}{3+d}}\right) \wedge \left((L+1)^{\frac{1}{2+d}}T^{\frac{1+d}{2+d}}\right)$$

is minimized by the first term which is of order $T^{\frac{1+d}{2+d}}$. Thus, using Proposition 18 with a single stationary phase $\mathcal{E}(T)$ gives lower bound of the right order. Such an environment clearly has total-variation $V_T = 0 \leq V$.

Let $\Delta \doteq \left[\left(\frac{T}{V} \right)^{\frac{2+d}{3+d}} \right] \leq \left[T^{\frac{1}{3+d}} \right]$ and consider $L + 1 = T/\Delta$ stationary phases of length Δ . Then, by 1094 the previous arguments we have the regret is lower bounded by 1095

$$(L+1)^{\frac{1}{2+d}} \cdot T^{\frac{1+d}{2+d}} = \frac{T}{\Delta^{\frac{1}{2+d}}} \ge \frac{T}{2^{\frac{1}{3+d}} (T/V)^{\frac{1}{3+d}}} \propto T^{\frac{2+d}{3+d}} \cdot V^{\frac{1+d}{3+d}}.$$

Additionally, $T^{\frac{2+d}{3+d}} \cdot V^{\frac{1+d}{3+d}}$ dominates $T^{\frac{1+d}{2+d}}$ since $V \ge \left(\frac{1}{T}\right)^{\frac{3+d}{2+d}}$. Thus, the regret lower bound is 1096 proven in terms of V. 1097

It remains to show the total-variation V_T is at most V in the above constructed environments so that 1098 it lies in the family $\mathcal{P}(V, L, T)$. 1099

Clearly, the instantaneous total-variation $\|\mathcal{D}_t - \mathcal{D}_{t-1}\|_{TV} = 0$ for all rounds t not being the start 1100 of a new stationary phase. On the other hand, for such a round t, we have that since conditioning 1101 increases the TV [Polyanskiy and Wu, 2022, Theorem 7.5(c)], the instantaneous TV is at most: 1102

$$\|\mathcal{D}_t - \mathcal{D}_{t-1}\|_{\mathrm{TV}} \le \mathbb{E}_{x \sim \mu_X} \left[\|\mathcal{D}_t(Y_t | X_t = x) - \mathcal{D}_{t-1}(Y_{t-1} | X_{t-1} = x) \|_{\mathrm{TV}} \right].$$

Since $Y_t^a | X_t = x \sim \text{Ber}(f_t^a(x))$, we have the RHS' inner TV quantity is just the total variation between Bernoulli's or $\max_{a \in [2]} |f_t^a(x) - f_{t-1}^a(x)|$. Carefully analyzing the variations in the constructed Lipechitz reward functions in the proof of Proposition 18 reveals this TV between Bernoulli's is at most $\frac{e^{\frac{1}{2+d}}}{8^{\frac{1}{2+d}}} \cdot \left(\frac{L+1}{T}\right)^{\frac{1}{2+d}}$ (note the attached constant is < 1 for all $d \in \mathbb{N} \cup \{0\}$). 1103 1104 1105 1106

- Summing over phases, we have 1107

$$V_T \le (L+1)^{\frac{3+d}{2+d}} \cdot T^{-\frac{1}{2+d}} = T \cdot \left(\frac{1}{\Delta}\right)^{\frac{3+d}{2+d}} = V.$$

1108